A peer-reviewed version of this preprint was published in PeerJ on 25 September 2018.

<u>View the peer-reviewed version</u> (peerj.com/articles/5677), which is the preferred citable publication unless you specifically need to cite this preprint.

Wang J, Yamada R. 2018. In silico study of medical decision-making for rare diseases: heterogeneity of decision-makers in a population improves overall benefit. PeerJ 6:e5677 https://doi.org/10.7717/peerj.5677

In silico study of medical decision-making for rare diseases: heterogeneity of decision-makers in a population improves overall benefit

Juan Wang¹, Ryo Yamada^{Corresp. 1}

¹ Unit of Statistical Genetics, Center for Genomic Medicine, Graduate School of Medicine, Kyoto University, Japan

Corresponding Author: Ryo Yamada Email address: statgenetjimu@gmail.com

Background: Medical decision-making is difficult when information is limited due to rareness. For example, there are two treatment options for patients affected by a rare disease with high lethality. The information about both treatment effects is unavailable or very limited. Patients are inclined to accept one of the interventions rather than waiting for death, but they are reluctant to be assigned the inferior one. While a single patient selects one treatment that seems better based on the limited information, he or she loses the chance to select the other treatment, which may be the better option. This is the socalled dilemma between exploitation (enjoying the benefits of using current knowledge) and exploration (taking the risk to obtain new knowledge). In clinical settings, the statistical advice for individual patients seems to be the maximum expected success rate or something equivalent and patients' selections tend to be homogeneous, which does not solve the dilemma. In this study, our aim is to investigate the effects of the heterogeneity of decision-makers in the decision process.

Methods: Here, we proposed a decision strategy that introduced the heterogeneity of decision-makers by considering patients' self-decisions where the patients' heterogeneous attitudes towards the treatment are integrated into the probabilistic utility function based on the Beta Bayesian posterior. Based on the context of two-armed bandit treatment options with limited information, we compared the overall success rate of treatment between our heterogeneous decision strategy and a homogeneous decision strategy that is defined to select the treatment with the largest posterior mean.

Results: The heterogeneity of decision-makers in a population improved the overall benefit of treatment under some conditions.

Discussion: In clinical settings, there exists heterogeneity of decision-making among patients. Our study investigated a targeting strategy by respecting the self-decision of all individuals and found that the heterogeneity of decision-making can improve the overall benefit under some conditions. In addition, this outperformance may suggest that heterogeneity of decision-making is of importance to human beings. Besides the ethical merit, our findings provide meaningful ideas for better strategies towards decision-making dilemmas in clinical settings for rare diseases or cases where only limited information is available. It would be further suggested to investigate the effects of heterogeneity of decision-making in other fashions, such as genetic heterogeneity and phenotypic heterogeneity.

1	In silico study of medical
2	decision-making for rare diseases: heterogeneity of decision-makers in a
3	population improves overall benefit
4	
5	Juan Wang, Ryo Yamada*
6	
7	Unit of Statistical Genetics, Center for Genomic Medicine, Graduate School of Medicine Kyoto
8	University, Kyoto, Japan
9	
10	* Corresponding author
11	E-mail: ryamada@genome.med.kyoto-u.ac.jp (RY)
12	
13	
14	
15	
16	
17	
18	
19	
20	

21 Abstract

22 **Background:** Medical decision-making is difficult when information is limited due to rareness. 23 For example, there are two treatment options for patients affected by a rare disease with high 24 lethality. The information about both treatment effects is unavailable or very limited. Patients are 25 inclined to accept one of the interventions rather than waiting for death, but they are reluctant to 26 be assigned the inferior one. While a single patient selects one treatment that seems better based 27 on the limited information, he or she loses the chance to select the other treatment, which may be 28 the better option. This is the so-called dilemma between exploitation (enjoying the benefits of 29 using current knowledge) and exploration (taking the risk to obtain new knowledge). In clinical 30 settings, the statistical advice for individual patients seems to be the maximum expected success 31 rate or something equivalent and patients' selections tend to be homogeneous, which does not 32 solve the dilemma. In this study, our aim is to investigate the effects of the heterogeneity of 33 decision-makers in the decision process.

Methods: Here, we proposed a decision strategy that introduced the heterogeneity of decisionmakers by considering patients' self-decisions where the patients' heterogeneous attitudes towards the treatment are integrated into the probabilistic utility function based on the Beta Bayesian posterior. Based on the context of two-armed bandit treatment options with limited information, we compared the overall success rate of treatment between our heterogeneous decision strategy and a homogeneous decision strategy that is defined to select the treatment with the largest posterior mean.

41 **Results:** The heterogeneity of decision-makers in a population improved the overall benefit of
42 treatment under some conditions.

43 Discussion: In clinical settings, there exists heterogeneity of decision-making among patients. Our 44 study investigated a targeting strategy by respecting the self-decision of all individuals and found that the heterogeneity of decision-making can improve the overall benefit under some conditions. 45 46 In addition, this outperformance may suggest that heterogeneity of decision-making is of 47 importance to human beings. Besides the ethical merit, our findings provide meaningful ideas for 48 better strategies towards decision-making dilemmas in clinical settings for rare diseases or cases 49 where only limited information is available. It would be further suggested to investigate the effects 50 of heterogeneity of decision-making in other fashions, such as genetic heterogeneity and 51 phenotypic heterogeneity.

53 Introduction

54 The Randomized Controlled Trial (RCT) is considered to be the gold standard for the evaluation of treatment effects in medical settings. Since the design that combines randomization 55 56 and blinding in the RCT minimizes selection bias and distributes confounders between the placebo 57 group and the intervention group, the assessment of outcomes is more objective and accurate [1,2]. 58 While the classic RCT design is suitable for common diseases, it is less feasible in rare diseases 59 because it is too time consuming to obtain a sufficient sample size [3-5]. Moreover, the outbreak 60 of acute lethal diseases requires patients to select one option among few with limited information. 61 A typical example of this was the Ebola infection outbreak of 2013-2016, when empirical 62 treatments were offered to patients without RCT. Parents may be reluctant to enrol their child in a 63 placebo-controlled trial where he or she may receive a placebo rather than undergo the intervention 64 of a treatment [6]. In fact, there are precedents for approval of orphan drugs treating rare 65 neurological diseases based only on pilot studies using smaller trial sizes and without the RCT 66 principles [7].

67 Here, we assume a medical condition in practice. Two new treatments are developed for a rare disease with a high lethality rate; however, there is no or very limited evidence for both 68 69 treatments' effect in the current status. Under these conditions, patients themselves are disposed 70 to try one of the two treatments rather than waiting for death. As we do not know the true effects 71 (the true success rate) of either treatment, the very first patient will select either treatment with 72 probability 0.5, that means she or he is assigned to the treatment with a lower success rate with probability 0.5. With the increase of sample size, strong evidence would eventually be obtained to 73 74 clarify which treatment is superior to the other. However, in the process, many of the benefits to 75 the patient are sacrificed, in particular individuals who are enrolled in the earlier stages of the trial.

Although we investigated the effects of heterogeneity of decision makings, it is important to state that we do not mean the heterogeneity-oriented approach should replace the well-established approaches for clinical trials where random assignment is critical to assure the results of the studies. Our objective is to study the heterogeneity of decision makings and its effects in more general settings.

81 One of the classic Bayesian decision rules is to take the treatment with the largest posterior 82 maximum expected success rate encoded in the posterior distribution. Then, patients are assigned to the treatment with the largest posterior maximum expected success rate [8,9]. However, this 83 84 causes a problem. While a single patient is assigned to one out of two treatments with a higher 85 posterior mean, it prevents exploration of the other treatment, which may be the true superior. 86 Thus, the problem is the dilemma between exploitation (enjoying the benefit of using the current 87 knowledge) and exploration (taking the risk of obtaining new knowledge) [10]. Many studies in 88 statistical decision theory have revealed that the validity of the loss function-based approach can 89 quantitatively demonstrate the optimal decision strategy and its limitation when information is 90 limited [11,12]. In addition to the statistical decision theory, there is another important aspect in 91 clinical decision-making by patients. As we clinicians see patients, we find them heterogeneous 92 regarding selection among options. Some aspects of the heterogeneity seem to be explained by 93 inadequate understanding of statistical information on the options. Some aspects of the 94 heterogeneity seem to be explained by the heterogeneity of personalized conditions or weighting 95 parameters of loss functions. In addition, however, there seems to be heterogeneity in risk-taking. One such example is the attitude towards clinical trials. Some patients are very positive towards 96 97 them, and some decline the idea, while others waver somewhere between. We considered that the

risk-takers tend to see the optimistic aspects of the unknowns and the risk-hesitaters tend to seethe pessimistic aspects.

100 Because we were interested in whether patient optimism/pessimism heterogeneity could 101 mitigate the problem of the exploration-exploitation dilemma, we designed this study to evaluate 102 the effects of heterogeneity of decision making on the overall success rate of treatments in a 103 population. In the study, we generated a simple decision-making model with the heterogeneity. 104 The heterogeneity of decision-making was parameterized for optimism/pessimism, called as the 105 targeting decision strategy (T-strategy). With the Bayesian decision-theoretic approach [12], 106 patients' belief in the current state of knowledge regarding the success rate of each treatment is 107 estimated from the Beta posterior distribution. Then, patients taking the T-strategy will select the 108 treatment with the higher posterior probability that success rate is more than a targeting value. This 109 value is calculated from a function of the larger posterior mean of success rate, but depending on 110 the patient's attitude. To evaluate the effects of T-strategy, we compared it with one classic 111 decision rule which is defined as selecting the treatment with the larger posterior mean only, called 112 as E-strategy. Using simulated datasets with two new treatments options, we compared the overall 113 success rate of two treatments based on some conditions between the patients who are taking T-114 strategy and E-strategy separately, and quantitated the effects of the heterogeneity of decision-115 making.

116 Materials & Methods

117 Context

In this paper, we study the two-armed-bandits problem in clinical settings in which there are two new treatments for a disease without knowledge of success, denoted *A* and *B*. Each patient's Bernoulli outcome, favourable (success) or unfavourable (failure), is recorded. The true

rates of favourable outcomes for *A* and *B* are unknown, and they are denoted *a* and *b*, with 0 < a, b < 1. A series of patients select *A* or *B*, one by one, and the next $(n + 1)^{th}$ patient is informed with the preceding outcomes that *n* patients have been treated in total and n_A and n_B have selected *A* and *B*, respectively, with n_{As} and n_{Bs} successful outcomes and n_{Af} and N_{Bf} failures, respectively. As shown in Table 1, $n = n_A + n_B$, $n_A = n_{As} + n_{Af}$, $n_B = n_{Bs} + n_{Bf}$, $n_s = n_{As} + n_{Bs}$, and $n_f = n_{Af} + n_{Bf}$.

127 Beta conjugate distribution to binomial outcomes

The outcome of each treatment is a Bernoulli outcome, and the parameter of unknown success rate follows a binomial distribution. In Bayes' theorem, whereby the posterior is proportional to the prior multiplied by likelihood, there is one advantage in that the beta distribution is the conjugate distribution to Binomial outcomes (shown as supplementary note 1). When no patient has been treated, we set a uniform Beta(1,1) as the initial prior, and then the two parameters of the prior are to be updated by the outcomes, successes $n_{*s} + 1$ and failures $n_{*f} + 1$.

134
$$p_{*}(\theta|n_{*s}, n_{*f}) = \beta \left(n_{*s} + 1, n_{*f} + 1\right) = \frac{1}{B(n_{*s} + 1, n_{*f} + 1)} \theta^{n_{*s}} (1 - \theta)^{n_{*f}}, \qquad (1)$$

135 where * indicates *A* or *B*, β and B indicate the beta distribution and function, respectively, and 0 136 $\leq \theta \leq 1$.

138 of decision-making

We assumed that every individual selects one out of two treatments (*A* or *B*) with a higher value that is estimated from the beta posterior distribution with given current outcomes (successes and failures) of each treatment, $p_*(\theta | n_{*s}, n_{*f})$ (Equation 1). In this study, we modelled two types

(2)

(3)

142 of individuals; one type of individual selects a treatment based on the posterior mean v_{*E} 143 (Equation 2). They select the treatment with the larger posterior mean/larger maximum expected 144 success rate (Equation 5). We call this type of individuals' selection as the E-strategy, where E 145 stands for "Expected". The other type of individual is somehow optimistic or pessimistic and 146 selects the treatment with a higher value that is different from the expected value and depends on 147 each individual's optimistic/pessimistic preference. We set an attitude index, w, to parameterize 148 the two preferences of individuals, where the w of pessimistic individuals ranges from -1 to 0 149 and the optimistic w ranges from 0 to 1. In fact, in terms of treatment assignment in clinical 150 settings, we assumed that optimistic individuals care whether the treatments are adequately 151 successful or not and that they set a target value (t) higher than the maximum posterior mean (v_{*E}) (Equation 4), and calculate the probability that success rate is higher than the t. This 152 probability is denoted by $v_{*T}(t)$ (Equation 3). By contrast, we assumed that pessimistic 153 individuals set t lower than the maximum posterior mean(v_{*E}), and calculate the probability that 154 155 success rate is higher than t. The modelled t is calculated from a function of maximum posterior mean (v_{maxE}) but depending on this individual's attitude index w (Equation 4). Those individuals 156 who are optimistic or pessimistic select the treatment with higher v_{*T} (Equation 5). We refer to 157 158 this type of individual's decision as a T-strategy, where T stands for "Target".

159 The posterior mean of the success rate is

160
$$v_{*E} = \int_0^1 \theta \times p_*(\theta | n_{*S'} n_{*f}) d\theta = \frac{n_{*S} + 1}{n_{*S} + n_{*f} + 2},$$

- 161 where * indicates A or B
- 162 The probability of a success rate more than a target value is
- 163 $v_{*T}(t) = \int_t^1 \mathbf{p}_*(\theta | n_{*S}, n_{*f}) d\theta,$

164 where * indicates A or B, and the target value t is calculated by $v_{maxE} = max(v_{AE}, v_{BE})$ and

165 depends on the attitude index *w*.

166
$$t = \begin{cases} w + (1 - w)v_{maxE}, & 0 \le w \le 1\\ (1 + w)v_{maxE}, & -1 \le w < 0 \end{cases}$$
(4)

167 where positive and negative *w* stands for optimism and pessimism, respectively. Correspondingly, 168 the $v_{maxE} \le t \le w$ specifies optimism, and $0 \le t \le v_{maxE}$ specifies the pessimism.

169 Subsequently, the probability of selecting A, Prob(A), is given as

170
$$Prob(A) = \begin{cases} 1: & v_{A*} > v_{B*} \\ 0.5: & v_{A*} = v_{B*} \\ 0: & v_{A*} < v_{B*} \end{cases}$$
(5)

where * indicates E or *T*. Actually the selection of every individual is deterministic based on the values that are calculated from the 2 by 2 table values in principle. The selection is stochastic only when the values are equal. When we assumed a population was homogeneous, their selection was deterministic except for the stochastic selection due to the identical values for two arms. When we assumed a population was heterogeneous, the individuals' optimistic/pessimistic attitude vary among them and the sequence of individuals were stochastically generated in the experiment.

We assumed that the population is a mixture of individuals with various levels of optimism/pessimism. To specify the heterogeneous population in this model, we assumed that need w is symmetric around zero and is in a monomodal distribution ranging from -1 to 1. With this assumption, the majority of people are almost neutral and relatively few people are strongly optimistic or pessimistic. As a simple model for this distribution, we assumed

182
$$w \sim 2 (\beta(u,u) - 0.5),$$
 (6)

183 where *u* parameterizes the shape of distribution of *w*.

184 An example for the selection was given in the Figure 1, where the Beta posterior 185 distributions of two treatments were drawn with the information: the eighteen (18 = (14 - 1) + (6 - 1)) patients have been treated with the treatment A, resulting in 186 187 outcomes of 13 successes and 5 failures, and three (3 = (3 - 2) + (2 - 1)) patients have been 188 treated with B, resulting in outcomes of 2 successes and 1 failure. Two sets of decision values 189 based on E-strategy (v_{*E}) and T-strategy (v_{*T}) were calculated separately, and in the Figure 1 v_{*E} are indicated as vertical lines ($v_{AE} = 0.7$ and $v_{BE} = 0.6$), and v_{*T} are indicated as the 190 191 area under the curve truncated by a vertical line ($v_{AT}(t) = 0.16$ and $v_{BT}(t) = 0.18$). Since 0.7 > 0.6, individuals with E-strategy should select A. Since 0.16 < 0.18, individuals with T-192 193 strategy (t = 0.8 corresponds to an optimistic individual with w approximately 0.3) should 194 select *B*.

195 **Experimental conditions**

196 The true but unknown success rate of the two arms, a and b, and the number of total patients 197 , N, are parameterized. The total 5050 pairs of combination of a and b was generated with $0 < b \le a < 1$. For the T-strategy population, w as an attitude index is parameterized within the 198 199 range of -1 and 1. In addition, we evaluated relatively small N values, from 1 to 100, because 200 our objective is to study the effects of optimism/pessimism of decision-makers for treatment 201 assignment with a rarer disease in a clinical setting where the patient population size is small. 202 Under the same condition of a, b, and N, we first compared the overall success rate between the 203 homogeneous population who take the E-strategy and the homogeneous population who take the 204 T-strategy, where all T-strategy individuals have the same w values.

After the comparison between the homogeneous E-strategy population and the homogeneous T-strategy population, we investigated the effects of heterogeneity. We generated a

207 population with the T-strategy whose *w* values were not same but distributed as shown in Equation

208 6, where *u* was 5, 30, 70 and 500.

209 Measure of the overall benefit

210 Calculation of the exact probability of every Bernoulli outcome of the two arms

211 In this study, we enumerated all possible combinations of successes and failures for each

- 212 treatment as a 2 \times 2 table (shown as supplementary note 2); then, the exact probability of every
- 213 table consisting of N_{As} , N_{Af} , N_{Bs} , and N_{Bf} was calculated as the below equation.

214
$$Pr(N_{As}, N_{Af}, N_{Bs}, N_{Bf}) = Prob(A | (N_{As} - 1, N_{Af}, N_{Bs}, N_{Bf})) \times a \times Pr(N_{As} - 1, N_{Af}, N_{Bs}, N_{Bf}) + Prob(A | (N_{As}, N_{Af} - 1, N_{Bs}, N_{Bf})) \times (1 - a) \times Pr(N_{As}, N_{Af} - 1, N_{Bs}, N_{Bf}) + (1 - Prob(A | (N_{As}, N_{Af}, N_{Bs} - 1, N_{Bf}))) \times b \times Pr(N_{As}, N_{Af}, N_{Bs} - 1, N_{Bf})$$

215 +
$$(1 - Prob(A|(N_{As'}, N_{Af'}, N_{Bs'}, N_{Bf} - 1))) \times (1 - b) \times Pr(N_{As'}, N_{Af'}, N_{Bs'}, N_{Bf} - 1),$$
 (7)

216 When patient number N=0, then
$$Pr(0,0,0,0) = 1$$
.

For heterogeneous decision-makers, we assigned w to a series of patients from the distribution in the equation (6) with the indicated u value. Because the stochastic processes vary with the sequence of w values, we iterated 3000 random Monte Carlo patient sequences [13] up to N = 50, and we calculated the average of $Pr(N_{As}, N_{Af}, N_{Bs}, N_{Bf})$.

221 Measure of the Overall Success Rate

We emphasized the evaluation of the overall benefit of treatments in a population rather than an individual's best benefit; thus, we measured the average fraction of favourable outcomes (successes) for the series of N patients as a whole when a and b were given regardless of the selected arm, named the Overall Success Rate (OSR).

226
$$OSR(N) = \sum_{\{(N_{AS'}, N_{Af'}, N_{BS'}, N_{Bf}) | N_{AS} + N_{Af} + N_{BS} + N_{Bf} = N\}} \sum_{N=1}^{N_{AS} + N_{BS}} Pr(N_{AS'}, N_{Af'}, N_{BS'}, N_{Bf}), \quad (8)$$

All calculation was performed with the R, and the code is available at the following URL:

228 <u>https://github.com/statgenetJimu/SelfDecABP/blob/master/SelfDecABP(1).%20r-package</u>.

229

230 **Results**

First, we showed a typical case of homogeneous E-strategy (*E.st*) and homogeneous Tstrategy (*T.st*). Second, we demonstrated the detailed effects of conditions of *a*, *b* and *N* on the difference between homogeneous *E.st* and homogeneous *T.st*. Third, we showed the benefit of the heterogeneous *T.st*.

235 Typical case of homogeneous E-strategy (*E.st*) and homogeneous T-

236 strategy (T.st).

237 Individuals in the homogeneous *E.st* population select treatments based on the posterior mean, v_{*E} . Individuals in the homogeneous T.st population select treatments based on v_{*T} , 238 239 which represents the optimism/pessimism attitude and is shared by all the individuals in the 240 population. Figure 2 is the results of the experiment, where the success rate of the two treatments 241 were a = 0.8 and b = 0.6, and the optimistic attitude w of the T.st population was 0.5. The total 242 number of patients was N = 100. The panels A and B of Figure 2 are the 2-dimensional 243 histograms, where one axis is the fraction of individuals who selected A and the other axis is the 244 overall success rate when all the processes reached N = 100 for *E.st* and *T.st*, respectively. The 245 processes of selection by a series of individuals are stochastic; the fraction of selecting A and the 246 overall success rate take distribution. The exact distribution of the fraction and the rate were 247 calculated and displayed. In the case of *E.st*, as shown in panel A of Figure 2, the distribution was 248 bimodal, with the higher peak corresponding to the occasions in which a majority of N patients

249 had selected the better treatment arm, A, and the lower peak indicating that the minority had 250 selected the inferior treatment arm, B, with a subsequently lower OSR. In the case of T.st, as 251 shown in panel B of Figure 2, the distribution was monomodal with the peak towards the A-arm 252 selection. The bimodality was the result of the exploitation-exploration dilemma. In some cases, the patients who selected the inferior one turn out to be successful with the expected success rate 253 254 higher than the true success rate because this is a stochastic process. In this case, the following patients tend to select the inferior treatment arm with the belief that this treatment arm has a high 255 256 success rate, and they lose the chance to select the other treatment arm that was truly better. Panels 257 A and B indicate that the decision strategy of the population affected the exploitation-exploration 258 pattern.

Panels C and D in Figure 2 show the average fraction of individuals who selected the *A*arm and the OSRs among a number of patients from 1 to 100 of the two strategies, respectively. Panel C shows that the A-arm fraction of *E.st* and *T.st* at N = 100 was 0.714 and 0.855, respectively, and the fraction was higher for *T.st* throughout for the number of patients. Panel D shows that OSR of *E.st* and *T.st* at N = 100 was 0.743 and 0.771, respectively, and the rate was higher for *T.st* throughout. This finding was also related to the lack of exploration that occurred in *E.st* in this particular scenario.

266 Comparison of overall benefit of homogeneous decision-makers 267 between E-strategy (*E.st*) and T-strategy (*T.st*) with the same

268 attitude index w value

The typical case above showed that the homogeneous decision-makers of *T.st* with w = 0.5outperformed *E.st* when a = 0.8 and b = 0.6 for N = 1, 2, ..., 100. However, such superiority

is not always true for all conditions of *a*, *b*, *N*, and *w*. In fact, under some conditions, homogeneous *T.st* decision-makers outperformed the homogeneous *E.st*, but under other conditions, the *E.st*decision-makers outperformed the *T.st*.

We evaluated the difference of Overall Success Rates (OSRs) between homogeneous *T.st* and homogeneous *E.st* decision-makers with the same attitudes of *w* values for various conditions; N = 1, 2, ..., 100 and $(a,b) = \{(a,b) | a,b \in \{0.01, 0.02, ..., 0.99\}, a \ge b\}, w = \{-0.0005, -0.001, ..., 0.999, 0.9995\}$. The number of (a,b) pairs was 5050.

Panel A of Figure 3 visualizes the difference of OSR of 5050 (*a,b*) pairs at N = 10,30, and 100 for w = -0.8, -0.4, 0.4, and 0.8. The (*a,b*) pairs form a triangular space. The horizontal axis is *a*, and the vertical axis is *b*. Red indicates (*a,b*) pairs where *T.st* outperforms *E.st*, and blue indicates (*a,b*) pairs where *E.st* outperforms *T.st*. Colour intensity stands for the value of the difference of OSRs as indicated in the colour bar on the right. The black curves stand for the (*a,b*) pairs without difference between *E.st* and *T.st*.

The colour patterns of 12 conditions of the panel A of Figure 3 show that the superiority 284 285 of two strategies is the function of (a,b) conditional to N and w. The homogeneous optimistic 286 decision-makers (*T.st* with w > 0) outperformed *E.st* when both a and b were relatively large, 287 but their performance was worse than *E.st* when both a and b were relatively small, as shown in 288 the first and second rows of w = 0.4 and 0.8 of panel A of Figure 3. When N increases, the colour 289 intensity tends to become stronger. In contrast, the performance of the homogeneously pessimistic 290 attitude decision-makers (T.st with w < 0) was worse than E.st for the majority of (a,b) pairs, 291 and *T.st* outperformed only when both a and b were small. When N increases, the colour intensity 292 tends to become stronger.

293 Next, we investigated the relation between (a,b) pairs and the superiority of E.st and T.st, 294 regardless of the intensity of optimism or regardless of w values as far as w > 0. We calculated 295 OSRs for $w = \{0, \dots, 0.999, 0.9995\}$. Panel B of Figure 3 coloured the (a,b) pairs with red, grey 296 and blue, where red indicates that the average OSR of T.st is higher than the OSR of E.st for all 297 optimistic w values, and blue indicates that the average OSR of T.st is lower than the OSR of E.st 298 for all optimistic w values, and grey indicates otherwise. In general, T.st tends to outperform when 299 both treatments have a relatively high success rate and *E.st* outperforms when both treatments 300 have a relatively low success rate. In this comparison, we evaluated homogeneous T.st populations 301 from which all individuals in a population were the same w values, and we set different such same 302 w (ranging from -1 to 1) values for each of homogenous populations. In the next experiment, 303 we modelled populations that are heterogeneous for decision attitudes that consisted of individuals 304 whose optimism/pessimism attitude index w varied and compared their performance with E.st.

305 The effect of heterogeneity of decision-makers in a population

In reality, populations seem to consist of individuals with various attitudes. We modelled heterogeneity of optimism/pessimism attitude index w with equation 6, where w distributes symmetric around 0. Figure 4 is the colour plot to display the superiority of *E.st* and heterogeneous *T.st*. Four panels show the result of four different distributions of optimism/pessimism index w, specified with u = 5, 30, 70 and 500. The distribution of w is displayed in a window of each panel.

The panels of Figure 4 show some tendencies. The heterogeneous T.st tended to outperform *E.st* (red) when both *a* and *b* were relatively high. With smaller *u* or larger variance of *w*, the difference of OSR was bigger. Actually, the right most panel indicates essentially no difference between *E.st* and heterogeneous *T.st* when the variance of *w* is very small. When *u* is small (u = 5), the triangular area was divided into two coloured subareas almost evenly, and when u is larger (u = 30 and 70), the area where the heterogeneous *T.st* was superior was bigger than the area where *E.st* was better.

319

320 **Discussion**

321 We assumed that decision-makers might vary in a population due to the heterogeneity of 322 individuals' decision attitudes in medical decision-making. In reality, it is ethical to respect 323 individuals' self-decision particularly when there is no sufficient information to make decisions 324 for certain. In clinical settings, these conditions correspond to rare diseases or patients with a 325 common disease that is complicated by various conditions, where some information is available 326 but is not conclusive. If all individuals of a population are E-strategy decision-makers, the outcome 327 for the whole population may not be optimal, which is the result of the trade-off between 328 exploitation and exploration where an individual selects one arm to optimize the outcome for 329 herself/himself but the population loses the chance to take another arm that might be better; this is 330 the exploration-exploitation dilemma and the multi-armed bandit problem [10,14-16]. This 331 phenomenon was shown in Figure 2. Although statistical studies on the multi-armed bandit 332 problem developed decisions to optimize the outcome of the whole population strategies, such as 333 Gittin's index [17-19], they did not respect individuals' self-decisions. Thus, we studied the effects 334 of heterogeneity of decision-making on the overall benefit of the whole population in a simple 335 clinical setting. Although we stated in the introduction section, it should be stated again that we 336 have investigated the possible benefit in heterogeneity of decision makings in population but that 337 we do not mean that patients should select a treatment from multiple options and the importance 338 of clinical trials and the evidence-based medicine is the gold standard in clinics.

339 The question here is whether we, as humans, are truly heterogeneous in our selections. It 340 is true the individuals' background heterogeneity, such as comorbidity, cost, life-style and age and 341 so on, in case of clinical conditions, can cause their decisions heterogeneous and these hidden 342 factors might explain all the components of heterogeneity of decision making among people. 343 However it is also true that we, clinicians, face the heterogeneous attitudes among patients that do 344 not seem to be well explained by the factors. The particular example of this heterogeneity is the attitude heterogeneity towards clinical trials. We recruit patients for a clinical study who are 345 346 homogeneous enough to meet inclusion/exclusion criteria, and some patients participate in it and 347 others decline the idea. We believe the enrollment or no-enrollment should not be heavily biased 348 with hidden factors and we are not sure why some participate and some not. We investigated this 349 unclear heterogeneity in our study. Another example of the heterogeneity of decision makings 350 among people can be seen in attitude towards gambles. It is obvious that nobody should anticipate 351 gain as a whole, but still some people keep gambling and some won't. Therefore the assumption 352 of heterogeneity of decision making seemed reasonable to be investigated. In this study we 353 modelled the heterogeneity in decision makings very simply. When two treatments have never 354 been applied and someone is the first patient to be treated with either of one of the two, he or she 355 will select either treatment with a probability of 0.5. When one treatment has been used 18 times 356 with 13 favourable and five unfavourable outcomes and the other treatment has been subjected to 357 3 attempts with two favourable and one unfavourable outcome, which treatment would be 358 selected? If you ask this question to many people without further information, their answers would 359 vary. If you add the information with the knowledge of the posterior mean and the two treatments 360 are (13 + 1)/(13 + 1 + 5 + 1) = 14/20 = 0.7 and (2 + 1)/(2 + 1 + 1 + 1) = 3/5 = 0.6, 361 then all or the vast majority would likely select the first treatment (shown in the red and green

362 vertical lines, respectively in Figure 1). In other words, some people who may initially prefer the 363 second treatment with the information of (2 out of 3) have to change their preference to select the first treatment with the information (13 out of 18) if depending only on the posterior mean. Is this 364 change in preference due to the lack of "statistical literacy"? The answer could be yes or no. Given 365 the information of 13 out of 18 of the first treatment, the distribution of the favourable outcome 366 367 probability is a beta distribution with shape parameters 14 and 6 whose posterior mean is 0.7. 368 Using the same distribution, the probability that the first treatment could have a success rate of more than 0.8 is 0.16 (shown in the grey areas under the red and green curve separately in the 369 370 Figure 1). Based on the outcome information of two treatments, we have two sets of values. One 371 set is the expected success probability of two treatments, (the first treatment 0.7, the second 372 treatment 0.6). The other set is the probability that success probability should be more than 0.8373 (the first treatment 0.16, the second treatment 0.18). If we select the treatment with higher 374 expected probability, we should use the value set, (the first treatment 0.7, the second treatment 0.6), and because 0.7 > 0.6, we should take the first treatment. If we select the treatment with higher 375 376 probability that the first treatment could have a success rate of more than 0.8, we should use the 377 value set, (the first treatment 0.16, the second treatment 0.18), and because 0.18 > 0.16, we should take the second treatment. Based on this hypothetical evaluation, it can be said that the individuals 378 379 who selected the treatment with 2 vs. 1 might bet on it because of its potential to be a truly good 380 treatment. We call this attitude "optimistic". In fact, when we design a clinical trial to test a newer 381 treatment against a standard treatment, we should be optimistic enough to believe that there is 382 some chance that the newer treatment might be better than the standard treatment. Based on this 383 idea, we modelled a decision attitude, the T-strategy, which compared the likelihood that the

384 success rates are higher than the targeted value that reflects optimism/pessimism and their 385 intensity.

386 Through the comparison of the overall benefit in a population between the homogenous 387 decision-makers of E-strategies and homogeneous decision-makers of T-strategies with fixed 388 optimism/pessimism parameters, our study revealed the following: First, the optimistic 389 homogeneous decision-makers of the T-strategy outperformed the E-strategy when the true 390 success rates of both arms were relatively high, and the pessimistic homogeneous decision-makers 391 of T-strategy performed best when the true success rates of both were relatively low. Second, the 392 effects of optimism and pessimism were asymmetric. The area of the optimistic homogeneous 393 decision-makers of T.st with better performance than that of E.st tended to be wider when 394 compared with the pessimistic homogeneous decision-maker of T.st with better performance than 395 that of *E.st*. Additionally, homogeneous optimism worked better regardless of the intensity of the 396 optimism for some conditions, but homogeneous pessimism did not have any such conditions 397 (Figure 3). Furthermore, through the comparison of the overall benefit in a population between the homogenous decision-makers of E-strategies and the heterogeneous decision-makers with a 398 symmetric mixture of optimists and pessimists of various intensities, our study revealed that the 399 400 heterogeneous decision-makers in a population outperformed the homogeneous E-strategy when 401 the true success rates of the two arms were relatively high. When the variation in 402 optimism/pessimism was small, the degree of benefit was small, but the conditions a and b in 403 which heterogeneity outperformed the homogeneity of the E-strategy decision-maker were wide 404 and vice versa (Figure 4). These findings suggested that when the success rates of two treatment 405 arms for patients with rare diseases are believed to be relatively high, the decision-makers with an 406 optimistic decision attitude in a population would be the best. In addition, due to the wider

407 conditions of beneficial effects of the homogeneous optimistic than pessimistic attitudes, optimism 408 should be encouraged over pessimism if the decision-makers are homogeneous. When the 409 decision-makers in a population are heterogeneous, a large variation performs better under 410 narrower conditions with stronger intension. Those findings are practical not only in medical 411 decision-making but also in other fields of decision-making. For example, in a complex system in 412 network science, recent studies in the field have reported that the heterogeneity of factors increased 413 the structural vulnerability of the system [20]. In the context of our study, the sequence of 414 individuals can be considered to be a directed line graph in which information flows and the 415 heterogeneity of the factors of the network is attributed to the heterogeneity of each individual.

416 In our study, we evaluated only a small population size by enumerating all combinations 417 of outcomes of each treatment at each status by calculating the exact probabilities, each of which 418 is formed as a 2 \times 2 table pattern consisting of four integer numbers. For the larger population 419 sizes, we evaluated the cases with N = 500 using the Monte-Carlo simulation methods [13] rather 420 than the exact probability calculation, which showed a qualitatively similar phenomena to the ones 421 we observed for the smaller size. Because our investigation was limited to a very specific scenario 422 and an artificial attitude model, further studies should be performed; the following seemed to be 423 hypothesized. The benefit of a single individual will be maximized by selecting the option with 424 the higher expected success rate. However, when all the individuals of a population take the same 425 decision strategy, the overall benefit of the whole population may not be optimized in some cases. 426 This is the phenomenon of the exploitation-exploration dilemma. When the individuals of the 427 population are heterogeneous regarding decision-making, the dilemma seemed to be mitigated at 428 least partially. In addition, this idea is compatible with the clinical scenario where patients' self-429 decisions should be respected. Although we do not know whether human beings are heterogeneous

430 in decision-making, it is possible because human beings are heterogeneous in many ways, such as 431 genetically and phenotypically, and because the heterogeneity in various aspects is believed to be 432 important for the sustainability of the species. One more interesting finding was that the optimistic 433 attitude and the heterogeneity of optimism/pessimism performed better when both options had a 434 higher success rate. Because all species including human beings participate in the survival game 435 of evolutional history, they keep trying to find ways with higher success rates. Thus, it may be the 436 case that the majority of selection tasks are selections among options with relatively higher success 437 rates. If all of these assumptions are true, heterogeneity with some inclination towards the 438 optimistic side could be one of the best strategies for populations. Again, these hypotheses were 439 based on our limited investigations, and further studies are necessary.

Although the overall benefit of treating is improved if the heterogeneity of decision-making in a population is considered by respecting every individual's decision attitude, in fact, other realistic factors might be combined in our proposed heterogeneity of decision-making, e.g., cost, life-style and age. Considering that there are far more complicated cases in real clinical works, it would require effective cooperation between statisticians and clinicians for further investigation when more factors are introduced into the heterogeneity model.

446

447 **Conclusions**

We modelled the heterogeneity of decision-making in populations in terms of optimism and pessimism and compared them with the decision rule based on the expected success rate. We identified that the optimistic or pessimistic strategy outperforms the expected value-based strategy when success rates of options are in particular conditions. In addition, when a population consists of individuals with heterogeneous optimistic/pessimistic attitudes, it was able to outperform when

453 it pursues options with a high success rate. This outperformance is achieved by respecting the self-454 decision of all individuals, which is ethically important. Our findings may provide meaningful 455 ways to find better strategies for the decision-making dilemma in clinical settings for rare diseases 456 or cases where only limited information is available. It would be further suggested to investigate 457 the effects of heterogeneity of decision-making in other aspects, such as genetic heterogeneity and 458 phenotypic heterogeneity.

- 459
- 460
- 461

462 **References**

- Abel U, Koch A. 1999. The role of randomization in clinical studies: myths and beliefs. J
 Clin Epidemiol 52:487-497.
- 465 2. Schulz KF, Grimes DA. 2002. Blinding in randomized trials: hiding who got what. Lancet
 466 359:696–700.
- 467 3. Edwards SJ, Lilford RJ, Braunholtz D, Jackson J. 1997. Why Underpowered Trials are not
 468 Necessarily Unethical. Lancet 350:804–847.
- 469 4. Wilcken B. 2001. Rare Diseases and the Assessment of Intervention: What sorts of Clinical
 470 Trials can we use? J Inherit Metab Dis 24:291–298.
- 471 5. Gerss JW, Kopcke W. 2010. Clinical trials and rare diseases. Adv Exp Med Biol 686:173–
- 472 190 DOI: 10.1007/978-90-481-9485-8_11.
- 473 6. 2010 The needs of the few. Nature 466:160.

474	7.	Mitsumoto J, Dorsey ER, Beck CA, Kieburtz K, Griggs RC. 2009. Pivotal studies of
475		orphan drugs approved for neurological diseases. Ann Neurol 66:184-190 DOI:
476		10.1002/ana.21676.
477	8.	James OB. 1985. Statistical decision theory and Bayesian Analysis (2nd edition). New
478		York: Springer-Verlag, 118-387.
479	9.	Donald AB, Bert F. 1985. Bandit problems: sequential allocation of experiments. London;
480		New York: Springer Netherlands, 83-85; 191-206.
481	10.	Berger-Tal O, Nathan J, Meron E, Saltz D. 2014. The exploration-exploitation dilemma: a
482		multidisciplinary framework. PLoS One 9:e95693 DOI: 10.1371/journal.pone.0095693.
483		eCollection 2014.
484	11.	Savage LJ. 1972. The Foundation of Statistics (2nd Revised Edition). New York: Dover
485		Publications.
486	12.	Bernardo JM, Smith AFM. 1994. Bayesian Theory. Chichester: Wiley, 105-164; 377-426.
487	13.	Kroese D. P, Brereton T, Taimre T, Botev Z. I. 2014. Why the Monte Carlo method is so
488		important today. WIREs. Comput Stat 6:386-392.
489	14.	Robbins H. 1952. Some aspects of the sequential design of experiments. Bull Am Math
490		Soc 58:527-535.
491	15.	Auer P, Cesa-Bianchi N, Fischer P. 2002. Finite-time analysis of the multiarmed bandit
492		problem. Mach Learn 47:235-256.

- 493 16. Press WH. 2009. Bandit solutions provide unified ethical models for randomized clinical
- trials and comparative effectiveness research. Proc Natl Acad Sci 106:22387-22392 DOI:
- 495 10.1073/pnas.0912378106. Epub 2009 Dec 14.

496	17.	Gittins JC. 1979. Bandit processes and dynamic allocation indices. J R Stat Soc Series B				
497		41:148-177.				
498	18.	Gittins JC, Jones DM. 1979. A Dynamic allocation index for the discounted multiarmed				
499		bandit problem. Biometrika 66:561-565.				
500	19.	Karoui NE, Karatzas I. 1993. General Gittins index processes in discrete time. Proc Natl				
501		Acad Sci 90:1232-1236.				
502	20.	Sun S, Wu Y, Ma Y, Wang L, Gao Z, Xia C. 2016. Impact of Degree Heterogeneity on				
503		Attack Vulnerability of Interdependent Networks. Sci Rep 6:32983 DOI:				
504		10.1038/srep32983.				

Figure 1

The visual explanation of decision values of E-strategy and T-strategy (v_{*E} and $v_{*\tau}$).

The red curve shows the probability density function (PDF) of the beta posterior distribution of the success rate of treatment *A* with 13 successes and 5 failures; and the green curve shows one of the treatment *B* with 2 successes and 1 failure. The red and green vertical lines indicate the posterior means of two distributions, v_{AE} and v_{BE} , respectively. Actually $v_{AE} = 0.7$ and $v_{BE} = 0.6$. The vertical line that demarcates the gray areas under the curves indicates the target value, t = 0.8. The gray areas under the curves indicate the probability that their success rates are higher than the target value t = 0.8 for the two strategies, v_{AT} and v_{BT} . Actually $v_{AT} = 0.16$ and $v_{BT} = 0.18$. Because $v_{AE} > v_{BE}$, people with E-strategy will select the treatment *A* and because $v_{AT} < v_{BT}$, people with T-strategy will select the treatment *B*.



Figure 2

Comparison of a homogeneous population with E-strategy (*E.st*) and a homogeneous population with T-strategy (*T.st*) when the true success rates of A and B were 0.8 and 0.6 and the optimistic attitude index of *T.st* was 0.5.

The panels A and B show the results when 100 patients' outcomes have been recorded for *E.st* and *T.st*, respectively. Each panel is the two-dimensional histogram where the cyan area indicates the support plane and one axis "Success rate" indicates the overall success rate for 100 patients and the other axis "A-arm Fraction" indicates the fraction of patients who selected treatment A. The vertical axis indicates the exact probability of occurrence in the stochastic process. The panel A for *E.st* shows two peaks; one peak's overall success rate was around 0.8 and its A-arm fraction was close to 1 and the other peak's overall success rate was around 0.6 and its A-arm fraction was close to 0. The first peak was higher than the other peak. These findings indicated that the majority of individuals with *E.st* selected treatment A in many occasions but that in some occasions, they selected treatment B rather than treatment A. The panel B for T.st shows one peak and its overall success rate was around 0.8 and its A-arm fraction was close to 1. The mountain in Panel B was lower than the mountain located nearby in Panel A. These findings indicated that the majority of individuals with T.st selected treatment A in almost all the occasions, although the A-arm fraction tended to be lower than *E.st.* (C) and (D) show how the two measures change based on the homogeneous *E.st* versus *T.st* while the patient number changed, N=1,2,...,100, and horizontal axis shows the patient's number, and vertical axis shows the measure " A-arm fraction" in the Panel C and the measure "overall success rate (OSR)" in the panel D, where homogeneous *E.st* is labeled in black and *T.st* is in red.



Figure 3

The relation between true success rates of two treatments and overall success rates of two strategie.

The difference of the OSRs (*T.st* – *E.st*) between homogeneous *E.st* decision-makers and homogeneous T.st decision-makers with the fixed attitude w values, on the conditions where 5,050 of true success rates pairs ($(a,b) = \{(a,b) \mid a, b \in 0.01, 0.02, ..., 0.99\}, 1 > a >= b >0\}$) were calculated for multiple w values of T.st population (w = 0.8, 0.4, -0.4, and -0.8). The twelve plots in the panel A indicate N = 10, 30 and 100 for every w values. Each plot has a triangle area that corresponds to the (a,b) pairs where a >= b. The negative and positive values of difference of OSRs are coded in blue and red, respectively. The black curves in the triangle areas indicate the (a, b) pairs without difference in OSRs between two strategies. In the panel B, the red indicates the (a,b) pairs where the optimistic homogeneous *T.st* decision-makers performed better than *E.st* with regardless of the optimistic w values where the number of the fixed w is 2000, with w $\in \{0,...,0.999,0.9995\}$, the blue indicates the area of (a,b) pairs where *E.st* performed better, and the gray indicates the (a,b) pairs where the *T.st* or *E.st* performed better with depending on the w value.

NOT PEER-REVIEWED



Figure 4

Comparison of the Homogeneous *E.st* population and Heterogeneous *T.st* population.

The figure 4 indicates the difference of the OSRs between the decision-makers in a population with heterogeneity of decision attitudes (w values of individuals are various) and the homogeneous *E.st* decision-makers. The four panels separately indicate the difference of OSRs based on the four distributions of attitudes with given various beta parameter u values of 5,30,70, and 500 separately, where the distribution of w is located on the top-left of each image (w values were generated from the distribution shown as the equation 6 in the method section). The triangle area in each image corresponds to 5050 success rate pairs (a,b)= $\{(a,b) | a,b \in \{0.01,0.02,...,0.99\}, a \ge b\}$, and the difference of OSRs in the each triangle was coded with blue and red with representing negatives and positives, respectively. The sample size N=50. Such OSRs for the heterogeneity population were the average of 3000 Monte Carlo iterations.



Table 1(on next page)

Bernoulli Outcomes for Two Treatments After N Decision Processes.

 N_{As} and N_{Af} are favorable (successes) and unfavorable (failures) outcomes of patients who selected treatment A. Correspondingly, N_{Bs} and N_{Bf} are favorable and unfavorable outcomes of patients who selected treatment B.

	Favorable Outcome	Unfavorable Outcome	
A	N _{As}	N _{Af}	N _A
В	N _{Bs}	N _{Bf}	N _B
	N _s	N _f	N

1 Table 1. Bernoulli Outcomes for Two Treatments After N Decision Processes.

- 2 N_{As} and N_{Af} are favorable (successes) and unfavorable (failures) outcomes of patients who selected
- 3 treatment A. Correspondingly, N_{Bs} and N_{Bf} are favorable and unfavorable outcomes of patients
- 4 who selected treatment *B*.

5