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Psychometric properties of the perceived stress scale (PSS): measurement invariance between athletes and non-athletes and construct validity

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Background: Although Perceived Stress Scale (PSS, Cohen, Kamarack, Mermelstein, 1983) has been validated and widely used in many domains, there is still no validation in sport by comparing athletes and non-athletes and examination of related psychometric indices. **Purpose:** The purpose of this study was to examine the measurement invariance of PSS between athletes and non-athletes, and examine construct validity in the sport contexts. **Methods:** Study 1 sampled 359 college student-athletes (males = 233; females = 126) and 242 non-athletes (males=124; females=118) and examined factorial structure, measurement invariance and internal consistency. Study 2 sampled 196 student-athletes (males = 139, females = 57, M_{age} =19.88 yrs, SD = 1.35) and examined discriminant validity and convergent validity of PSS. **Results:** Results found that 2-factor PSS-10 fitted the model the best and had appropriate reliability. Also, there was a measurement invariance between athletes and non-athletes; and PSS positively correlated with athlete burnout and life stress but negatively correlated with coping efficacy provided evidences of discriminant validity and convergent validity. **Discussion:** It is suggested that 2-factor PSS-10 can be a useful tool in assessing perceived stress either in sport or non-sport settings. We suggest future study may use 2-factor PSS-10 in examining the effects of stress on athletic injury, burnout, and psychiatry disorders.

Psychometric Properties of the Perceived Stress Scale (PSS): Measurement Invariance Between Athletes and Non-athletes and Construct Validity

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Abstract

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Discussion: It is suggested that 2-factor PSS-10 can be a useful tool in assessing perceived stress either in sport or non-sport settings. We suggest future study may use 2-factor PSS-10 in examining the effects of stress on athletic injury, burnout, and psychiatry disorders.

Key words: cognitive-transactional model of stress, multiple group comparisons, nested model, perceived coping

Since the development of the perceived stress scale (PSS, Cohen, Kamarack, Mermelstein, 1983) it has been widely used in various research such as the degree of global stress of a given situation (Leon, Hyre , Ompad , DeSalvo, & Muntner, 2007; McAlonan et al., 2007), or effectiveness of an intervention on psychological stress (Holzel et al., 2010; Seskevich & Pieper, 2007; Taylor-Piliae, Haskell, Waters, & Froelicher, 2006), or the associations of perceived stress and psychiatric/physical disorder (Culhane, Rauh, McCollum, Hogan, Agnew, & Wadhwa, 2001; Garg et al., 2001). In addition, many studies used PSS to examine its relationship with quality of life (Golden-Kreutz, Browne, Frierson, & Anderson, 2004; Golden-Kreutz et al., 2005), job satisfaction (Norvell, Walden, Gettelman, & Murrin, 1993), immune functioning (Burns, Drayson, Ring, & Carroll, 2002; Maes & Van Bockstaele, 1999), depression (Carpenter et al., 2004), and sleep quality (Cohen & Williamson, 1988). Therefore, it can be said PSS is a very important tool in assessing stress.

Built on Lazarus and Folkman's (1984) transactional model of stress, the development of PSS is to assess one's perceived nonspecific stress in a given situation or a daily life situation. Lazarus and Folkman's (1984) transactional model of stress contends that an individual's stress perception derived from the imblance between one's appraisal of situational demands and coping resources. If one perceives that situational demands over resources, and the consequences of such failure will be severe; then it will lead to psychophysiological responses such as fast heart beats, pale face, cold and sweat hands, tense muscles...etc. Cohen and colleagues (Cohen et al., 1983; Cohen & William, 1988) constructed this global stress perception by two important components --- something that one can control (i.e., counter stress) and something that one can't control (i.e., perceived stress). In such manner, PSS is not only a measure to assess an extant of how a given situation might hurt oneself but also to assess the degree of how this give situation is controllable or uncontrollable (Golden-Kreutz et al., 2004; Örücü & Demir, 2009; Roberti, Harrington, & Storch, 2006).

Given that the applicability of PSS in assessing perceived stress in many domains, the number of the item, the factorial structure, and the reliabilityof PSS have been intensively examined by many researchers. For example, athough the initial version of PSS (Cohen et al., 1983) has been developed as 14-item with a unidimensional measure, Cohen and colleagues (Cohen & William, 1988)

continuingly examined the appropriateness of the item's number and factorial structure of PSS. They sampled 960 male and 1,427 female US residents (*Mage*=42.8±17.2 years) and examined its factorial structure, criterion validity, and internal consistency of the PSS. Results found 10-item, by deleting item 4, 5, 12, and 13, 2-factor PSS-10 can be a better measuring tool of perceived stress because they found the revised version of PSS-10 accounted 48.9% of the variance, and had better reliabilities (Cronbach's α=.84~.86), and correlated with anxiety, depression and life events which indicated good construct validity.

Followed Cohen and colleagues (Cohen & William, 1988), Hewitt and colleagues sampled psychiary patient as participants and examined factorial structure and reliability of PSS, the exploratory factor analysis (EFA) found PSS had two factors --- perceived distress (Cronbach's α=.81) and perceived coping(Cronbach's α=.72) with 11 items. Similar findings also found in a Mexico sample where Ramírez and Hernández (2007) found two factors PSS-10 had better reliability and accounted 48.02% of variance. Recently, Barbosa-Leiker and colleagues (Barbosa-Leiker, Kostick, McPherson, Roper, Hoekstra, & Wright, 2013) examined measurement invariance of PSS across gender and time with clinical sample. Results indicated the 2-factor PSS model provided acceptable fit in both men and women at each time point. Therefore, it is understandable that researchers keep examining the psychometric properties of PSS is to seek the answers of the validity and reliability of the PSS in order to choose the best version of PSS for research.

In sport, stress is an important issue that has been recieved much of attention; specifically in the studies of athlete burnout and athletic injury. For example, Smith (1986) proposed a cognitive-affective model of athletic burnout to explain the influences of stress on burnout. Smith (1986) explains that in the stressful sport settings, athletes keep appraising the contextual stressors and personal coping resources. If athletes perceived situational demands surpass personal resources and consequences will be severe, the cognitive appraisal leads to severe physiological and psychological responses--- anxiety, tension, insomnia, and illness, which eventually lead to burnout. With same line of conceptualization, Andersen and Williams (1988) also proposed a 'stress-athletic injury model' which stating that athletic injury is the interaction between personality, history of stressors, coping resources and cognitive appraisal. In the stress appraisal process, athletes' perceived stress is influence by

above mentioned factors such as personality, history of stressors, and coping rescources. The consequences of this interaction can lead to either attenuate or deteriorate the perceived stress, and eventually cause athletic injury.

Many researchers borrow either stress-burnout or stress-injury concepts to examine the role of perceived stress on athletes' burnout and injury. For example, in examining whether stress and affect as the mediator of hope-stress relationship, Gustafsson and colleagues (Gustafsson, Skoop, Podlog, Lundqvist, & Wagnsson, 2013) administered 238 Swedish soccer players with trait hope, Swedish version of PSS (i.e., PSS-10), positive and negative affect and athlete burnout scales. Results found athletes' hope and burnout were fully mediated by stress and positive affect. Similarly, Tashman,Tenenbaum, and Eklund (2010) sampled 177 college coaches and examined the relationship between coaches' perfectionism and burnout. Results indicated that perceived stress (measured by PSS-14) mediated the relationship between self-evaluative perfectionism and burnout, and a significant direct link to burnout, accounting for 56% of its variance. In contrast, conscientious perfectionism did not directly predict burnout, nor mediate the relationship between perceived stress and conscientious perfectionism. Similar sport burnout studies that using either PSS-10 or PSS-14 can be found in Raedeke and Smith (2004), Smith, Gustafsson and Hassmén (2010), and Gustafsson and Skoog (2012) studies.

In terms of stress- athletic injury relationship, Galambos and colleagues (Galambos, Terry, Moyle, & Locke, 2005) investigated 845 Australian youth athletes' injury rates and its relations with psychological variables. Participants (males=433; Females=412) completed demographica questionnaire, health history, Brunel Mood Scale and PSS-10. Results found mood and stress collectively predicted injury characteristics. In a similar study, Malinauskas (2010) sampled 123 college athletes and administered with social support scale, PSS-10, and life satisfaction scale. Results found greater perceived stress was associated with diminished life satisfaction for major injury athletes than minor injury athletes. Also, the interaction between perceived stress and perceived social support was associated the most with diminished life satisfaction for athletes with major injury.

Although sport researchers used either PSS-10 or PSS-14 in examining their relationships with athlete burnout or injury. The psychometric properties of PSS have never been examined and remained several questions. First, whether two-factor or

unidimensional PSS will be suitable in sports? Second, if we compare athletes and non-athletes will measurement characteristics remains the same? Third, what is the reliability and validity of PSS? Therefore, there are three puroses in this study. First, we intended to examine the factorial structure of the PSS-10 and PSS-14, and internal consistency. Second, we intended to examine the measurement invariance of PSS between athletes and non-athletes. Third, we attempted to examine the construct validity of PSS.

Study 1

The purpose of study1 was threefold: (a) to compare the factorial structure of the PSS-10 and PSS-14; (b) to examine internal consistency of PSS-10 and PSS-14; (c) to examine measurement invariance of PSS between athletes and non-athletes.

Methods

Participants & procedure

359 college student-athletes (males = 233; females = 126) and 242 non-athletes (males=124; females=118) with mean age 20.08 (*SD*=± 1.51) were recruited as participants (M_{age} = 20.10 yrs, *SD* = ± 1.55). The participants were recruited at two sport-colleges and three universities in Taiwan. At the time of data collection, athletic participants were all in their regular training seasons and had been participating in a variety of individual and team sports, such as gymnastics, track and field, golf, weight-lifting, basketball, volleyball, soccer, Tae-kwon-do, badminton and baseball for 8.93 years (SD=3.14) of training and competition experiences. For non-athlete participants the data was collected during their regular classes.

After the approval by a local institutional review board, the first author contacted the coach/teacher of a target team and class, and asked permission to use his/her team as participants. Once the coach/teacher agreed to use his/her team as

participants, we visited target team/class one hour before they finished the regular training/class. Before administering the questionnaire package, the first author explained the general purpose of the study, the method to complete questionnaires, and rights of being a participant. To prevent social desirability effects, we informed participants that this is a study to explore college students' life experiences, and there were no right or wrong answers. Additionally, we asked them to answer the questions as truthfully as possible, and all responses would be confidential. After the briefing, participants who interested in this study then signed a consent form, and completed the demographic questionnaire and 14-item PSS. It took about 15 minutes to complete the questionnaires.

Measurements

Demographic Information

The demographic questionnaire was designed to gather information about participants' age, gender, types of sports, and years of athletic experiences.

Perceived Stress Scale (PSS)

The PSS measures is a self-report measure designed to assess one's perception about the degree of a given situation in daily life is considered stressful (Cohen et al., 1983). The PSS-14 contains seven positively worded 'stress' items (e.g., How often have you felt upset because of something that happened unexpectedly?) and seven negatively worded 'counter-stress' items (e.g., How often have you felt confident about your ability to handle personal problems?). Items are rated on a 5-point Likert scale of occurrence these statements over the past 4 weeks (0 = never, 1 = almost never, 2 = sometimes, 3 = fairly often, 4 = very often). Because reverse-coding may confound the counter-stress factor (Golden-Kreutz et al., 2004), we did not reverse-code the items.

Analytic strategy

All primary statistical testing was conducted in AMOS version 22. Models of the PSS-14 and the PSS-10 (i.e., by deleting item 4, 5, 12, and 13) were first estimated

separately for the 1-factor and 2-factor as Table 1. Overall, model fit was evaluated using the following indices suggested by Hu and Bentler (1999) as follow: the comparative fit index (CFI; study criterion ≥0.950 as ideal and ≥0.90 as the minimum acceptable level), the root mean square error of approximation (RMSEA; study criterion ≤0.080) and the standardized root mean square residual (SRMR; study criterion ≤0.080). To examine the internal consistency of the factors, Cronbach's α coefficient was used as an index.

For testing measurement invariance, we adopted earlier suggestion (Barbosa-Leiker et al., 2011) by following procedures: (a) once the confirmatory factor models for each group established that the overall model was acceptable, a series of analyses to examine measurement invariance were performed sequentially between comparison and nested model; (b) each model was added equality constraints and was tested against the less-constrained model including following indices:

1. Configural invariance (Horn & McArdle, 1992) (also referred to as 'equal form'). This step examined the pattern of salient and non-salient loadings across groups (Vandenberg & Lance, 2000). This step took the measurement model and examined if the theoretical framework of the PSS is the same for athletes and non-athletes.

2. Metric invariance (Horn & McArdle, 1992) (also referred to as 'equal loadings'). This step constrained the factor loadings for like items across groups to determine whether the expected changes in observed values of the indicators per unit change of the construct were equal (Vandenberg & Lance, 2000). This step tested if the relationships of the PSS-14 or PSS-10 items were equivalent for like indicators in athletes and non-athletes.

3. Factor variance/covariance invariance (also referred to as 'equal factor variances'). This step constrained the like factor variances across the groups (Vandenberg & Lance, 2000). If factor variance invariance holds, then the amount of within group variability of the latent factor is equal across groups (Brown, 2006). This step tested whether athletes and non-athletes use equivalent ranges of the latent constructs (stress and counter-stress) to respond to the PSS-14 or PSS-10 items.

4. Error variance invariance. This step constrained error variances across the groups. If the same level of measurement error is present for each item between groups. This step tested whether the error of stress and counter-stress all items were related equivalently across athletes and non-athletes.

For tests of invariance, χ^2 difference tests are typically used to compare nested models. However, the χ^2 difference test may also be influenced by sample size (Chen, Sousa, & West, 2005); thus, a change in comparative fit index (CFI) between comparison and nested models of greater than or equal to -0.010. In addition, we examined the change in root mean square error of approximation (RMSEA) ≥0.015 or a change in standardized root mean square residual (SRMR) ≥0.030 (for loading invariance) and ≥0.010 (for intercept invariance) is recommended as an appropriate criterion indicating a decrement in fit between models (Chen, 2007; Chen, Sousa, & West, 2005; Cheung & Rensvold, 2002). Additionally, a χ^2 difference test for a small difference between models (rather than 0) was also conducted (χ^2critical 0.05; MacCallum, Browne, & Cai, 2006).

Results

Table 1 shows the fit of the 2-factor model of PSS-10 and PSS-14, and for comparison purposes. The 1-factor model whether 14-items (RMSEA = 0.156> 0.080; CFI = 0.483 <0.90; SRMR = 0.157> 0.080) and 10-items (RMSEA = 0.146> 0.080; CFI = 0.685 <0.90; SRMR = 0.146> 0.080) adaptation indicators show the overall pattern did not accept. However, the 2-factor model provided acceptable fit. The 2-factor model displayed PSS-14 the overall pattern did not accept (RMSEA = 0.084 >0.080; CFI = 0.853 <0.90; SRMR = 0.084 >0.080). The PSS-10 (RMSEA = 0.070< 0.080; CFI = 0.929> 0.90; SRMR = 0.060 <0.080) overall fit indices are acceptable adaptation. The CFI and SRMR show acceptable, especially SRMR achieve good adaptation indicators. The results showed that 2-factor PSS-10 model was the best model. Also, it was found Cronbach's α coefficients for 2-factor PSS-10 were .81 (perceived stress) and .71 (counter stress).

(Insert Table 1 about here)

Table 2 shows the athletes and non-athletes 2-factor PSS-10 model of the measurement invariance, M1 was configuration invariance model, M2 metric invariance, M3 variation \ covariance invariance, M4 error variance invariance are shown to have an acceptable adaptation indicators. △CFI indicated that 2-factor PSS-10 model of athletes and non-athletes in M1, M2, M3 measurement invariance model display equivalent (△CFI <0.01), however, M4 shows the residuals are not equal (△CFI> 0.01). We will discuss this later in the discussion.

Study 2

The purpose of study 2 was to examine the construct validity of 2- factor PSS-10, which is the convergent and discriminant validity, via correlational analyses surrounding the relationships among PSS-stress, PSS counter-stress, college student-athletes' life stress, coping self-efficacy and burnout experiences.

(Insert Table 2 about here)

Methods

Participants & procedure

A new sample of the targeted population was recruited. Valid data of 196 student-athletes from ten different universities were collected (males = 139, females = 57, M_{age} =19.88 yrs, SD = 1.35). The recruiting procedure was similar to study 1.

Measurements

The measurements included the Demographic Questionnaire and the 10-item PSS. In addition, the researchers administered the following measures for examining convergent and discriminant validity.

Athlete Burnout (ABQ)

ABQ (Raedeke & Smith, 2001) is a self-reported inventory that assesses athletes' burnout experiences. The initial factor analyses by Raedeke and Smith (2001) revealed that ABQ has three subscales including: (a) 5 items for reduced sense of athletic accomplishment, (b) 5 items for perceived emotional and physical exhaustion, and, (c) 5 item for devaluation of sports participation. Participants identify their athletic burnout experiences using a six-point Likert scale that ranged from 1 (*never*) to 6 (*always*). In the present study, the result of CFA confirmed that the factorial structure was suitable for the data. The Cronbach's α for the three subscales ranged from .63 to .86 and the reliability for all items was .90. To further identify convergent validity it is expected that the burnout scale should be positively correlated with the PSS stress factor because athletes' stress has been identified as leading factor of athlete burnout (Lewis, 1991; Nicholls Backhouse, & McKenna, 2009; Galambo *et al*., 2005; Johnson & Ivarsson, 2011). The PSS counter-stress factor was expected to have negative relation with ABQ.

College Student-Athletes' Life Stress Scale (CSALSS)

The 24-item CSALSS (Lu, Hsu, Chan, Cheen, & Kao, 2012) was used to assess situations that athletes encountered in their daily life and sports, and considered as major stressors in their lives. The questionnaire asked questions such as "I am annoyed with my coach's bias against me." There are eight factors in the 24-item CSALSS including: (a) sports injury, (b) performance demand, (c) coach relationships, (d) training adaptation, (e) interpersonal relationships, (f) romantic relationships, (g) family relationships, and (h) academic requirements. Lu and colleagues (Lu et al., 2012) reported that CSALSS can be categorized into two major

components --- general life stressors (by adding factor e, f, g, h) and sport-specific stressors (by adding factor a,b,c, d). Participants indicated the frequency of the event on a six-point Likert scale ranged from 1 (*Never*) to 6 (*Always*). Cronbach's α of these factors ranged from .69 to .87 and the reliability for all items was .92 in this study, indicating that the result was reliable. Given that CSALSS represents an individual's life stress, the PSS counter-stress factor was expected to have negative with CSALSS, another the PSS stress factor was expected to have positive relation with CSALSS.

Coping Self-Efficacy Scale (CSE)

CSE (Chesney, Neilands, Chambers, Taylor, & Folkman, 2006) is a self-reported inventory that assesses one's confidence in performing coping behaviors when faced with life challenges. The initial factor analyses by Chesney et al. (2006) revealed that CSE has three subscales including (a) problem- focused coping (6 items), (b) stop unpleasant emotions and thoughts (4 items), and (c) get support from friends and family (3 items). The CSE uses an eleven-point Likert scale that ranged from 0 ('cannot do at all'), 5 ('moderately certain can do') to 10 ('certain can do'). In the present study, the Cronbach's α for the three subscales ranged from .70 to .78 and the reliability for all items was .86. It is expected that CSE will be negatively correlated with the PSS stress factor, but positively correlated with PSS counter-stress.

Statistical Analyses

We used SPSS 18.0 to examine the convergent and discriminant validity by examining the correlations among PSS, CSALSS, CSE and ABQ.

Results

The Pearson correlations indicated that the stress factor of PSS negatively correlated with CSE, and positively related with ABQ and CASLSS. Moreover, the counter-stress factor of PSS negatively correlated with ABQ and CASLSS, and positively related with CSE. Most subscales of two-factor PSS-10 showed low to moderate associations with the all scales (see Table3).

(Insert Table 3 about here)

Discussion

In line with past research examining the validity of PSS, the purpose of this study was to examining psychometric properties of PSS in sport settings. Specifically, this study attempted to examine factorial structure, measurement invariance, reliability, and construct validity of PSS. By two studies we found 2-factor PSS-10 has better fit of model. Also, we found 2-factor PSS-10 had appropriate reliability and measurement invariance across athletes and non-athletes. Further, we found 2-factor PSS-10 positively correlated with athletes' life stress and burnout but negatively correlated with coping self-efficacy which indicated appropriate construct validity.

Therefore, the psychometric properties of PSS gain solid supports in the sport context. The results of factorial structure suggest that 2-factor PSS-10 has better measuring quality than unidimensional 2-factor PSS-10 or 2-factor PSS-14. The 2-factor PSS-10 reduces 4 items it allows researchers collect data in a short period of time (Shacham, 1983). Although researchers have different arguments regarding to dimensionality of PSS (Hewitt, Flett, & Mosher, 1992; Mitchell, Crane & Kim, 2008; Örücü, & Demir, 2009) it is for sure that 2-factor PSS receive better support in sport context. The results are consistent with earlier study by Hewitt and colleagues (Hewitt et al., 1992). We suggest future researchers may use 2-factor PSS to examine gender differences, or to compare which variable predict related psychological disorders/constructs the most.

In terms of measurement invariance we found configuration invariance, metric invariance, and variance/covariance invariance are all equivalent except error variance invariance. Therefore, it means that the same level of measurement error for each item between athletes and non-athletes is not the same. However, Lee (2006) suggested that most research that using CFA focus on the equivalence of the factor loadings and factorial covariance. If these indicators meet criteria they can assure that means measurement invariance across observed groups are held, while residual restrain model maybe too critical to be reached. Tabachnick and Fidell (2001) also suggest that when factor loadings and factorial covariance are equivalent across group it is indicated that measurement invariance holds true.

A significant feature of this study is the sample recruited from Taiwanese student-athletes. Although PSS has been validated in different culture such as Spain (Remor, 2006), Swedish (Eskin, & Parr, 1996), Japanese (Mimura, & Griffiths, 2004), Portugal (Ramírez & Hernández, 2007), Turkish (Örücü, & Demir, 2009), or Chines (Leung, Lam, & Chan, 2010), this is the first study using healthy young athletes as participants in examining psychometric properties of PSS and its relationship with athlete burnout, coping self-efficacy and life stress. The results of psychometric validation and measurement invariance can be forwarded to those researchers who interested in psychometric properties of PSS. Further, our construct validity analyses found 2-factor PSS-10 correlated with athlete burnout and life stress, but negatively correlated with coping self-efficacy, are worthy of forwarding these messages to sport professionales. As we all know the young athletes face many challenges in their life (Lewis, 1991). They engage in heavy and intensive training/competition all year round. The intensive and heavy training bring lots of stress for young athletes (Weinberg & Gould, 2015). If coaches and sport professionals fail to monitor athletes' training loading and arrange them with appropriate competition plans they may induce lots of stress and cause burnout. Therefore, teaching young athletes effective coping skills (e.g., fostering social support and time management) and teaching psychological skills (e.g., goal-setting, relaxation, and imagery) are very important because they can help young athletes to cope with stress from training and competitions.

Further, as previous mentioned that most research in sport either using Smith (1986) cognitive-affective model of athletic burnout or Anderson and Williams (1988) stress-athletic injury model to examine the effects of stress on athlete burnout/injury, we suggest future research may use 2-factor PSS-10 in empirical studies. Also, we suggest researchers may extend research beyond effects of stress on athletes' behavior. In stead we suggest that future research may examine the antecedents of perceived stress in sport. For example, sport literature suggests that motivatonal climate created by coaches and team (e.g., ego-involving climate) is one of the major sources of athletes' stress (Hogue, Fry, Fry, & Pressman, 2013). Researchers may use 2-factor PSS-10 in examining how a sport team's motivational climate predict athletes' perceived stress. In addition, it is well-documented that the leadership style implemented by coaches (e.g., autocratic) may produce stress for athletes (Horn, Bloom, Berglund & Packard, 2011). We suggeste esearchers may examine how coaches' leadership and situational conditions (e.g., competition performance) to predict athletes' perceived stress. Further, since athletic world is a challenging settings it is found that some athletes become substance abuse and eating disorder to cope stress (Ansel, 2010). We suggest researchers may use 2-factor PSS-10 as a measuring tool in assessing athletes' perceived stress during off- season, pre-season, and after-season in order to understand their stress level and provided with appropriate interventions.

There are several limitations that needs to be addressed. First, our sample are all recruted from Dividion I college student-athletes whether our results could be generalized to other athletes, such as professional athletes or junior athletes, need to be further examined. Additionally, the data were collected from Taiwanese student-athletes; hence the results may not be generalizable to different cultures. We recommend researchers adopt similar approaches to test measurement invariance in different culture and populations. Further, although we examined PSS-10 by internal consistency coefficients we did not examine other reliabilities indices such as composite reliability and test retest reliability. Future study may examine 2-factor PSS-10 reliability by these indices to provide more evidences of reliability.

Conclusion

To acknowlege that PSS is a widely used measure in assessing stress, we have conducted two studies to examine the factor structure, measurement invariance between athletes and non-athletes, reliability, and construct validity in the sport context. Results indicated that 2-factor PSS-10 can be an ideal measure for the research in sport. We suggest future research may use 2-factor PSS-10 in conducting various stress research.

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594

Table 1(on next page)

including 3 tables of factor structure measurement invariance

Table 1 The PSS one-factor -model and two-factor- model comparing table

	df	χ^2	CFI	RMSEA (LO90-HI90)	SRMR
1-factor					
PSS14	77	748.265**	0.483	0.156 (.146-.166)	0.157
PSS10	35	302.092**	0.685	0.146 (.131-.161)	0.123
2-factor					
PSS14	76	266.244**	0.853	0.084 (.073-.095)	0.084
PSS10	34	94.008**	0.929	0.070 (.054-.087)	0.060

p<.001**

Table 2 the athletes and non-athletes two-factor PSS-10 model of the measurement invariance table

		df	χ^2	CFI	RMSEA	SRMR	ΔCFI
athlete v.s. non-athlete	M1(configural)	68	178.961**	0.927	0.052	0.059	
	M2(metric)	76	193.196**	0.923	0.051	0.061	-0.004
	M3(variance\ covariance)	79	196.174**	0.923	0.050	0.062	0.000
	M4(residual)	89	265.011**	0.884	0.057	0.061	-0.039

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Table 3 Descriptive Statistics, Reliability Coefficients, and the Correlation Matrix for Study Variables

	1	2	3	4	5
1.Stress	.77				
2.Count-stress	-.126	**.68**			
3.ABQ	.370*	-.287*	**.90**		
4.CSE	-.285*	.408*	-.300*	**.86**	
5.CASLSS	.453*	-.183*	.459*	-.290*	**.92**
Mean	2.014	2.030	3.033	6.648	2.820
SD	0.670	0.703	0.854	1.365	0.794

Note: *** *p* < .05; Cronbach alphas are presented on the diagonal as bold font;**