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Twisted tale of the tiger: the case of inappropriate data and deficient science

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Publications in peer reviewed journals are often looked upon as tenets on which future scientific thought is built. Published information is not always flawless and errors in published research should be expediently reported, preferably by a peer review process. We review a recent publication by Gopalaswamy et al (2015;

doi:10.1111/2041-210X.12351) that challenges the use of "double sampling" in large scale animal surveys. Double sampling is often resorted to as an established economical and practical approach for large scale surveys since it calibrates abundance indices against absolute abundance, thereby potentially addressing the statistical shortfalls of indices. Empirical data used by Gopalaswamy et al. (2015) to test their theoretical model, relate to tiger sign and tiger abundance referred to as an Index Calibration experiment (IC-Karanth). These data on tiger abundance and signs should be paired in time and space to qualify as a calibration experiment for double sampling, but original data of IC-Karanth show lags of (up to) several years. Further, data points used in the paper do not match the original sources. We show that by use of inappropriate and incorrect data collected through a faulty experimental design, poor parameterization of their theoretical model, and selectively-picked estimates from literature on detection probability, the inferences of this paper are highly questionable. We highlight how the results of Gopalaswamy et al. were further distorted in popular media. If left unaddressed, Gopalaswamy et al. paper could have serious implications on statistical design of large-scale animal surveys by propagating unreliable inferences.

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1 Twisted tale of the tiger: the case of inappropriate data and deficient science. 2 3 Authors: 1) Qamar Qureshi, email: qnq@wii.gov.in, Wildlife Institute of India, Chandrabani, 4 5 Dehradun, India 2) Rajesh Gopal, email: rajeshgopal.sg.gtf@gmail.com, Global Tiger Forum, Jor Baug, 6 7 New Delhi, India 8 3) Yadvendradev Jhala*, email: jhalay@wii.gov.in, Wildlife Institute of India, Dehradun, 9 India. 10 11 * Corresponding author 12 13 **Abstract:** 14 Publications in peer reviewed journals are often looked upon as tenets on which future scientific 15 thought is built. Published information is not always flawless and errors in published research 16 should be expediently reported, preferably by a peer review process. We review a recent 17 publication by Gopalaswamy et al (2015; doi:10.1111/2041-210X.12351) that challenges the use 18 of "double sampling" in large scale animal surveys. Double sampling is often resorted to as an 19 established economical and practical approach for large scale surveys since it calibrates 20 abundance indices against absolute abundance, thereby potentially addressing the statistical 21 shortfalls of indices. Empirical data used by Gopalaswamy et al. (2015) to test their theoretical 22 model, relate to tiger sign and tiger abundance referred to as an Index Calibration experiment 23 (IC-Karanth). These data on tiger abundance and signs should be paired in time and space to 24 qualify as a calibration experiment for double sampling, but original data of IC-Karanth show 25 lags of (up to) several years. Further, data points used in the paper do not match the original 26 sources. We show that by use of inappropriate and incorrect data collected through a faulty 27 experimental design, poor parameterization of their theoretical model, and selectively-picked 28 estimates from literature on detection probability, the inferences of this paper are highly 29 questionable. We highlight how the results of Gopalaswamy et al. were further distorted in 30 popular media. If left unaddressed, Gopalaswamy et al. paper could have serious implications on 31 statistical design of large-scale animal surveys by propagating unreliable inferences.

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34	Scientific method operates by testing competing hypothesis or by choosing between
35	alternate models that best explain observed data. Hypothesis and models that survive
36	repeated testing by careful experimentation are published through rigorous scrutiny by a
37	peer review process, these subsequently become scientific theory (Gauch 2012). An
38	incorrect experimental design, inappropriate data collection protocol, and selective data,
39	used for analysis from telemetered Florida panthers (Puma concolor) (Gross 2005)
40	resulted in a peer reviewed publication of habitat use and preference (Maehr & Cox 1995)
41	in Conservation Biology. The results were subsequently used for land use planning and
42	policy (Maehr & Deason 2002) which resulted in the best panther habitat being lost to
43	developmental projects (Gross 2005). In the ideal world, response to deficiencies in science
44	is best made through a peer review process, since scientists understand the intricacies of
45	the scientific method probably more than others (Parsons & Wright 2015).
46	
47	In a recent paper "An examination of index-calibration experiments: counting tigers at
48	macroecological scales" published in the journal Methods of Ecology and Evolution,
49	Gopalaswamy et al (2015a) supposedly demonstrate that as part of their long-term, large-scale
50	data on tiger abundance and index (IC-Karanth) they did not find any relationship between tiger
51	abundance and scat index. They conclude that attempting to use double sampling (Cochran 1977
52	Eberhardt & Simmons 1987; Pollock et al. 2002) to establish relationships between any index of
53	abundance and actual abundance is a futile effort. In particular, they claim that the relationship
54	between tiger sign index and tiger abundance published by Jhala et al. (2011a) to be improbable
55	since they could not reproduce it by their data or theoretical model. We review Gopalaswamy et
56	al (2015a) to show that by the use of a) wrong ecological parameters for their theoretical model,
57	b) selectively picked references from literature, c) inappropriate and incorrect data, and d) data
58	not collected in an experimental setup, the inferences drawn by their paper are questionable.
59	a) Use of inadequate Ecological Parameters:
60	The basic premise for index calibration by double sampling is that animal sign intensity or count
61	data should reflect underlying animal abundance. Often due to logistic and economic constraints
62	large scale estimates of abundance are not possible through statistically rigorous methods that

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63 explicitly estimate and correct for detection (e.g. capture-mark-recapture or DISTANCE 64 sampling). Double sampling approach as described initially by Cochran (1977) and applied to 65 wildlife surveys by Eberhardt & Simmons (1987), allows us to address this limitation by 66 measuring a relatively easy and economically less expensive, but potentially biased index of 67 abundance across all sampling units, while simultaneously estimating detection corrected 68 abundance from within a subset of these sampling units (Conroy & Carroll 2009, Williams et al. 69 2002). Subsequently, the potentially biased index is calibrated against the unbiased estimate of 70 abundance or actual abundance using a ratio or regression approach (Skalski et al. 2005). Pollock 71 et al. (2002) recommend double sampling as a sensible large-scale survey design for most 72 species. 73 To prove their point of view, Gopalaswamy et al (2015a) use detection probability (p) 74 estimates from tiger occupancy studies as a surrogate for detection probability of tiger scat for 75 parameterizing their theoretical model. This p is the probability of finding (or not finding) tiger 76 sign on a single survey in an area occupied by tigers. Gopalaswamy et. al. (2015) confuse p of 77 occupancy surveys with the probability of finding (or missing) an individual sign (in this case 78 tiger scat) (r). In other words, p represents the number of surveys out of the total surveys that are 79 likely to detect presence of tigers in an occupied site, while r represents the proportion of tiger 80 signs that are detected (or missed) in a single survey. The two are not the same i.e. $p \neq r$. For 81 example, a survey that detected nine out of 10 signs present or another that detected one sign out 82 of 10 signs are both considered as having 100% detection of tiger presence (p=1) for an 83 occupancy survey, but r for each of these surveys is 0.9 and 0.1 respectively. Thus, detection 84 probability (p) of occupancy surveys is not informative on per capita detection rates (r) of tiger 85 sign. For estimating r the correct approach would be to use a double blind observer experimental 86 design (Buckland et al. 2010, Nichols et al. 2000), where two observers would walk the same 87 trail some distance apart and record observed tiger scat without communicating with each other. 88 The scats being missed by each of them could then be used to estimate the probability of missing 89 scats entirely. 90 Also, in occupancy surveys all kinds of signs are often used to detect tigers (pugmarks, 91 scat, scrape, rake marks, direct sightings, vocalization, tiger kills, etc). Karanth et al. (2011a)

have used both tiger scat and tiger pugmark to detect tigers in a grid for estimating occupancy.

Thus, detection probability of occupancy in these surveys is the compounded probability of



94 occurrence and detection of both scat and pugmark on a single survey which cannot be teased 95 apart and used as a surrogate for detecting individual scats. From the above it is clear that the use 96 of occupancy detection probability to parameterize detection probability of tiger scat in the 97 theoretical model of Gopalaswamy et al. (2015) is wrong. Typically in a double sample survey 98 the index is measured without an estimate of its detection, by calibrating this potentially biased 99 index against abundance, double sampling elegantly addresses the issue of detection and other 100 sources of variability in the index (Conroy and Carroll 2009). 101 102 b) Selectively picked references. 103 Not only do Gopalswamy et al (2015) use an incorrect detection probability (derived for 104 occupancy studies) in place of a double observer based detection probability for sign intensity for 105 their theoretical model, they were selective in picking low estimates of detection probability with 106 high coefficient of variation (CV) from those available in published literature. The estimates of 107 detection probability p at 1 km segments (0.17) and its CV (1) from Karanth et al (2011a) were 108 used, claiming that these were the only parameter estimates available. The use of low p and 109 extraordinarily high CV to suggest that detection of tiger presence for occupancy survey is in 110 general low and highly variable. These parameters play an important role in subsequent 111 derivations in the paper. Gopalswamy et al (2015) have ignored other published estimates of these parameters obtained by sampling large areas and derived by following the same field and 112 113 analytical protocols. These publications report far higher p with much smaller CV (Harihar and 114 Panday (2012), p = 0.951 SE 0.05; Barbara-Meyer (2013), p = 0.65 SE 0.08). The low p and high 115 CV reported by Karanth et al (2011a) is likely due to poor design and not a norm in detecting 116 tiger presence. In our experience tigers uses scat, scrape and rakes to advertise their presence and 117 it is highly unlikely that tiger signs will have such a low detectability unless the population is 118 very low, survey design is poor, or data is collected by inexperienced/untrained persons. c) Inappropriate and incorrect data 119 120 Throughout the paper the authors have used data and parameters related to tigers published by K. 121 Ullas Karanth (a co-author on the paper) and colleagues, which they refer to as Index-calibration 122 experiment – (IC-Karanth). The authors have presented eight paired data points on tiger density 123 and tiger signs (in fact only scats) in figure 5 of the paper. This graph shows no relationship 124 between tiger scat encounter rate and tiger density, considered as an empirical test in support of



125	their theoretical model based only on eight data points. On perusal of the references cited in
126	Gopalaswamy et al (2015a), we noticed several irregularities which invalidate the use of these
127	data as a scientific experiment to test this relationship. It is relevant to point out that for
128	calibration of any index with abundance as done in a double sampling experimental approach
129	(Eberhardt & Simmons 1987), both index and abundance, should be sampled contemporaneously
130	and over the same spatial extent (paired in time and space). In three data points out of eight
131	presented in figure 5 of Gopalaswamy et al. (2015), tiger signs and tiger density were not
132	collected contemporaneously. Tiger density can fluctuate substantially between years (Karanth et
133	al. 2006) and tiger signs have short persistence time. Yet, the data Gopalaswamy et al. (2015) use
134	for their paired experiment has lags of several years (two to seven years) between estimating
135	tiger density and tiger sign (Figure 1). In particular, the data point from Bandipur has a lag of
136	seven years (density estimated in 1999, scat sampling in 2006), data point representing Melghat
137	has a lag of three years (density estimated in 2002, scat sampling done in 2005) and data point
138	from Pench Maharashtra has a lag of two years (density estimated in 2002, scat sampling done in
139	2004) (Karanth & Nichols 2000, 2002; Karanth et al. 2004; Karanth & Kumar 2005; Andheria
140	2006, see supplementary material for relevant sections of these publications). The authors do
141	have concurrent density estimates from one of these sites (Bandipur) with smaller variance
142	(Gopalswamy et al 2012), but curiously have not chosen to use or refer to this. At one data point
143	(Tadoba), an extreme outlier at right corner of figure 5 of Gopalaswamy et al (2015a) (Figure 1),
144	the data on scat encounters does not match the original source (scat encounter rate $3.6/10~\mathrm{km}$ as
145	given in figure 5 of Goplaswamy et al (2015) vs. 1.99/10 km as given in the original source
146	(Karanth & Kumar 2005; but addressed this by mentioning that the original reference was
147	incorrect in a corrigendum to the original paper Gopalaswamy et al (2015b)). Yet two data points
148	(Melghat and Pench Maharashtra) continue to differ in their Fig 5 (Gopalaswamy et al (2015a)
149	from the cited references in the corrigendum Gopalaswamy et al (2015b).
150	Methods for recording scat encounter rates differed between source reference sites used
151	for IC-Karanth. Andheria (2006) removed all scats encountered on the first sample and
152	discarded them from data analysis, a practice which is not uniformly followed for recording
153	tiger scat encounter rates in other studies. For studies referenced for IC-Karanth, camera
154	trap sampling was done in small areas within larger protected areas for estimating tiger
155	density, whereas tiger scats were collected for studying tiger diet (Karanth & Nichols 2000,



156 2002; Karanth et al. 2004; Karanth & Kumar 2005; Andheria 2006) possibly opportunistically from the entire reserve. Any intent of calibrating these tiger scat data to 157 158 tiger density obtained through camera trap sampling is not mentioned in any of the original 159 sources. In the original studies cited by Gopalaswamy et al. (2015) referred to as IC-160 Karanth experiment, there seems to be no intent of designing an experiment to evaluate the 161 relationship between tiger sign encounter rate and tiger density, the sources are unclear if 162 the scat sampling was done within the same spatial extent as the camera trap survey for 163 estimating tiger density. The basic premise of a double sampling experimental approach, 164 wherein data from both samples (index and density) need to be paired in time and space is 165 violated in the field experiment (IC-Karanth) of Goplaswamy et al (2015) invalidating their 166 conclusions. 167 [Figure 1. Here] 168 169 170 171 d) Variability in tiger capture probability and density estimates from camera trap capture-172 mark-recapture. As with occupancy detection probability, Gopalaswamy et al. (2015) restrict themselves entirely 173 to 11 estimates of tiger density published by Karanth et al (2004) for their models. On multiple 174 occasions they point out the highly variable capture probability p and variance associated with 175 176 tiger density estimates. In fact, in light of the large number of published tiger density estimates 177 with higher precision (e.g. 21 estimates in Jhala et al 2011a), these authors should have 178 considered Karanth et al (2004) estimates as particularly lacking in precision. When, estimates 179 with large sampling errors are used to guide development of theoretical models it would be 180 difficult to deduce any relationship between tiger signs and tiger density. Poor precision of tiger 181 density estimates in Karanth et al (2004) were likely due to poor sampling design and not 182 something that is inherent in tiger population estimation, e.g. for data presented in Karanth et al 183 (2004) CV of tiger density increases with increase in sampled area and p decreases with the 184 sampled area (r = 0.4, and -0.63 respectively). Overstating the case of sampling uncertainty can 185 only do harm to the development and adoption of sound and practical methods.



186 187 e) Repeating non peer-reviewed literature to advance unsubstantiated claims 188 Gopalaswamy et al. (2015) claim that the methods followed by Jhala et al. (2011a) have resulted 189 in "improbable estimates of 49% increase in tiger density over 4 years". Gopalaswamy et al 190 (2015a) do not explain how they arrived at the figure of 49% increase, they cite a letter to 191 Science, commenting on a news article (Karanth et al 2011b), but they have not explained the 192 49% increase in tiger abundance in this letter as well (Jhala et al. 2011c). The fact is that in 2006 193 India's mean tiger population was estimated at about 1400 while in 2010 the estimate was about 194 1700 but included estimates from some new areas like Sundarbans that were not assessed in 195 2006. Comparing tiger numbers between common areas sampled in 2006 and 2010 an increase 196 of 17.6% was estimated in four years, or about 4% per year; which is very probable for large 197 carnivores. It is inexplicable to us how Gopalaswamy et al (2015a) arrived at a 49% increase in 198 abundance or why they continue to perpetuate this obviously erroneous inference. 199 200 201 f) Propaganda that is not consistent with facts 202 The paper of Gopalaswamy et al (2015a) is, as the title suggests, about "index calibration 203 experiment" especially referring to estimation of tiger abundance. To this extent the reference to Jhala et al (2011a) that demonstrates a strong relation between tiger sign index and tiger 204 205 abundance as IC-Jhala and several publications of U. Karanth as IC-Karanth is relevant. 206 Gopalaswamy et al. (2015a) seem to have gone through the methods employed for estimating the 207 status of tigers in India thoroughly (Jhala et al. 2008, 2011b, and 2015), since they have 208 meticulously computed parameters from these reports for their paper. K. U. Karanth is also an 209 author on several chapters in Jhala et al (2015). They should know that national tiger status 210 assessments (Jhala et al 2008, 2010, 2015) were never based on tiger sign index alone. Tiger sign 211 index was one amongst the many ecologically important covariates that included human 212 footprint, prey abundance and landscape characteristics, that were used for modeling tiger 213 density. Yet, the blog of the journal Methods in Ecology and Evolution titled "flawed method puts tiger rise in doubt " states "amongst recent studies thought to be based on this method is 214 215 India's national tiger survey " (Grives 2015) which the blog then discredits as being inaccurate 216 based on conclusions of Gopalaswamy et al. (2015a). The fact is India's national tiger survey of



217	2014 (Jhala et al. 2015) used spatially explicit capture-recapture (SECR) in a joint likelihood
218	based framework (Efford 2012) with covariates of prey abundance, tiger sign intensity, habitat
219	characteristics, and human footprint. The SECR and Joint likelihood analysis are a recent
220	development (Brochers & Efford 2008, Efford 2011) and therefore could not have been used for
221	earlier national tiger assessments which used general linear models (Jhala et al 2008, 2011b).
222	The misleading reports that subsequently followed in the media had forgotten that the MEE
223	paper by Gopalaswamy et al. (2015a) is a debate on index calibration using double sampling
224	approach (Eberhardt & Simmons 1987) with simple linear regression and not about national tiger
225	status assessment. The 2014 national tiger status assessment was based on photo-captures of
226	1506 individual tigers, capture-histories of these were subsequently modeled in SECR with
227	covariates of prey, habitat and human impacts to estimate 2226 (SE range 1945 to 2491, >1.5
228	year old) tigers from across India (Jhala et al 2015). This amounts to 68% of the total tiger
229	population being photo-captured and 77% (1722; 95% CI 1573 to 2221 tigers) of the total tiger
230	population being estimated by capture-mark-recapture without any extrapolation using
231	covariates/indices. By muddling index calibration with the national tiger survey in the paper
232	(Gopalaswamy et al. 2015) and in all subsequent press releases and interviews Dr. Ullas Karanth
233	and coauthors incorrectly use the Gopalaswamy et al. (2015a) paper results (which are
234	themselves highly questionable) to discredit the national tiger survey results as being inaccurate
235	(Bagla 2016, Chauhan 2015, Croke 2015, Grives 2015, Karanth 2015, 2016, Rohit 2015, Sinha
236	and Bhattacharyal 2015, Vaughn 2015, and Vishnoi 2015) and mislead the readers.
237	Peer reviewed publications form the basis for advancement of science and are often cited and
238	used as a basis from which to move ahead. Indeed, the Gopalaswamy et al. (2015a,b) paper has
239	been subsequently cited in papers addressing methodological reviews, advances and policies
240	(Darimont et al 2018, Hayward et al. 2015), abundance estimation papers (Broekhuis and
241	Goplaswamy 2016, Caley 2015, Elliot and Gopalaswamy 2017, Falcy et al 2016, Mahard et al
242	2016) and in some Masters and Ph.D. thesis (Walker 2016, Moorcroft 2017). Published scientific
243	literature can have errors, these can occur through negligence of scientists or deliberate
244	misleading of science (Macilwain 2014), and can pass the peer review process due to ignorance,
245	poor diligence, or vested interest (Parsons & Wright 2015). Mistakes in published science
246	should be corrected expediently, as these are detrimental to the scientific progress in the specific
247	field and propagate a wrong basis for further research. In our opinion, Gopalaswamy et al.



248	(2015a, b) results are misleading, due to inappropriate scientific process and data, and have
249	therefore not contributed to the wider debate on the usefulness of double sampling (Eberhardt &
250	Simmons 1987; Pollock et al. 2002) for large-scale animal surveys.
251	We stress that landscape scale surveys need to be a blend of robust statistical design and analysis
252	that are pragmatic (economic and logistically possible) to achieve. The national tiger surveys of
253	India (Jhala et al. 2008, 2011, 2015) have striven to keep pace with modern advancement in
254	animal abundance techniques and analysis and have used robust statistical tools available within
255	the constraints of large-scale data coverage, resources, and timeframe. The concept and
256	philosophy of double sampling (Cochran 1977) forms the basis for modern statistical and
257	analytical approaches that infer relationships between actual abundance and counts, indices, and
258	covariates. The family of general linear models, generalized additive models (Zuur et al. 2009),
259	joint likelihood (Conroy et al. 2008), SECR with habitat covariates (Efford and Fewster 2013),
260	and SECR joint likelihood (Chandler and Clark 2014) take the relationship between an
261	index/covariates and absolute abundance to various levels of analytical complexity. There seems
262	to be some agreement on the best analytical approach to use for landscape scale abundance
263	estimation of tigers between Gopalaswamy et al. (2015a) and us (Jhala et al. 2015).
264	Gopalaswamy et al. (2015a) recommend using the joint likelihood approach, while the tiger
265	status assessment for India for the year 2014 used spatially explicit joint likelihood with camera-
266	trap data of tigers, and covariates of tiger sign index, prey abundance and human footprint
267	indices (Jhala et al. 2015). Yet, we stress the relevance and importance of first exploring
268	relationships of abundance with indices and covariates, based on sound ecological logic before
269	attempting complex statistical analysis, and refrain from putting the proverbial cart (statistical)
270	before the horse (ecology) (Krebs 1989).
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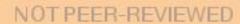
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482	Figure 1. Recreation of figure 5 from Gopalaswamy et al. (2015) highlighting the data
483	discrepancies in the index-calibration experiment. The names of tiger reserves from central
484	Indian landscape and Western Ghat landscape, where sampling was done are mentioned. MR
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Figure 1(on next page)

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