2 3

4

5 6 7

8

9

10

11

12

13 14 15

16 17

> 18 19

Weak effects of fire, large herbivores, and their interaction on regrowth of harvested trees in two West African savannas

Andrew T. Tredennick¹, Moussa Karembé², Fadiala Dembélé³, Justin Dohn⁴, and Niall P. Hanan⁵

Postal addresses:

¹Department of Wildland Resources, Utah State University, Logan, UT, USA

²Universite du Bamako, Bamako, Mali

³Institut Polytechnique Rurale du Katibougou, Katibougou, Mali

⁴Natural Resource Ecology Laboratory and Graduate Degree Program in Ecology, Colorado

State University, Fort Collins, CO, USA

⁴Geospatial Science Center of Excellence, South Dakota State University, Brookings, SD, USA

Keywords: savanna, tree harvest, fire, herbivory, Mali, West Africa

Corresponding author: Andrew Tredennick (atredenn@gmail.com)

Theory and empirical evidence for the impacts of fire and herbivory in savannas is well 21 established - they are top-down disturbances that maintain savannas in disequilibrium states 22 away from potential tree cover. In African savannas the demand for fuelwood is extremely high, 23 so tree harvest likely also has an impact, both directly and indirectly, on tree cover, density, and 24 biomass. Many savanna trees resprout vigorously from the base after harvest. However, 25 harvested trees regenerate as saplings susceptible to fire and browsing, so harvest may have 26 important demographic consequences. Here, we report the effects of tree harvest, and its 27 interaction with fire and herbivory, on savanna dynamics by analyzing woody regrowth 28 29 following a harvest in arid Sahelian and mesic Guinean savannas in Mali, West Africa. Tree 30 harvest resulted in an overall reduction in wood production per tree compared to growth in nonharvested trees. Regrowth, either biomass or height, did not differ among fire and herbivory 31 32 treatments. Our results suggest that the resprouting abilities that savanna trees have evolved to cope with frequent fire are essential for surviving tree harvest and subsequent disturbance. In 33 these savannas, regrowth is rapid enough in the first growing season to escape the impact of dry 34 season fires. 35

Top-down disturbances are essential for determining and maintaining the savanna biome 37 (Sankaran, Ratnam & Hanan, 2004; Bond, 2008). In particular, fire and herbivory act to 38 maintain savanna systems in a state of disequilibrium where bioclimatic potential tree cover is 39 rarely attained (Sankaran et al., 2005; Staver, Archibald & Levin, 2011b). However, other top-40 41 down forces exist that may contribute to the co-dominance of trees and grasses in tropical savannas. Tree harvest for fuel, particularly in African savannas, represents an under-explored 42 and potentially important disturbance due to the large annual demand in Africa and the density of 43 44 human populations in savanna regions. Tree harvest may play an independent role in determining savanna tree cover through direct removal, but may also amplify the well-known impacts of fire 45 and herbivory (Tredennick & Hanan, In Press). 46

47 Fire is likely the most important driver of savanna structure (i.e., tree cover and biomass) beyond climatic constraints (Bucini & Hanan, 2007; Staver, Archibald & Levin, 2011a). The 48 49 impact of fire is observed at landscape and regional scales in savannas but the direct effect of fire operates at the tree population level by reducing recruitment rates (Higgins, Bond & Trollope, 50 2000) more than through removal of adult tree biomass (Hanan et al., 2008). Many savanna tree 51 52 species are able to resprout vigorously when top-kill occurs during fire; a key functional trait that reduces mortality and provides an opportunity for rapid recovery of adult populations (Bond & 53 Midgley, 2001; Hoffmann, Orthen & Nascimento, 2003; Hoffmann et al., 2012; Clarke et al., 54 55 2013). In particular, the "Bottleneck Hypothesis" suggests recurrent fire increases tree sapling mortality, thus decreasing recruitment rates of trees to adult size-class (Higgins et al., 2000; 56 Sankaran et al., 2004). 57

The impact of tree harvest, on the other hand, is likely to contrast that of fire because harvest generally impacts adults but not seedlings. The removal of aboveground biomass of adult trees during harvest, much like fire, results in resprouting of cut individuals – individuals now potentially caught as saplings in the fire (Bond, 2008) and browse traps (Staver & Bond, 2014) as saplings. Thus, while tree harvest does result in a net loss of biomass from savannas, the indirect effect of harvest in terms of its interaction with fire may also be important at the population and community levels (Tredennick & Hanan, In Press).

The effects of herbivory on savanna tree structure are less general, often depending upon 65 feeding type (grazing vs. browsing) and how herbivore disturbance interacts with fire. For 66 67 example, grazing can reduce grass fuel loads thereby increasing sapling recruitment while reducing sapling-grass competition (Holdo et al., 2009; Riginos, 2009; February et al., 2013). 68 69 However, browsing results in loss of biomass and can, potentially, inhibit reproduction of adult trees and reduce sapling survival and growth directly (Augustine & McNaughton, 2004; Staver 70 et al., 2009; Moncrieff et al., 2011; Staver & Bond, 2014) or indirectly through an interaction 71 with fire (Staver et al., 2009; Staver & Bond, 2014). 72

In African savannas it is clear that climate, top-down disturbances, and the interactions of these factors are important in determining landscape scale tree cover or density. But how does tree harvest for fuelwood, an essential ecosystem service in rural Africa (Arnold, Köhlin & Persson, 2006), modify these interactions? Recent modeling work predicts that tree harvest can have a large impact at forest-savanna and savanna-grassland ecotones, but within savanna the effects should be minimal due to the resprouting ability of trees (Tredennick & Hanan, In Press). Even with the ability to resprout, a negative impact of fire on regrowth after harvest is expected.

80	Here we report a four-year manipulative study on the isolated and interactive effects of					
81	tree harvest, fire, and herbivory conducted in two savanna sites in Mali, West Africa - one semi-					
82	arid South Sahelian site and a mesic North Guinean site. In 2008 we deployed fully factorial					
83	experimental structures at each site that included fire and herbivory exclusion. We simulated tree					
84	harvest in 2010. Our study focuses on woody regrowth (one year and three years post-harvest)					
85	following the simulated harvest and how regrowth is impacted by herbivory, fire, and climate.					
86	Based on our knowledge of savanna systems and our recent modeling work (Tredennick &					
87	Hanan, In Press), we developed five hypotheses:					
88	1. In the absence of herbivory, fire will have a depressing effect on regrowth					
89	(biomass and height) after harvest due to the potential for top-kill of shoots by					
90	fires.					
91	2. The effect of fire will be smaller in plots with herbivory because grassy fuel loads					
92	will be reduced and fires will be less intense.					
93	3. In isolation, the presence of herbivory will have no impact on regrowth (biomass					
94	and height) after harvest if dominated by grazers (e.g. cattle) but a direct negative					
95	impact if dominated by browsers (sheep and goats).					
96	4. The effects described in hypotheses 1 and 2 will be additive over time (larger					
97	effects in 2013 relative to 2011).					
98	5. On average, across treatments, regrowth will be greater in the mesic savanna					
99	relative to the arid savanna.					
100						
101	Materials and Methods					
102	Study Area					

The study was conducted at two sites that bookend the savanna tropical rainfall gradient across Mali, West Africa (Figure 1A). Both sites are underlain by similar silty-sand soils as identified by local expert knowledge. Mean annual precipitation is 577 mm year⁻¹ at Lakamané and 1,132 mm year⁻¹ at Tiendéga (mean annual precipitations calculated for 1981-2010 from the CRU TS3.10 Dataset; Harris et al. 2014).

Each site receives varying degrees of fire, herbivory, and fuelwood harvesting. All field 108 sites were established in designated State Forests and Grazing Lands where they are protected 109 from agriculture and other development, but not from fuelwood harvest, fire, or herbivory by 110 111 domestic herbivores. Tree cover at Tiendéga is approximately 60% and at Lakamané is approximately 12%. The most common large grazers are domestic mixed herds of cattle, sheep, 112 113 and goats – large wild herbivores are effectively absent. The herds at Tiendéga are primarily cattle, while at Lakamané herds are predominately mixtures of sheep and goats. Herbivory varies 114 temporally and spatially as herds migrate south to north in the wet season (late June – August). 115 All sites are under some constant herbivory throughout other months. Other grazers include 116 termites and small rodents. Fire is an important disturbance at both sites and ambient fire return 117 time is approximately two years at Tiendéga and one year at Lakamané. However, we note 118 below that we imposed annual burns in our fire treatment plots. 119

120

121 Treatments and wood harvest simulation

At each of the sites, a replicated 2×2 factorial plot design with fire (*F*)/no-fire (*f*) and large herbivore (*H*)/no large herbivore (*h*) treatments was established in 2008 (Figure 1B). Four replicate plots (50×50 m each) were deployed at each site. The control (fire and herbivory, *FH*) allows for grazing of all large herbivores in the area and fire at ambient levels. Herbivore

exclosures (fire/no fire and no herbivory; Fh and fh) are fenced with 1.5 m wire-mesh fencing 126 with 8 cm mesh size that excludes all domestic and larger wild herbivores (but not rodents or 127 small mammals). Fire exclosures (no fire and herbivory/no herbivory; fH and fh) were 128 established using 5 m fire breaks cleared of all small trees and most large trees, all shrubs, and all 129 herbaceous cover. Firebreaks were re-cleared at the end of each rainy season. The herbivory 130 131 treatments (FH and fH) were open to grazing at ambient levels to mimic grazing and browsing pressure in the surrounding savanna matrix. Fire treatments (FH and Fh) were burned annually 132 in December. 133

134 To implement the fuelwood harvest simulation in each fire/herbivory treatment we first delineated a 25×25 m plot in two of four treatment plots for each treatment (Figure 1B). The 135 plots were chosen based on the availability of 10 or more individuals of the preferred fuelwood 136 species (as identified by local knowledge) at each site. While species selection was not random, 137 our study better reflects reality by focusing on the tree species actually used for fuelwood. 138 139 Between the two plots we selected 20 individuals that were tagged at the base and given unique identification numbers. We chose *Deterium microcarpum* Guill. And Perr. (n = 20 per treatment) 140 at Tiendéga and *Combretum glutinosum* Perr. (n = 20 per treatment) at Lakamané. Tree selection 141 142 was haphazard, but the sample reflected the size structure of the tree community at each site. We only selected trees with diameters at the base greater or equal to 2 cm. 143

We simulated tree harvest under each combination of fire and herbivory (*FH*, *fH*, *Fh*, and *fh*). Before harvesting the trees we took baseline measurements of basal diameter to develop allometric relationships between those variables and biomass, and also to assess the relationship between initial biomass and subsequent regrowth. Trees were cut at 10 cm from the soil surface using bow saws. All tree biomass was removed from the site following harvest. Following initial harvest during the peak-growing season of 2010, biomass regrowth was harvested and measured at peak growing season (July – August) in 2011 (t_{harv} + 1) and just after the growing season (October – November) in 2013 (t_{harv} + 3). A random and equal sample of trees (10 per treatment per year at each site) was measured in each observation year (2011 and 2013). This means that for trees in fire plots (*FH* and *Fh*), those measured in 2011 grew through one dry season fire, while those measured in 2013 grew through three annual dry season fires.

We measured woody biomass regrowth in 2011 and 2013, and in 2013 we measured the 155 heights of trees harvested in 2010 but not re-harvested for biomass measures in 2011. For the 156 157 biomass measures, all biomass associated with a harvested tree was collected and we separated leaf and wood biomass to obtain wet weights. Sub-samples of wood were taken and dried at 100° 158 159 C to account for water-content contribution to wet weights taken in the field. Species-specific 160 dry:wet weight ratios were applied to all wet weights to convert to dry biomass. Data on wet weight samples for Lakamané were damaged in a storm in 2013. Thus, for the 2013 data from 161 162 Lakamané we use a dry: wet weight ratio of 0.5 (near the value from 2011 of 0.52) and we performed a sensitivity analysis to ensure our results are robust to changes in the dry:wet weight 163 ratio (see Supplemental Information). Importantly, the loss of this data does not impact our 164 165 statistical tests for treatment effects.

We used non-harvested trees in the no fire/no herbivory plots (*fh*) as controls to measure annual wood growth in the absence of harvest or other disturbance. At each site we selected 20 trees distributed throughout the four *fh* plots that matched the species chosen for harvest. We measured initial circumference and installed dendrometer bands to measure wood growth. We measured dendrometer band change each year (2011 and 2013) during peak growing season. Throughout the rest of the text we refer to these as "control trees."

173 Data analysis

PeerJ PrePrints

To convert diameter and height measures to estimates of biomass we used allometric 174 relationships. We log-transformed the response variable (biomass) and the predictor variable 175 (diameter or height) because the data had a multiplicative error distribution (see Tredennick, 176 Bentley & Hanan, 2013). We compared candidate models using Akaike Information Criterion 177 (see Supplementary Information). The best model includes log(diameter) as the predictor 178 variable and is interspecific ($R^2 = 0.92$, P < 0.0001). We used the model to estimate initial 179 aboveground wood biomass of harvested trees immediately before harvest, initial biomass of 180 control trees, and annual biomass growth of control trees in 2011 and 2013 (see Supplementary 181 Information for details). Regrowth biomass of harvested trees was measured directly as 182 183 described above.

Our analysis focuses on annual wood biomass change, defined as relative growth rate (*rgr*) of wood for control (non-harvested) trees and relative regrowth rate (*rrgr*) of wood for harvested trees. We calculated *rgr* for control trees as

187

188
$$rgr_t = \frac{\ln(B_t) - \ln(B_{initial})}{year_t - year_{initial}}$$
 [Eq. 1]

189

and *rrgr* for harvested trees as

191

192
$$rrgr_{t} = \frac{\ln(B_{t} + B_{initial}) - \ln(B_{initial})}{year_{t} - year_{initial}}$$
[Eq. 2]

where rgr_t and $rrgr_t$ are relative growth and relative regrowth rates for measurement year *t*, *B*_t is biomass at year *t*, and *B*_{*initial*} is initial biomass measured in 2010. The term *year*_t - *year*_{*initial*} simply accounts for the number of years between measurements so that *rgr* and *rrgr* are consistent per year rates. Note that in Eq. 2 *B*_t represents the regrowth of a harvested tree.

We aimed to determine the quantitative difference in regrowth attributable to fire, 198 199 herbivory, and their interaction, both overall and within sites. To estimate the effects of fire and herbivory on *rrgr* we analyzed the biomass data using factorial ANOVA at two levels: 1) data 200 pooled across sites but with a site random effect on the intercept (mixed effects model), and 2) 201 202 within-site. For the factorial ANOVAs, we used the 'lme4' package (Bates et al., 2014) in the program R (2012) to fit the mixed effects model using the 'lmer' function and the 'lm' procedure 203 to fit the within-site models. We report F- and P-values from the 'anova' function in R applied to 204 205 the 'lmer' and 'lm' models. We also performed two-way ANOVAs to compare harvested tree biomass regrowth vs. control tree wood growth using the 'aov' procedure in R and performed 206 207 post hoc Tukey's tests when applicable using the 'TukeyHSD' procedure in R. We used the same approach to test for biomass regrowth differences between the two sites using pooled treatment 208 data. For all analyses, *rrgr* and *rgr* were log-transformed to meet model assumptions of 209 210 normality of residuals and homoscedasticity. So all reported statistics and statistically significant differences refer to the log-transformed version of rrgr and rgr, but in the figures we show non-211 transformed values. 212

We tested for treatment effects on post-harvest heights in 2013 using factorial ANOVA. Height was log transformed to meet model assumptions. We also used logistic regression to estimate the probability of a tree growing to a specific escape height three years after harvest as a function of initial tree size. We focused on escape heights of 1, 2, and 3 meters because these

reflect the range relevant to escaping the effects of fire and herbivory as reported elsewhere 217 (Bond & Midgley, 2001; Bond, 2008; Staver & Bond, 2014). 218 All code and data to reproduce our results has been deposited on Dryad (DRYAD LINK 219 HERE AFTER ACCEPTANCE) and archived at 220 http://atredennick.github.com/HarvestExperiment. 221

222

225

Results 223

Tree harvest depressed woody growth relative to that of non-harvested trees at both sites (Figure 2). On average, non-harvested trees had relative growth rates that were 6.9 times greater than harvested trees in Tiendéga. In Lakamané, non-harvested trees had relative regrowth rates 5 times greater than harvested trees. Mortality of harvested trees was low (Table 1). Percent mortality did not differ among treatments, years, or sites.

Since we initiated annual fires, we know the fire treatments were effective. Herbivore 229 230 exclosures were also effective, as indicated by grass biomass being significantly greater within herbivore exclosures relative to biomass outside of exclosures in 2010, 2011, and 2012 (P < 0.05231 in all years at each site except for 2012 at Lakamané). Unfortunately, we do not have data on 232 233 grazer and browser use via dung counts. Fire and herbivory had little effect on post-harvest regrowth (Figure 3). We failed to detect any significant effects of fire, herbivory, or their 234 interaction at both sites and in both years (Table 2). While not statistically significant, the effect 235 236 of fire is stronger at both sites in 2013 relative to 2011 (Table 2), giving some qualitative support for our hypothesis that treatment effects would emerge over time. Given the lack of statistically 237 238 strong treatment effects we do not report effect sizes of herbivore and fire exclusion on tree 239 relative regrowth rates (but the magnitude of the effects can be visually interpreted in Figure 3).

When we pooled *rrgr* among treatments, *rrgr* at Lakamané was greater than at Tiendéga in 2011 (P = 0.006), but in 2013 *rrgr* was greater in Tiendéga (P = 0.004; but see Supplementary Information for caveats about this test). *rrgr* at Tiendéga was greater in 2013 relative to 2011 (P= 0.001; Figure 4). *rrgr* among years at Lakamané were likely similar (we do not provide a Pvalue due to the sensitivity of this test to the dry:wet weight ratio; see Supplementary Information). See Supplementary Information for the effects of different dry:wet weight ratios in Lakamané in 2013 on these statistical tests.

There were no significant treatment effects on tree height three years after harvest, but average height was greater at Tiendéga than Lakamané (P = 0.012; Figure 4A). Initial tree biomass is a significant predictor of the probability that a tree will reach 2 and 3 meters three years after harvest in Tiendéga (P = 0.004 and P = 0.023, respectively) and for 2 meters in Lakamané (P = 0.022) (Figure 4B,C).

253 **Discussion**

A defining trait of savanna tree species is their ability to resprout vigorously. While this 254 trait likely evolved by natural selection in response to frequent fires in savanna ecosystems 255 256 (Bond & Midgley, 2001; Shackleton, 2001), the same trait benefits savanna trees that are harvested for fuelwood (Shackleton, 2001). Indeed, we observed near zero mortality after 257 harvesting savanna trees in two West African savannas (Table 1). This is unsurprising, as prior 258 259 work in a South African savanna reported similar results (Shackleton, 2001). More surprising is that relative regrowth rate and height of harvested trees did not differ among treatments of fire 260 exclusion, large herbivore exclusion, or the exclusion of both fire and herbivores. Given that 261 262 recent modeling work suggests an important interaction between tree harvest and fire, we

expected that woody regrowth in fire plots would be lower because of shoot mortality in fires.
However, we found no support for an interaction between harvest and fire, or for any of the other
hypotheses advanced in the Introduction (Table 2, Figure 4A). Perhaps most surprisingly,
treatment effects did not emerge over time, at least statistically.

In the case of the herbivore treatments, it could be that grazing by cattle has very little impact, either through reduced grass competition or reduced fuel loads, on post harvest regrowth. Other experiments have shown that the direct effect of browsers is much greater than the indirect effects of grazers (Staver and Bond, 2014). Thus, it is not surprising that at the grazer-dominated mesic site (Tiendéga) we found no effect of herbivory, especially since our treatments only ran for three years. More time may be required for the indirect effects of grass-tree competition and increased fuel loads to emerge.

274 Contrary to the possible indirect effects of grazing, a positive effect of browsing release in Lakamané, where goats and sheep dominate herds, should have been immediate. Yet we found 275 no statistical support for an herbivore effect (Table 2, Figure 4A). The most likely explanation is 276 that harvested trees regrow very quickly and are able to escape from the 'browse trap' within one 277 year (Staver and Bond, 2014) (Figure 4B,C). Three years after harvest the average height of 278 279 regrowing trees across all treatments at Lakamané was 1.74 meters (s.d. = 0.66) (Figure 4A). Thus, as we discuss below for the case of fire, harvested trees are able to regrow rapidly and 280 avoid disturbance impacts, unlike seedlings and saplings growing from seed. 281

The lack of an effect on *rrgr* or height by excluding fire, either in the presence or absence of herbivores, is more perplexing. Decades of work in savanna systems points toward the key role of fire in limiting seedling and sapling transitions to larger size classes, so why did we not observe a positive effect of fire exclusion here? It is likely that post-harvest saplings simply grow

faster than saplings growing from seed. In that case, post-harvest saplings are able to escape the 286 firetrap quickly by drawing on large stores of belowground carbohydrates (Hoffmann et al., 287 2003; Hoffmann, Orthen & Franco, 2004). We found evidence of this since initial tree size was 288 a good predictor of a tree reaching specific escape heights (Figure 4B,C). By growing fast, post-289 harvest saplings avoid subsequent top kill during the dry season. This scenario could occur 290 291 within a single growing season if the harvest event occurs early enough, or the scenario could play out over two growing seasons: in the first season the tree is harvested, but remains virtually 292 dormant through the dry season fires, then can grow quickly in the subsequent growing season. 293 294 Therefore, even after three annual burns, for trees measured in 2013, we did not detect a fire effect on regrowth (Table 2, Figure 4A). 295

Most other work on savanna tree demographics has not considered the effect of tree 296 297 harvest (but see Zida et al., 2007). Thus, the insight that fire can maintain the savanna state by limiting sapling transitions to adult trees is constrained to systems where the saplings that are 298 299 negatively impacted by fire have grown from seed or have resprouted from a sapling partially or completely top-killed by a previous fire. In other words, most work on this topic does not 300 consider that coppice saplings growing from the rootstock of large trees may have faster growth 301 rates than even the fastest growing "regular" saplings (those from seed or regrowth from 302 saplings). In our previous work we assumed equal growth rates among "new" and "harvested" 303 saplings (Tredennick & Hanan, In Press). Clearly this assumption needs to be refined. To do so 304 305 will require comparative studies of sapling growth rates when grown from seed and across a size spectrum at time of harvest. 306

307 Another surprising result is that biomass regrowth at the mesic site (Tiendéga) was not 308 consistently greater than at the arid site (Lakamané). In fact, in the first year after harvest,

309 biomass regrowth was statistically larger at the arid site. This runs counter to our naïve hypothesis based on climate constraints alone (e.g., a positive relationship between growth rates 310 and precipitation). However, there are many factors that may interact with precipitation to inhibit 311 regrowth in mesic savannas relative to arid savannas, like competition for light with adult trees 312 (Iponga, Milton & Richardson, 2008) or competition with grass (Riginos, 2009) and tree roots 313 314 (Sea & Hanan, 2012) for soil water. And, we did find that heights three years after harvest were greater in the mesic site than in the arid site (Figure 4A). Fully understanding the susceptibility 315 of regenerating trees to fire and herbivory will require disentangling those interacting resource-316 317 based drivers.

To conclude, while fuelwood harvest decreases overall tree population wood-growth 318 319 rates in post-harvest years, we found little evidence that presence of fire and large herbivores, or 320 both, affect post-harvest regrowth of trees in semi-arid and a mesic savannas of West Africa. Our findings are broadly consistent with work from Burkina Faso that also found no interaction 321 between tree harvest, fire, and grazing (Zida et al., 2007). Our work suggests that the effects of 322 disturbance, however small, are most important during the first year of growth after harvest. 323 Savanna trees resprouting after harvest are able to draw on large nutrient stores held in 324 325 substantial root systems. Thus, their growth rates are rapid, allowing them to escape the impacts of disturbance quickly. The timing of disturbance is likely important for determining the success 326 327 and magnitude of post-harvest regrowth.

328

329 Acknowledgements

Many thanks to Salif Traouré, Labassoun Keita, and numerous field assistants in Mali for help
with biomass harvests. The National Science Foundation, through a grant to N.P.H. (DEB-

335 **References**

- ARNOLD, J. E. M., KÖHLIN, G. & PERSSON, R. (2006) Woodfuels, livelihoods, and policy
 interventions: Changing Perspectives. *World Development*, 34, 596-611.
- AUGUSTINE, D. J. & MCNAUGHTON, S. J. (2004) Regulation of shrub dynamics by native
 browsing ungulates on East African rangeland. *Journal of Applied Ecology*, 41, 45-58.
 - BATES, D., MARTIN, M., BOLKER, B. M. & WALKER, S. (2014) lme4: Linear mixed effects models using Eigen and S4.
 - BOND, W. J. (2008) What Limits Trees in C4 Grasslands and Savannas? *Annual Review of Ecology, Evolution, and Systematics*, **39**, 641-659.
 - BOND, W. J. & MIDGLEY, J. J. (2001) Ecology of sprouting in woody plants: the persistence niche. *Trends in Ecology & Colution*, **16**, 45-51.
- BUCINI, G. & HANAN, N. P. (2007) A continental-scale analysis of tree cover in African
 savannas. *Global Ecology and Biogeography*, 16, 593-605.
- 348 CLARKE, P. J., LAWES, M. J., MIDGLEY, J. J., LAMONT, B. B., OJEDA, F., BURROWS,
- G. E., ENRIGHT, N. J. & KNOX, K. J. E. (2013) Resprouting as a key functional trait:
- how buds, protection and resources drive persistence after fire. *New Phytologist*, **197**, 19-
- 351 35.
- 352 FEBRUARY, E. C., HIGGINS, S. I., BOND, W. J. & SWEMMER, L. (2013) Influence of
- competition and rainfall manipulation on the growth responses of savanna trees and
- 354 grasses. *Ecology*, **94**, 1155-1164.

HANAN, N. P., SEA, W. B., DANGELMAYR, G. & GOVENDER, N. (2008) Do fires in
savannas consume woody biomass? A comment on approaches to modeling savanna
dynamics. *Am Nat*, **171**, 851-856.

358 HARRIS, I., JONES, P. D., OSBORN, T. J. & LISTER, D. H. (2014) Updated high-resolution

- 359 grids of monthly climatic observations the CRU TS3.10 Dataset. *International Journal* 360 of Climatology, 34, 623-642.
- HIGGINS, S. I., BOND, W. J. & TROLLOPE, W. S. W. (2000) Fire, resprouting and variability:
 a recipe for grass-tree coexistence in savanna. *Journal of Ecology*, 88, 213-229.

HOFFMANN, W., ORTHEN, B. & FRANCO, A. (2004) Constraints to seedling success of

savanna and forest trees across the savanna-forest boundary. Oecologia, 140, 252-260.

HOFFMANN, W. A., GEIGER, E. L., GOTSCH, S. G., ROSSATTO, D. R., SILVA, L. C. R.,

LAU, O. L., HARIDASAN, M. & FRANCO, A. C. (2012) Ecological thresholds at the savanna-forest boundary: how plant traits, resources and fire govern the distribution of tropical biomes. *Ecology Letters*, **15**, 759-768.

HOFFMANN, W. A., ORTHEN, B. & NASCIMENTO, P. K. V. D. (2003) Comparative Fire
 Ecology of Tropical Savanna and Forest Trees. *Functional Ecology*, **17**, 720-726.

HOLDO, R. M., SINCLAIR, A. R. E., DOBSON, A. P., METZGER, K. L., BOLKER, B. M.,

372 RITCHIE, M. E. & HOLT, R. D. (2009) A Disease-Mediated Trophic Cascade in the
373 Serengeti and its Implications for Ecosystem C. *PLoS Biol*, 7, e1000210.

374 IPONGA, D. M., MILTON, S. J. & RICHARDSON, D. M. (2008) Superiority in competition for

- light: A crucial attribute defining the impact of the invasive alien tree Schinus molle
- 376 (Anacardiaceae) in South African savanna. *Journal of Arid Environments*, **72**, 612-623.

377 MONCRIEFF, G. R., CHAMAILLÉ-JAMMES, S., HIGGINS, S. I., O'HARA, R. B. & BOND,

- W. J. (2011) Tree allometries reflect a lifetime of herbivory in an African savanna.
 Ecology, 92, 2310-2315.
- RIGINOS, C. (2009) Grass competition suppresses savanna tree growth across multiple
 demographic stages. *Ecology*, **90**, 335-340.
- SANKARAN, M., HANAN, N. P., SCHOLES, R. J., RATNAM, J., AUGUSTINE, D. J.,
 - CADE, B. S., GIGNOUX, J., HIGGINS, S. I., LE ROUX, X., LUDWIG, F., ARDO, J.,
 - BANYIKWA, F., BRONN, A., BUCINI, G., CAYLOR, K. K., COUGHENOUR, M. B.,
 - DIOUF, A., EKAYA, W., FERAL, C. J., FEBRUARY, E. C., FROST, P. G. H.,
 - HIERNAUX, P., HRABAR, H., METZGER, K. L., PRINS, H. H. T., RINGROSE, S.,
 - SEA, W., TEWS, J., WORDEN, J. & ZAMBATIS, N. (2005) Determinants of woody cover in African savannas. *Nature*, **438**, 846-849.

SANKARAN, M., RATNAM, J. & HANAN, N. P. (2004) Tree-grass coexistence in savannas

- revisited insights from an examination of assumptions and mechanisms invoked in
 existing models. *Ecology Letters*, 7, 480-490.
- 392 SEA, W. B. & HANAN, N. P. (2012) Self-thinning and Tree Competition in Savannas.
- *Biotropica*, **44**, 189-196.
- 394 SHACKLETON, C. (2001) Managing regrowth of an indigenous savanna tree species
- (Terminalia sericea) for fuelwood: the influence of stump dimensions and post-harvest
 coppice pruning. *Biomass and Bioenergy*, 20, 261-270.
- 397 STAVER, A. C., ARCHIBALD, S. & LEVIN, S. (2011a) Tree cover in sub-Saharan Africa:
- Rainfall and fire constrain forest and savanna as alternative stable states. *Ecology*, 92,
 1063-1072.

4(STAVER, A. C., ARCHIBALD, S. & LEVIN, S. A. (2011b) The global extent	and determinants
40	of savanna and forest as alternative biome states. Science, 334, 230-232	2.
40	STAVER, A. C. & BOND, W. J. (2014) Is there a 'browse trap'? Dynamics of	herbivore impacts
40	on trees and grasses in an African savanna. Journal of Ecology, 102, 59	95-602.
40	STAVER, A. C., BOND, W. J., STOCK, W. D., VAN RENSBURG, S. J. & W	VALDRAM, M. S.
40	(2009) Browsing and fire interact to suppress tree density in an African	savanna.
2 40	Ecological Applications, 19, 1909-1919.	
40	R CORE TEAM (2012) R: a language environment for statistical computing. F	R Foundation for
5 40	Statistical Computing, Vienna, Austria.	
4(TREDENNICK, A. T., BENTLEY, L. P. & HANAN, N. P. (2013) Allometric	Convergence in
4	Savanna Trees and Implications for the Use of Plant Scaling Models in	Variable
4	Ecosystems. PLoS ONE, 8, e58241.	
4	TREDENNICK, A. T. & HANAN, N. P. (In Press) Effects of tree harvest on the	ne stable state
4	dynamics of savanna and forest. The American Naturalist.	
4	ZIDA, D., SAWADOGO, L., TIGABU, M., TIVEAU, D. & ODÉN, P. C. (200	07) Dynamics of
4	sapling population in savanna woodlands of Burkina Faso subjected to	grazing, early fire
4	and selective tree cutting for a decade. Forest Ecology and Managemer	<i>ıt,</i> 243, 102-115.
4		
4		

PeerJ PrePrints | http://dx.doi.org/10.7287/peerj.preprints.718v1 | CC-BY 4.0 Open Access | rec: 18 Dec 2014, publ: 18 Dec 2014

420

Tables

421 **Table 1.** Percent mortality of harvested trees each measurement year. The raw number of dead

trees out of the total sample is shown in the parentheses.

423

Site	Percent mortality					
	2011	2013	Years combined			
Tiendéga	2.5% (1/40)	2.5% (1/40)	2.5% (2/80)			
Lakamané	0% (0/38)	5.4% (2/37)	2.7% (2/75)			

Table 2. Results from the factorial ANOVA for the effects of herbivore and fire exclusion on

426 post-harvest biomass regrowth (*rrgr*) one and three years after harvest. Measurement periods are

427 denoted as in the main text where $t_{harv} = 2010$. Statistical tests result from running an ANOVA

			Fire		Herbivroy		Fire × Herbivory	
Measurement								
Period	Site	n^{c}	F _{1,n-3}	Р	F _{1,n-3}	Р	F _{1,n-3}	P
$t_{\rm harv} + 1$	All ^a	75	0.062	na ^b	0.01	na ^b	0.061	na ^b
	Tiendéga	38	0.004	0.95	0.164	0.688	0.059	0.809
	Lakamané	37	0.053	0.819	1.482	0.232	2.207	0.146
$t_{\rm harv} + 3$	All ^a	74	0.062	na ^b	0.01	na ^b	0.061	na ^b
	Tiendéga	39	0.327	0.571	0.749	0.393	0.723	0.401
	Lakamané	35	2.739	0.108	0.874	0.357	0^{d}	1 ^d

428 ('anova') on the results from a linear model ('lm') in the program R.

^aTo estimate the effects of fire and herbivory across both sites we used a mixed-effects model with a site random effect on the intercept.

^bP values are not statistically appropriate when computed based on mixed-effects models, so we do not report them here. However, the very small F values indicate that the fire and herbivory effects are not important.

^cSample sizes may be different from those in Table 2 because we excluded dead trees and one observation with an unreliable estimate of initial biomass.

⁴³⁹ ^dThese are values rounded to three digits.

440

Figure Legends

Figure 1. (A) Map of Mali, West Africa showing locations of the two study sites (Tiendéga in the South; Lakamané in the North) and broad trends in mean annual precipitation. (B) The experimental layout at each site. F = fire present; f = fire excluded. H = large herbivores present; h = large herbivores excluded. As described in the main text, a 25 x 25 meter plot in two of four treatment replicates was used for this study.

Figure 2. Boxplots comparing relative (re)growth rates by site for control (growth rates) and harvested trees (regrowth rates) across both measurement years. Letters denote statistically significant (P < 0.1) pairwise comparisons from post-hoc Tukey's HSD tests. Growth rates are expressed per unit aboveground wood biomass of trees in 2010 (prior to cutting in harvested treatments; see Eqs. 1 and 2).

Figure 3. Boxplots of relative biomass regrowth rates (*rrgr*) 1 year (2011) and 3 years (2013) after harvest for each treatment at each site. There are no strong statistical differences among treatments within each year; however, treatment effects are further diminished by 2013 (see Table 1). We tested for interactions between treatments and the year harvested, but found no statistically important interaction effects.

459

Figure 4. Boxplots of harvested tree heights after three years (A) and results from logistic
regression estimating the probability of reaching an escape height (2 or 3 meters) after three
years as a function of pre-harvest tree size (B and C). In A, there are no statistical differences
among treatments, but average height after three years was greater in Tiendéga than in Lakamané

- 464 (P = 0.012). In C, only the regression for a 2 meter escape height has initial biomass as
- significant, but we still show the 3 meter regression for context.



Figure 1. (A) Map of Mali, West Africa showing locations of the two study sites (Tiendéga in the South; Lakamané in the North) and broad trends in mean annual precipitation. (B) The experimental layout at each site. F = fire present; f = fire excluded. H = large herbivores present; h = large herbivores excluded. As described in the main text, a 25 x 25 meter plot in two of four treatment replicates was used for this study.



Figure 2. Boxplots comparing relative (re)growth rates by site for control (growth rates) and harvested trees (regrowth rates) across both measurement years. Letters denote statistically significant (P < 0.1) pairwise comparisons from post-hoc Tukey's HSD tests. Growth rates are expressed per unit aboveground wood biomass of trees in 2010 (prior to cutting in harvested treatments; see Eqs. 1 and 2).



Figure 3. Boxplots of relative biomass regrowth rates (*rrgr*) 1 year (2011) and 3 years (2013) after harvest for each treatment at each site. There are no strong statistical differences among treatments within each year; however, treatment effects are further diminished by 2013 (see Table 1). We tested for interactions between treatments and the year harvested, but found no statistically important interaction effects.

PeerJ PrePrints



Figure 4. Boxplots of harvested tree heights after three years (A) and results from logistic regression estimating the probability of reaching an escape height (2 or 3 meters) after three years as a function of pre-harvest tree size (B and C). In A, there are no statistical differences among treatments, but average height after three years was greater in Tiendéga than in Lakamané (P = 0.012). In C, only the regression for a 2 meter escape height has initial biomass as significant, but we still show the 3 meter regression for context.

2

3 4

5

6 7 8

9

10

11

12 13

14 15

16 17

18 19

20

21

22

23

24

Supplemental Information for: Weak effects of fire, large herbivores, and their interaction on regrowth of harvested trees in two West African savannas

Andrew T. Tredennick, Moussa Karembé, Fadiala Dembélé, Justin Dohn, and Niall P. Hanan

Supplemental Information 1: Assessing impact of dry:wet weight ratio on statistical results

There are two statistical tests in our paper that could be biased by an incorrect dry:wet wood ratio for Lakamané in 2013:

- 1. The comparison of relative regrowth rate between sites in 2013.
- 2. The comparison of relative regrowth rate between years in Lakamané.

Comparisons in Lakamané within the year 2013 are the same regardless of the dry:wet weight ratio because the ratio is constant among treatments within sites. To test the sensitivity of our statistical results for the two comparisons above we used a simple simulation where we performed the same statistical test after applying a range of dry:wet weight ratios to the wet weights observed in Lakamané. In the main text we base our conclusions on a dry:wet ratio of 0.5 - a value similar to the one estimated from the 2011 data in Lakamané, which was collected at nearly the same time of year.

26 Our simulation test shows that a statistical difference (p < 0.1) between Lakamané regrowth

(pooled across treatments) and Tiendéga regrowth in 2013 occurs when the dry:wet weight ratio 27

is less than 0.65 (Figure S1; open circles). For the comparison between years in Lakamané, a 28

29 significant difference occurs when the dry:wet ratio is less than 0.54 (Figure S1; grey circles).

Thus, we are confident in our conclusion that regrowth rates between sites in 2013 are 30

statistically different, as reported in the main text. However, since the threshold for a significant 31

difference between years in Lakamané is very near the dry:wet weight ratio we use, we are less 32

confident in that result. Therefore, we do not report a significant difference between years in 33

Lakamané in the main text. 34

35

36 We note that all statistical tests other than the two listed above are unaffected by this particular

dry:wet weight ratio. 37



Figure S1. Results of our simulation test for bias induced by the wet:dry weight ratio. The open circles show the simulation results for the difference between sites in 2013. The grey filled circles show the simulation results for the difference between years in Lakamané. The solid black line shows a p-value of 0.1, and the dashed vertical line shows the dry:wet ratio of 0.5 that we used for the analyses presented in the main text.

PeerJ PrePrints | http://dx.doi.org/10.7287/peerj.preprints.718v1 | CC-BY 4.0 Open Access | rec: 18 Dec 2014, publ: 18 Dec 2014

56

57

58 59

60

61

62

63

64

65

66

67

68

69 70 71

72 73

45 Supplemental Information 2: Converting diameter measurements to biomass 46 using allometric models

We measured the growth of non-harvested trees within our fully protected plots (*fh*; fire and large herbivores excluded) using dendrometer bands. Each measurement year (2011 and 2013) we recorded the change in diameter from the initial diameter measured in 2010. To compare these growth rates to those of the harvested trees, we converted diameter measurements to biomass estimated using a simple allometric model. We used data collected in 2010 from harvested trees to estimate the parameters in a log-log allometric regression:

55 $\log(y) = a + b\log(D)$ [Eq. S1]

where y is tree wood biomass, a is a normalizing constant, b is the scaling parameter, and D is tree diameter (see Tredennick et al. 2013 for more details on allometric models in savannas and the data used here).

We fit two competing models to the data: (1) a mixed-effects model with a random effect of species on the slope and (2) an interspecific model with no random effects. AICc (AIC for corrected for small datasets) among the two models were roughly equivalent (46.93 for (1) and 47.99 for (2)). So, we decided to use the interspecific model with the following parameters: a = 3.26 and b = 2.78 (P < 0.0001, $R^2 = 0.92$; Figure S2). We used this equation to estimate biomass of non-harvested trees in 2010, 2011, and 2013. Then, using those estimates, we were able to estimated the relative biomass growth rate (rgr_{year} = [B_{year}/B₂₀₁₀]/year-2010) of non-harvested trees in each measurement year.

Below are the details of the log-log allometric model fit in R. Y_2 is wood biomass for each tree and x is the diameter of each tree.

```
74
    Call:
    lm(formula = log(Y2) \sim log(x))
75
76
77
    Residuals:
                    10
                                       30
78
         Min
                         Median
                                               Max
    -0.94569 -0.22800 -0.04236 0.17408
                                           1.09724
79
80
81
    Coefficients:
                 Estimate Std. Error t value Pr(>|t|)
82
    (Intercept)
                   3.2568
                              0.2990
                                        10.89 6.03e-13
83
                   2.7763
                              0.1399
                                        19.85
                                               < 2e-16 ***
84
    log(x)
    - - -
85
                     0 (**** 0.001 (*** 0.01 (** 0.05 (. 0.1 ( 1
    Signif. codes:
86
87
    Residual standard error: 0.3974 on 36 degrees of freedom
88
    Multiple R-squared: 0.9163, Adjusted R-squared:
89
                                                         0.914
                    394 on 1 and 36 DF, p-value: < 2.2e-16
    F-statistic:
90
91
```



Figure S2. Fitted regression of log(Biomass) ~ log(Diameter). The fitted line is the mean
prediction from the log-log regression described above.

References

Tredennick, A.T., Bentley, L.P., and Hanan, N.P. (2013) Allometric Convergence in Savanna Trees and Implications for the Use of Plant Scaling Models in Variable Ecosystems. PLoS One 8(3): e58241. doi:10.1371/journal.pone.0058241