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1 **Long-term evolution of preferences for conservation projects**
2 **in Seto Inland Sea, Japan: A comprehensive analytic**
3 **framework**

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14 Abstract

15 The long-term evolution of preferences for nature is crucial to conservation projects given their targeted
16 long-term horizons. Neglecting to account for this evolution could lead to undesirable human–nature
17 relationships. This study compares the willingness to pay (WTP) for three coastal conservation projects
18 in the Seto Inland Sea, Japan, at two distant time points (1998 and 2015) and tests for temporal
19 transferability. It also compares protest responses that are often overlooked in WTP practices,
20 regardless of their utility for conservation projects. Given the lack of a unanimous protocol for protest
21 response analyses and their use in estimating WTP, we propose a comprehensive analytic framework
22 that integrates the two. We show that while preferences for coastal ecosystem services were overall
23 stable and temporarily transferable, the preferences for certain aspects of conservation projects
24 considerably changed. This suggests the need to reconsider the projects' scheme, not the ecosystem
25 services themselves, along with the clarification of beneficiaries and those responsible for past
26 destruction. We conclude by suggesting further studies with focus on regions experiencing significant
27 social-ecological changes, such as developing countries, by exploiting the rich asset of existing
28 valuations. This could contribute to the database for more temporal-sensitive ecosystem service
29 valuations utilized for benefit transfers.

30 Keywords:

31 Willingness to pay; Protest responses; Coastal ecosystem services; Benefit transfer; Contingent valuation
32 method

33 1. Introduction

34 Understanding people's valuations of nature and how they change in the long run is of crucial
35 importance to establishing and sustaining desired relationships with nature (Uehara and Mineo, 2017;
36 Uehara et al., 2016). The potential evolution of preferences for nature indicates that a conservation
37 project assuming constant preferences could mislead us to an undesired state. Ideally, conservation
38 projects should be adaptive to the evolution (Skourtos et al., 2010).

39 A key approach to understanding how people value nature (i.e., welfare) is measuring willingness to pay
40 (WTP) for ecosystem services, whose techniques have been well-developed in environmental economics
41 (Freeman et al., 2014; Gómez-Baggethun et al., 2010). However, there is scope for improvement in
42 WTP-based studies; in particular, two aspects could foster better contributions to the literature. First,
43 WTP studies generally conduct one-time estimates and relative to the time horizon for conservation
44 projects, related research on the evolution of WTP are short termed (e.g., from few weeks to a year)
45 (Skourtos et al., 2010). These short-term studies tend to focus on stability and equality rather than
46 evolution or changes (Jakus, 2005). Moreover, they often assume the temporal stability of WTP, rather
47 than conducting explicit tests (Brouwer and Bateman, 2005; Costanza et al., 1997, 2014) or using simple
48 variations in previous WTP estimates for future projections (Kubiszewski et al., 2017). Second, WTP
49 estimates commonly exclude protest respondents (Brouwer and Martín-Ortega, 2012), that is,
50 respondents who reject certain aspects of a conservation project presented in a survey by saying no to a
51 proposed bid for the project, even though they positively value the ecosystem services (Freeman et al.,
52 2014). Protest responses provide non-negligible information for conservation projects in a real-world
53 context, not in a vacuum (García-Llorente et al., 2011). While WTP reveals preferences for ecosystem
54 services that benefit from conservation projects, protest response analyses highlight preferences for
55 project design and implementation. Reviewing past environmental valuation studies, Meyerhoff and

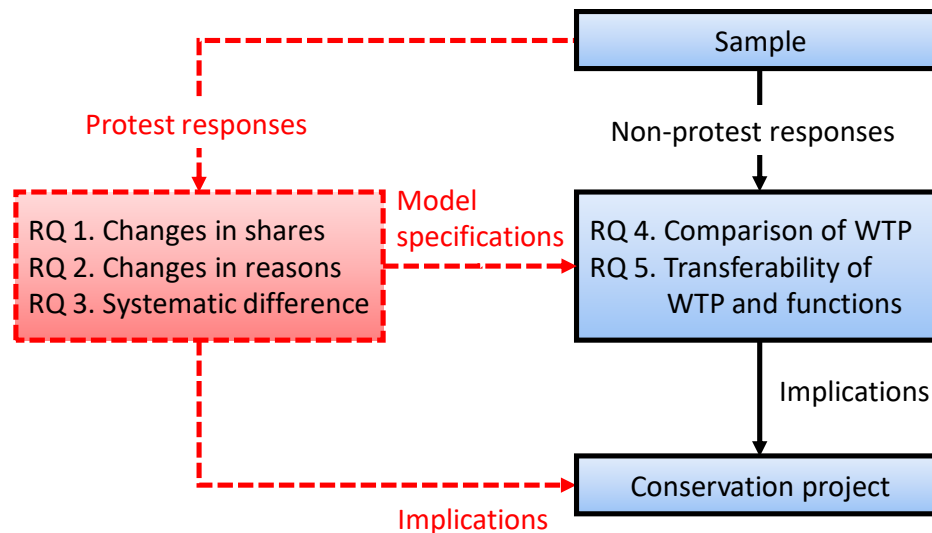
56 Liebe (2010) found that, on average, the rate of protest responses is 17.69%, indicating that a simple
57 disposal could result in a significant loss of information. In addition, it could lead to a biased WTP
58 estimate if people who protest systematically differ from those who do not (Brouwer and Martín-
59 Ortega, 2012; Freeman et al., 2014).

60 Our study aims to understand the long-term evolution of preferences for coastal ecosystem services by
61 addressing the abovementioned underdeveloped yet crucial topics: evolution of welfare measured in
62 WTP and that of protest responses. Since common WTP practices exclude protest responses and there is
63 no unanimous protocol on how to deal with them (Meyerhoff and Liebe, 2010), we propose a
64 comprehensive analytic framework that integrates WTP estimation and protest response analyses and
65 comprises five research questions. We compare the coastal and non-coastal residents' preferences for
66 three hypothetical projects that provide coastal ecosystem services in the Seto Inland Sea (SIS), Japan, at
67 two distant time points, 1998 and 2015. A 17-year difference is sufficient to include select next
68 generations that were not included in the 1998 survey.

69 2. Materials and Methods

70 Figure 1 presents the comprehensive analytic framework with the five research questions. While
71 common practices focus on the temporal comparison of WTP at two time points (RQ4) as well as the
72 temporal transferability of WTP and functions of WTP estimates (RQ5) (Brouwer and Bateman, 2005;
73 Brouwer and Spaniks, 1999; Downing and Ozuna, 1996; Rosenberger, 2015; Zandersen et al., 2007), the
74 present framework adds three research questions. How have the shares of protest responses changed
75 (RQ 1)? How have the reasons for protest responses changed (RQ 2)? Is there a systematic difference
76 between protesters and non-protesters (RQ 3)? RQ 3 could provide important information about model
77 specifications that could elicit unbiased WTP, as will be discussed later.

78



79

80 Figure 1. Comprehensive analytic framework for evolution of preferences for conservation projects. The
 81 red dashed lines indicate a procedure that is overlooked in common WTP practices and benefit
 82 transfers.

83

84 2.1 Three hypothetical projects in Seto Inland Sea

85 The three hypothetical projects were designed to elicit WTP for coastal ecosystem services in the Seto
 86 Inland Sea (SIS), an enclosed coastal sea in western Japan (Figure 2). The SIS was rich in ecosystem
 87 services; however, these ecosystems were destroyed or degraded with the rapid economic progress
 88 since the mid-20th century, resulting in, for example, declining fish catches, destruction of coastal zones
 89 for landfills and other anthropocentric uses, and water pollution (The Association for the Environmental
 90 Conservation of the Seto Inland Sea, 2015).

91 The projects include the restoration of the natural beauty of coastlines (project 1), conservation of
 92 seagrass beds as cradles of the sea (project 2), and protection of natural coastlines through a national
 93 trust (project 3) (see supplementary information 1 (SI1) for more details). We adopt the same three

94 projects in both 1998 and 2015; however, owing to certain changes in the SIS, we present changes in the
95 hypothetical projects in the 2015 survey.



96

97

Figure 2. Location of Seto Inland Sea, Japan

98 2.2 Data generating processes

99 An internet survey was conducted in 1998 and 2015, in which coastal and non-coastal residents were
100 asked to respond to a questionnaire on WTP estimates for the three projects. For the 1998 survey, we
101 utilized raw data collected by Tsuge and Washida (2003). The survey website was posted at the top page
102 of a national newspaper website, Asahi Shimbun (www.asahi.com). An average of 440,000 people
103 visited the website per month. The survey was conducted between December 1 and 12, 1998. The 2015
104 survey was also posted at Asahi Shimbun, but the sample was collected with help from a survey
105 company to ensure the sample was sufficiently comparable with the 1998 survey. The survey was
106 conducted between December 2 and 7, 2015.

107 2.3 Protest response analysis

108 For the protest response analysis, we used the statistical analysis software STATA (Version 14.2) by
109 StataCorp LP (<http://www.stata.com>).

110 2.3.1 Coding protest responses

111 Information on protest responses was generated from the reasons for rejecting bids for projects in the
112 questionnaire. Respondents could either choose from the reasons available or provide an independent
113 answer. To conduct a quantitative analysis, we coded the open answers and created eight categories, as
114 will be explained in the Results section. Some reasons for rejection were valid and thus, not considered
115 a protest response (e.g., “I support these projects, but the contribution amount is too high.”). Since
116 there is no clear-cut definition of protest responses (Brouwer and Martín-Ortega, 2012), we follow
117 discussions in well-established textbooks on the valuation of the environment (Freeman et al., 2014;
118 Rosenberger and Loomis, 2017) to choose protest responses from reasons coded in the survey.

119 2.3.2 Detection of systematic difference

120 While we note the potential impact of systematic differences between protestors and non-protestors on
121 WTP estimates (Freeman et al., 2014), there is no unanimous protocol to address such differences
122 (Meyerhoff and Liebe, 2010). Here, we chose a logit model for a binary response to detect the
123 systematic difference. That is,

$$124 \Pr(y_i = 1 | \mathbf{x}_i) = \frac{e^{\mathbf{x}_i \boldsymbol{\beta}}}{1 + e^{\mathbf{x}_i \boldsymbol{\beta}}}, \quad (1)$$

125 where y_i is a binary response regarding whether respondent i is a protestor (“1”) or non-protestor (“0”).
126 \mathbf{x}_i is a vector of explanatory variables and $\boldsymbol{\beta}$ is a vector of coefficients. With the logit model, we can
127 identify factors that determine whether a respondent is a protestor.

128 2.4 Welfare analysis

129 To analyze dichotomous choice-contingent valuation data (Version 0.0.15), we used a package by
130 Nakatani, Aizaki, and Sato (<https://cran.r-project.org/web/packages/DCchoice/citation.html>) run on R
131 (Version 3.3.2 for Windows (64 bit)) by the R foundation (<https://www.r-project.org/>).

132 2.4.1 WTP estimate

133 We used a single-bounded dichotomous choice format because it is less susceptible to a bias than open-
134 ended or payment card formats (Mitchell and Carson, 1989). As bids, each respondent was shown one
135 among six randomly selected amounts: 500, 1,000, 3,000, 8,000, 15,000, and 30,000 JPY. The
136 respondents were then asked if they were willing to pay the amount toward the implementation of each
137 project. We assumed that the payment would be made only once. To ensure that the respondents
138 recognized the payment burden, we explained that the donation amount would be deducted from the
139 money used for other household purposes. Those who agreed to donate the amount were asked to
140 specify the expenditures they forfeit for the donation.

141 The response data were analyzed using the binary logit model derived from the random utility model
142 (Hanemann, 1984). In the model, the following is assumed as the utility respondent k obtains from the
143 alternative:

$$144 \quad U_{ki} = V_{ki} + \varepsilon_{ki}, \quad (2)$$

145 where i takes the symbol y when respondent k answers “yes” to the bid and n when respondent k
146 answers “no.” V_{ki} and ε_{ki} represent the observable deterministic term and unobservable error term of
147 utility. It is assumed that respondent k considers cost and environmental improvement realized by the
148 conservation project and chooses an alternative with higher utility. The probability P_{ky} that the
149 respondent k will answer yes is equal to the probability that the utility from the alternative U_{ky} , is larger
150 than the utility from the alternatives n , U_{kn} , as described below:

151 $P_{ky} = Pr(U_{ky} > U_{kn}) = Pr(V_{ky} + \varepsilon_{ky} > V_{kn} + \varepsilon_{kn}).$ (3)

152 Assuming error term ε_{ki} follows a type-I extreme value distribution (Gumbel distribution), probability
 153 P_{ky} is described by the following binary logit model:

154 $P_{ky} = \frac{1}{1+e^{-\Delta V}},$ (4)

155 where ΔV denotes the utility difference function and the following log-linear function is assumed: $\Delta V =$
 156 $\alpha + \beta \ln T_k$. In the utility difference function, T_k represents the bid offered to respondent k and α and β
 157 indicate the utility obtained from environmental improvement and utility obtained from the payment.
 158 By extending the utility difference function as follows, it is possible to analyze the influence of other
 159 factors (e.g., household income) on the respondents' answers: $\Delta V = \alpha + \beta \ln T_k + \boldsymbol{\gamma} \mathbf{z}_k$, where, \mathbf{z}_k is a
 160 vector of other factors possibly affecting respondent k 's answer and $\boldsymbol{\gamma}$ is a vector of parameters for
 161 those factors.

162 The parameters are estimated by the maximum likelihood method (Greene, 2014). The log likelihood
 163 function can be written as follows:

164 $\ln L = \sum_k \sum_i \delta_{ki} \ln P_{ki},$ (5)

165 where δ_{ki} is a dummy variable such that $\delta_{ki} = 1$ when respondent k answers yes to a bid and $\delta_{ki} = 0$
 166 otherwise.

167 The mean WTP can be calculated using the estimated parameters, α and β (Hanemann, 1984). It is
 168 obtained by integrating the probability that the respondent will answer yes to the bid. However, since it
 169 is not realistic to integrate an extremely high amount, the maximum bid is often used as the integration
 170 upper limit. In this case, the mean WTP is calculated as follows:

171 $Mean\ WTP(truncation\ at\ T_{max}) = \int_0^{T_{max}} P_{ky} dT,$ (6)

172

173 where T_{max} is the maximum bid.

174 2.4.2 Confidence Intervals

175 We calculated the confidence intervals using Krinsky and Robb's technique (1986), which is often
176 employed in stated preference methods such as contingent valuation method (CVM) and conjoint
177 analysis (Downing and Ozuna, 1996; Lew and Wallmo, 2017; Matthews et al., 2017; Zandersen et al.,
178 2007). Using the technique, we draw 10,000 random coefficients and compute 10,000 mean WTP
179 measures. Then, we ordered the 10,000 mean WTP measures from the smallest to largest and selected
180 the 95% confidence limits.

181 2.4.3 Transferability test

182 Since it is impossible for people who are not born yet to report their future WTP and confirm the extent
183 to which a current project will be supported by future generations, we need to extrapolate a future WTP
184 value by exploiting value information currently available. A method that has been widely used is benefit
185 transfers, which involve transferring existing value information to a new context (Rosenberger and
186 Loomis, 2017). There are two primary types of benefit transfers: value and function transfers. Value
187 transfers are the direct application of summary statistics in existing research such as per unit measure of
188 WTP (Rosenberger and Loomis, 2017). It generally assumes constant preferences over time (e.g.,
189 Coastanza et al., 1997, 2014). Function transfers tailor value estimates by reflecting differences in the
190 characteristics of contexts in a model estimating WTP.

191 In addition to a strong interest in the transferability of WTP estimates given the scarcity of resources,
192 the temporal stability of the estimates and its testing methods have been extensively studied, although
193 most studies are limited to the short term (from a week to two years) (Skourtos et al., 2010). The key
194 focus is the statistical equality of WTP and coefficient parameters of models using various statistical

195 tests such as the t-test, Wald test, likelihood ratio test, Mann–Whitney test, and Kolgorov–Sminov test
196 (Brouwer and Spaninks, 1999).

197 However, we did not conduct these statistical tests for two reasons. First, the statistical tests examine
198 for statistical equality and ignore acceptable levels of accuracy in a real-world context (Rosenberger,
199 2015). A review by Rosenberger (2015) reveals that most studies failed to pass these tests. Second, the
200 coefficients estimated by the logit model used in this study are not purely parameters of the utility
201 function but products of parameters of the utility function and scale parameter (Train, 2009). Therefore,
202 testing the statistical equality of the estimated coefficients does not necessarily mean examining the
203 statistical equality of the parameters of the utility function and there is a possibility of erroneous
204 judgments on the latter. On the other hand, since WTP estimates are calculated from the ratio of
205 estimated coefficients, the scale parameters of the numerator and denominator are canceled out and
206 are not affected by them. Therefore, it is more meaningful to test for the statistical equality of WTP
207 estimates.

208 Hence, we evaluated the performance of value and function transfers by conducting a percentage
209 transfer error (PTE) test, which is a type of transfer error test (Rosenberger, 2015) that measures the
210 difference between the benefit transfer value (estimated using 1998 values or functions) and true value
211 (2015 estimates). While the abovementioned tests focus on equality, this test estimates maximum
212 transfer error. The percentage transfer test is calculated as

$$213 \text{ PTE} = \left[\frac{V_T - V_P}{V_P} \right] \times 100, \quad (7)$$

214 where V_T is the transfer estimate and V_P is the known or actual estimate for the policy site. PTE then
215 measures the degree of difference between the transferred and actual estimates at the policy site.
216 Typically, PTE requires both estimates to be available within the context of a primary study that has
217 derived them (Rosenberger, 2015, p.309).

218 For a function transfer, we used models that include income as an explanatory variable and incorporate
219 average income for 2015 in the 1998 models to estimate. There are three reasons to include income as
220 an explanatory variable. First, it is consistent with economic theory underpinning this method (Brouwer
221 and Bateman, 2005; Hanemann, 1984). Second, it is statistically significant (Brouwer and Bateman,
222 2005) in many empirical studies including the present analysis. Third, long-term income projection has
223 been well-studied and available from various sources such as government agencies and the Organisation
224 for Economic Co-operation and Development.

225 3. Results

226 While the sample sizes have the same order of magnitude (5,632 respondents for 1998 and 7,264
227 respondents for 2015), there are significant differences in the rates of internet accessibility that could
228 affect the compositions of the samples (SI2 for descriptive statistics) and the following analyses. In
229 Japan, personal accessibility to the Internet significantly increased from 13.4% in 1998 to 83.0% in 2015
230 (Ministry of Internal Affairs and Communications, 2017). While company employees account for the
231 largest share of respondents in both 1998 and 2015, which is consistent with population characteristics,
232 there are certain differences between the years. For instance, university students with internet access
233 accounted for 14% of the respondents in 1998 but 6% in 2015. Similarly, the rate of part-time workers,
234 unemployed, and housewives who did not have good internet access was 36% in 2015 but 3% in 1998.
235 Because of the potential biased samples, we calculated confidence intervals using the Krinsky–Robb
236 technique in the following analysis on welfare changes and temporal transferability.

237 3.1 Protest responses

238 3.1.1 RQ 1: share of protest responses

239 The share of protest responses is greater in 2015 than in 1998 and larger for non-coastal residents
240 compared with coastal residents for all three plans (Table 1). The analysis of variance (ANOVA) shows

241 that these differences in the shares by year and geographical origin are statistically significant at the
 242 10% level for all the projects.

	Project 1		Project 2		Project 3	
	1998	2015	1998	2015	1998	2015
Coastal residents	24.4%	37.3%	20.4%	32.6%	17.5%	33.7%
Non-coastal residents	25.8%	41.6%	22.4%	39.0%	18.9%	38.2%
Total	25.7%	41.3%	22.2%	38.6%	18.8%	37.9%

243 Table 1. Shares of protest responses by year and geographical origin

244 3.1.2 RQ 2: reasons for protest responses

245 We coded the protest responses into six types on the basis of multiple choices and open answers to
 246 reasons underpinning the rejection of a bid proposed for the projects. Here, we show the categorization
 247 of reasons by year and geographical origin. The patterns are similar across all the projects, except for the
 248 change in reason 1 for non-coastal residents.

249 The share of coastal residents increased for all three projects and they accounted for the second highest
 250 number of respondents protesting contribution to a fund (reason 1). The respondents were asked to
 251 contribute to a newly established local fund (The SIS Environment Conservation Fund) to implement the
 252 projects. In the open answers, some respondents stated that it should be funded from tax revenues
 253 because it is a public good whose cost should be incurred by everyone.

254 In both years, coastal and non-coastal residents accounted for the highest numbers in terms of the
 255 belief that funding was not their personal responsibility (reason 2). In particular, the number of non-
 256 coastal residents significantly increased for reason 2 and accounted for a larger share of protest
 257 responses for 2015. In the open answers, both coastal and non-coastal residents claimed that the

258 project(s) should be funded by people who are responsible for the environmental destruction, such as
 259 private companies and municipalities. A characteristic unique to non-coastal residents is that while they
 260 valued these projects, some preferred to conserve the environment closer to their place of residence.
 261 The rate of respondents opposed to the program itself (reason 3) was lower in 2015. In the open
 262 answers, certain respondents who chose reason 3 stated they were dubious about the effectiveness of
 263 the project(s). For example, some pointed out that the scales of the projects are too small to realize the
 264 benefits mentioned.

Reason to oppose bid	Coastal residents			Non-coastal residents		
	1998	2015	t-test	1998	2015	t-test
1. I support these projects, but I am against contributing to a fund.	19%	33%	***	23%	27%	***
2. I support these projects, but I don't think I need to personally take responsibility for funding.	42%	40%		37%	54%	***
3. I am opposed to the program itself.	34%	25%	**	33%	18%	***
4. I do not trust the survey.	3%	1%		3%	0%	***
5. Information is insufficient to make a judgment.	1%	2%		4%	1%	***
6. Did not understand the questionnaire.	1%	0%		1%	0%	***
Total	100%	100%		100%	100%	
N	96	170		1,330	2,762	

265 ***p < 0.01, ** p < 0.05, * p < 0.10

266 Table 2a. Composition of reasons for protest by year and geographical origin for project 1. The t-tests
 267 examine the null hypothesis of no difference in numbers of times each reason was cited between 1998
 268 and 2015.

Reason to oppose bid	Coastal residents			Non-coastal residents		
	1998	2015	t-test	1998	2015	t-test
1. I support these projects, but I am against contributing to a fund.	20%	38%	***	28%	29%	
2. I support these projects, but I don't think I need to personally take responsibility for funding.	46%	49%		41%	60%	***
3. I am opposed to the program itself.	26%	11%	***	23%	10%	***
4. I do not trust the survey.	4%	1%	**	4%	0%	***
5. Information is insufficient to make a judgment.	3%	1%		4%	1%	***
6. Did not understand the questionnaire.	1%	0%	*	1%	0%	***
Total	100%	100%		100%	100%	
N	80	149		1,155	2,595	

269 *** p < 0.01, ** p < 0.05, * p < 0.10

270 Table 2b. Composition of reasons for protest by year and geographical origin for project 2.

Reason to oppose the bid	Coastal residents			Non-coastal residents		
	1998	2015	t-test	1998	2015	t-test

1. I support these projects, but I am against contributing to a fund.	22%	38%	***	30%	29%
2. I support these projects, but I don't think I need to personally take responsibility for funding.	51%	47%		44%	59%
3. I am opposed to the program itself.	20%	11%	**	18%	10%
4. I do not trust the survey.	3%	2%		3%	1%
5. Information is insufficient to make a judgment.	3%	2%		3%	1%
6. Did not understand the questionnaire.	1%	0%	*	1%	0%
Total	100%	100%		100%	100%
N	69	154		974	2,538

271 Table 2c. Composition of reasons for protest by year and geographical origin for project 3.

272 3.1.3 RQ 3: systematic difference between protestors and non-protestors

273 Before building a model to estimate WTP, we tested the systematic difference between protestors and
 274 non-protestors. The WTP estimate could be biased if there is a systematic difference (Freeman et al.,
 275 2014). However, there is no unanimous protocol for the treatment of protest responses (Tobarra-
 276 González, 2015). Here, we adopted a logit model to explore factors influencing a respondent's choice to
 277 protest or not. We chose place of residence, income, and year as explanatory variables. Because of the
 278 17-year gap between 1998 and 2015, the samples were considered to be drawn from different
 279 populations. Given that respondents are geographically located in different areas, differentiating WTP
 280 by place of residence could also be informative for conservation projects (e.g., more targeted

281 fundraising). Income is a key variable consistent with economic theory (Brouwer and Bateman, 2005;
 282 Hanemann, 1984). The results revealed (Table 3) that all three variables explain the respondents' choice
 283 to protest at the statistically significant levels, indicating the possibility of a systematic difference
 284 between protestors and non-protestors by place of residence, income, and year. Therefore, it would be
 285 desirable to estimate WTP by constructing models on the basis of these three variables. However, since
 286 income has 15 categories and it is not realistic to model each category separately, we use income as an
 287 explanatory variable, which is also consistent with economic theory (Hanemann, 1984). Accordingly, we
 288 constructed four models for each project, resulting in a total of 12 models.

	Project 1			Project 2			Project 3		
	Coefficient	t-stat		Coefficient	t-stat		Coefficient	t-stat	
Coastal_dummy	-0.164	-1.95	*	-0.241	-2.74	***	-0.173	-1.93	*
Income	-0.014	-2.36	**	-0.015	-2.42	**	-0.020	-3.03	***
Year	0.035	13.92	***	0.041	15.53	***	0.051	18.62	***
Constant	-71.479	-14.01	***	-82.404	-15.63	***	-102.785	-18.73	***
N	10,933			10,938			10,937		
Log-likelihood	-6,808.384			-6,537.457			-6,244.689		

289 *** p < 0.01, ** p < 0.05, * p < 0.10

290 Table 3. Logit models for three projects

291 3.2 Welfare

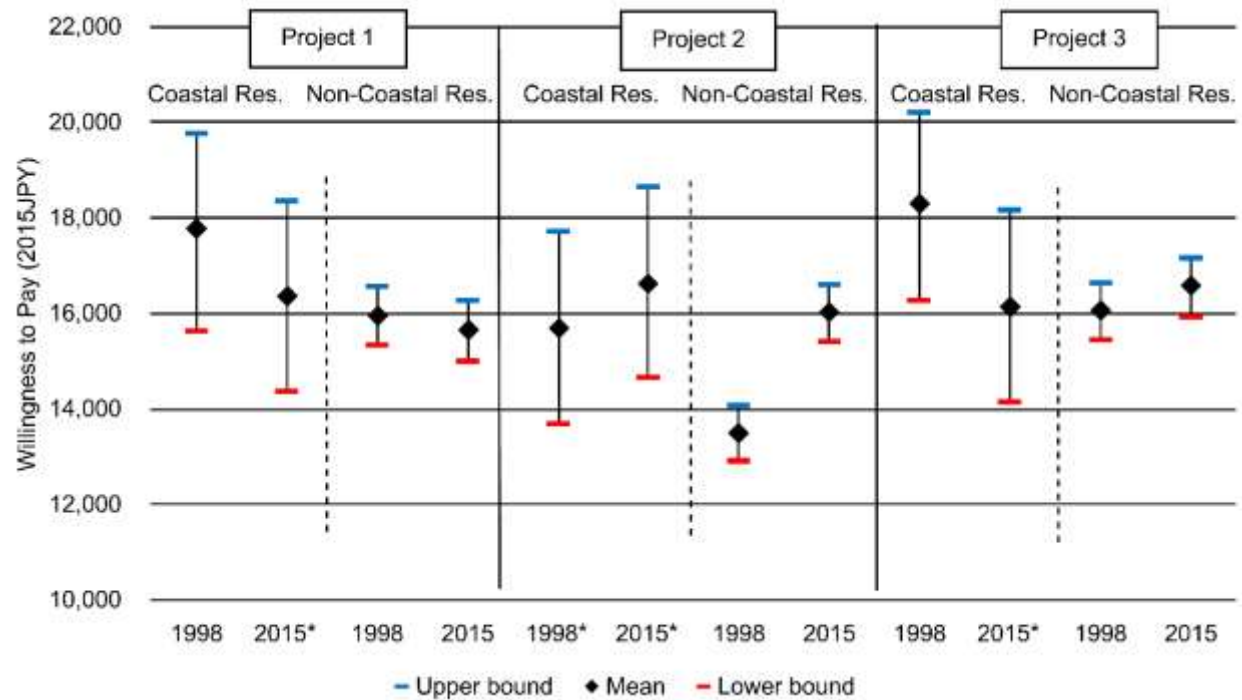
292 3.2.1 RQ 4: confidence intervals

293 This research question addresses the extent to which preferences for coastal ecosystem services (i.e.,
 294 welfare obtained from the services) have evolved over the 17 years by measuring changes in people's

295 WTP for the projects. On the basis of the protest response analysis, we built models to estimate WTP by
296 year and geographical origin with income as an explanatory variable when it is statistically significant
297 (SI3).

298 Figure 3 shows changes in the mean WTP with 95% confidence intervals across 17 years by geographical
299 origin. The sample sizes of each geographical origin have the same order of magnitude across time:
300 286–303 respondents for coastal residents and 3,146–3,772 respondents for non-coastal residents. The
301 confidence intervals were relatively wider for coastal residents due to their smaller sample sizes. The
302 confidence intervals among the same geographical origin are comparable because of the similar sample
303 sizes.

304 The confidence intervals overlap for all models, except for non-coastal residents in the case of project 2
305 (conservation of seagrass beds as cradles of the sea), indicating that only welfare obtained from project
306 2 increased for non-coastal residents at the statistically significant level. The mean WTP increased by
307 18.6 %, from 13,497 JPY in 1998 to 16,006 JPY in 2015. This change is also the largest among the point
308 estimates of the mean WTP: –7.9% for coastal residents in project 1, –1.8% for non-coastal residents in
309 project 1, 6.0% for coastal residents in project 2, –11.7% for coastal residents in project 3, and 3.3% for
310 non-coastal residents in project 3 (SI4).



311

312

Figure 3. Confidence intervals of mean WTP

313 “*” indicates that income was excluded from the model because it was not statistically significant (SI3
314 and 4).

315 3.2.2 RQ 5: temporal transferability

316 The research question is based on the extent of transferability of WTP estimates and models in 1998 to
317 those in 2015. Table 4 presents the absolute percentage transfer errors for value and function transfers.

318 PTE measures the difference in percentage between true values calculated using the 2015 data and
319 model and transferred values are estimated using value information for 1998. Of the six transfers, four
320 transfers performed better for the value transfer.

		Value transfer		Function transfer	
Residency		Mean PTE	Range	Mean PTE	Range
Project 1	Coastal	8.6	7.7–8.8	3.0	1.8–3.8

	Non-coastal	1.9	1.7–2.1	3.8	3.4–4.2
Project 2	Coastal	5.7	4.8–6.7	11.0	8.9–13.5
	Non-coastal	15.7	15.3–16.1	19.5	18.7–20.3
Project 3	Coastal	13.3	11.2–15.0	7.4	6.9–7.5
	Non-coastal	3.2	3.2–3.1	5.8	5.4–6.1

321 Table 4. Absolute percentage transfer errors with their ranges

322 4. Discussion

323 4.1 Protest responses

324 Protest response analyses are generally beyond the scope of WTP practices and benefit transfers.

325 However, our study revealed that it provides non-negligible information on successful conservation
 326 projects that are implemented in the real world, not in a vacuum. Changes in the share and composition
 327 of protest responses demonstrated those in preferences for other aspects of a conservation project
 328 rather than the value of ecosystem services measured in WTP.

329 Overall, these shares are larger than the average share of protest responses in previous studies (mean:
 330 17.69%; standard deviation: 11.30; median: 16.13; min.: 0; max.: 59.28) (Meyerhoff and Liebe, 2010).

331 The shares increase in 2015 and this indicates the growing importance of protest response analyses as a
 332 source of information for conservation projects. In addition to the possibility of an actual increase in the
 333 protest responses in the 2015 population, the drastic changes in internet accessibility resulted in biased
 334 samples with varying population attributes. However, since there appears to be no study on temporal
 335 changes in the share of protest responses, we are still unaware if an increase in the share of protest
 336 responses is a general trend and the factors influencing the increase. We leave this to further study.

337 The analysis on the reasons for protest highlights the need for policymakers to be adaptive and rethink
338 the manner in which projects are implemented. More specifically, there is a growing dislike for payment
339 methods (i.e., establishing a fund) (reason 1). In addition, respondents who do not want to personally
340 take responsibility for the funding (reason 2) account for the highest number of protestors, suggesting
341 the reconsideration of the payment method along with a reflection of those responsible for the past
342 destruction of ecosystems and project beneficiaries. Failure to account for these reasons could lead to
343 policymakers facing unexpected oppositions at the time of actual project implementation, even if the
344 WTP estimates that do not include protest responses indicate the projects as valuable. The choice of
345 payment method is an important aspect of a project scheme (Freeman et al., 2014) and this choice
346 should be sensitive to the social context (Fischhoff and Furby, 1988). In addition to the method of
347 payment, the recipient of these payments is an important aspect warranting consideration given that
348 according to protest responses, some non-coastal residents prefer spending money on a similar project
349 closer to their place of residence (reason 2).

350 The protest response analysis was also informative in correcting systematic differences caused by the
351 exclusion of the protest responses from the WTP estimate. To obtain an accurate WTP estimate, we
352 recommend the protest response analysis, especially when the share is not as small as that in our case
353 study.

354 4.2 Welfare

355 The confidence intervals showed that the welfare obtained from the projects measured in WTP was
356 stable over 17 years, except for non-coastal residents in the case of project 2 (conservation of seagrass
357 beds). These changes have the same order of magnitude as global estimates by Costanza et al. (2014),
358 who used the world database for valuation studies on ecosystem services: the unit value of estuaries
359 decreased by 8.2% (from 31,509 USD/ha per year in 1997 to 28,916 USD/ha per year in 2011) and the

360 unit value of seagrass or algae beds increased by 10.1% (from 26,226 USD/ha per year in 1997 to 28,916
361 2007 USD/ha per year in 2011). However, Pendleton et al. (2016) conducted a closer examination of the
362 data compiled by Costanza et al. (2014) and highlighted the lack of accuracy and comprehensiveness,
363 especially for marine and coastal areas. For example, the database includes estimates from more than
364 20 years ago by assuming the temporal stability (or non-changing) of a unit value. This supports the
365 importance of primary studies on the temporal valuation of marine and coastal ecosystem services.

366 The changes in WTP can be explained by factors affecting demand or supply of the ecosystem services
367 (Skourtos et al., 2010). Factors affecting demand could include income, prices of other goods, and
368 socioeconomic profile, while those influencing supply may be the amount and quality of ecosystem
369 services.

370 There are three possible reasons for a higher WTP for project 2 among non-coastal residents. First, while
371 project 2 was the same in both 2015 and 1998, the context was different. The 1998 survey presented
372 the *possibility* of the seagrass bed being destroyed: “Moreover, 50 hectares of the largest remaining
373 seagrass bed are currently proposed for reclamation as airport and harbor construction progresses.”
374 However, since the seagrass bed was destroyed by 2015, the corresponding survey addressed it as
375 *actually* destroyed. This can be considered a scarcity signal. Previous studies also show the sensitivity of
376 WTP to changes in the supply of ecosystem services in the SIS (Tokiyoshi et al., 2005). Second, the first
377 reason may attract attention from those who do not live in coastal zones because they benefit from the
378 airport and harbor construction. Finally, the restoration of the seagrass bed became a national agenda
379 and the Fisheries Agency launched an investigative committee for seagrass beds and mudflats about six
380 months prior to our survey.

381 The percentage transfer errors were not large compared with those in previous studies: the mean of the
382 mean PTE was 140 for the value transfer and 65 for the function transfer (Rosenberger and Loomis,

383 2017). The performance of a transfer is considered to depend on contextual similarity (Rosenberger and
384 Loomis, 2017). Since these previous studies are about a spatial transfer (i.e., between spatially different
385 sites) and not a temporal transfer (i.e., between temporally different but spatially same sites), the
386 contextual difference resulting from the 17-year gap is smaller than the spatial differences in previous
387 studies (Rosenberger and Loomis, 2017). It is difficult to judge whether these transfer errors are small
388 among long-term temporal transfer studies since such studies are limited. Zandersean et al. (2007)
389 conducted a study on forest ecosystem services in Denmark and reported a PTE of 25 for 52 forests
390 across a 20-year period. Boman et al. (2011) estimate this value at 17 in the case of Sweden.

391 In contrast to general tendency (Rosenberger, 2015; Rosenberger and Loomis, 2017), in the present
392 study, the value transfer performs better than the function transfer: of the six transfers, four are better
393 as value transfers. There are two possible explanations: temporal contextual similarity and insufficiency
394 of the function transfer. First, as Bateman et al. (2011, p.383) argued, “the choice of [value vs. function
395 transfer] depends crucially upon the degree of similarity of the sites under consideration.” As the
396 comparison of previous studies on spatial transfer revealed, study sites during 1998 and 2015 seemed
397 similar. Second, our function transfer did not sufficiently capture changes because in general, function
398 transfers perform better than value transfers as the former can increase transfer accuracy by reflecting
399 site characteristics (Rosenberger and Loomis, 2017). In technical terms, there are two types of changes
400 that affect WTP estimates. First are changes in the WTP estimate model’s arguments and the second are
401 those in the coefficients of the model (Whitehead and Hoban, 1999). The function transfers in our study
402 adjusted only income, an argument, and assumed that the coefficients are constant over 17 years.

403 In addition, it is notable that non-coastal residents’ mean PTE and its range for project 2 are the highest
404 for both value and function transfers. This is reasonable because neither the value nor function transfer
405 reflected the loss of the seagrass bed, a change in the supply side. This indicates that while welfare was
406 not as sensitive to time even in the long term, it was sensitive to changes in the supply of ecosystem

407 services (i.e., loss of seagrass beds for project 2). This calls a further study on contextually relevant
408 research with particular focus on supply-side changes. However, since Japan has been relatively stable in
409 a socioeconomic sense, our findings do not rule out the importance of other contextual changes that
410 affect the supply and demand of ecosystem services, such as income, demographics, perceptions of
411 nature, and the preference structure of individuals, through learning procedures or cultural
412 transmissions (Skourtos et al., 2010).

413 4.3 Limitations and future research

414 Our study is subject to two major limitations in terms of its implications for conservation projects: a
415 biased sample and context-dependent results. First, because of the significant changes in internet
416 accessibility from 1998 (13.4%) to 2015 (83.0%), the sample attributes may differ enough to influence
417 the results (see SI2 for descriptive statistics). Second, the overall stability of WTP could be attributed to
418 the specific context, Japan, where there is little drastic social-ecological change affecting the supply and
419 demand of ecosystem services. Therefore, the stability level found in this study might not be applicable
420 to other areas characterized by drastic social-ecological changes, such as developing countries.

421 To derive more general implications for conservation projects, further studies on the evolution of
422 preferences and development of methodology for protest response analyses are encouraged. Further,
423 given the asset of previous one-time studies on WTP estimates in various contexts and time periods
424 across the world (e.g., Ecosystem Services Valuation Database (ESVD) (Van der Ploeg and de Groot,
425 2010)), a similar temporal study should be conducted by exploiting the asset to elicit more general
426 findings about the long-term evolution of preferences for nature. The accumulation of such studies
427 would allow us to construct a more temporal-sensitive database for the valuation of ecosystem services
428 and conduct a better temporal and spatial benefit transfer. A caveat, however, is the availability of raw
429 data used for past studies. In particular, data for protest responses may be limited in their availability.

430 Priority should, thus, be given to cases in which drastic social-ecological changes and adaptive
431 conservation projects are expected. Furthermore, while our study used CVM, conjoint analysis, another
432 stated preference method to measure WTP could prove promising (Louviere et al., 2000). A conjoint
433 analysis uses profiles of a conservation project with various attributes (e.g., degree of conservation,
434 development, and cost) and elicits respondents' preferences for trade-offs among these attributes. This
435 could be particularly useful when a project faces serious trade-offs such as conservation vs.
436 development. In addition, a conjoint analysis can capture marginal changes in WTP caused by changes in
437 attributes such as quantity of ecosystem services supplied, and thus, could better capture factors
438 affecting supply and demand for ecosystem services.

439 As our study showed, protest responses could provide non-negligible information for a conservation
440 project in a real-world context. Protest responses are by no means residual information. In fact, several
441 attempts have been made to use protest responses (e.g., Cunha-e-Sá et al., 2012; Garcia-Llorente et al.,
442 2011), although such analyses remain underdeveloped (Brouwer and Martín-Ortega, 2012; Tobarra-
443 González, 2015).

444 5. Conclusions

445 This study investigated the evolution of preferences for conservation projects (i.e., welfare measured in
446 WTP and other project aspects assessed by protest responses). We compared preferences for three
447 conservation projects in the SIS, Japan, at two distant time points, 1998 and 2015. Owing to the lack of a
448 unanimous protocol for protest response analysis and its use for WTP estimate, we proposed a simple
449 comprehensive analytic framework that integrates protest response and WTP analyses.

450 Protest responses provide useful information to render a project adaptive to changes in the social-
451 ecological system, the SIS. For instance, the payment method should be reconsidered. The welfare
452 obtained from the projects was stable over a 17-year period, except for non-coastal residents in the case

453 of project 2. This possibly reflects the factors influencing changes in the demand and supply of the
454 ecosystem services. Since the percentage transfer errors for both value and functional transfers were
455 smaller than those in previous studies, they can be considered temporarily transferable. A function
456 transfer performs less than a value transfer because of the contextual similarity over time and the
457 insufficient specification of functions for the transfer.

458 Further temporal studies are highly encouraged with focus on locations where significant social-
459 ecological changes have occurred or are expected to occur, such as in developing countries. These
460 studies can contribute to not only the primary study site but also the accumulation and sophistication of
461 ecosystem services database such as ESVD. Doing so will enable a better benefit transfer when time and
462 budget are unavailable to conduct a primary study.

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