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The prognostic value of tumor length to resectable esophageal squamous cell carcinoma: a retrospective study

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ABSTRACT

Background. The current TNM classification system does not consider tumor length for patients with esophageal carcinoma (EC). This study explored the effect of tumor length, in addition to tumor depth and lymph node involvement, on survival in patients with esophageal squamous cell carcinoma (ESCC).

Methods. A total of 498 ESCC patients who underwent surgical resection as the primary treatment were selected in the retrospective study. Pathological details were collected, which included tumor type, TNM stage, differentiation. Other collected information were: the types of esophageal resection, ABO blood group, family history and demographic and lifestyle factors. A time-dependent receiver operating characteristic (ROC) curve and a regression tree for survival were used to identify the cut-off point of tumor length, which was 3 cm. Univariate and multivariate Cox proportional hazard regression models were used to identify the prognostic factors to ESCC.

Results & Discussion. The 1-, 3-, 5-year overall survival rates were found to be 82.5%, 55.6%, and 35.1%, respectively. Patients who had larger tumor length (>3 cm) had a higher risk for death than the rest patients. From the univariate Cox proportional hazards regression model, the overall survival rate was significantly influenced by the depth of the tumor and lymph node involvement (either as dummy or continuous variables), Sex, and tumor length. Using these four variables in the multivariate Cox proportional hazard regression model, we found that the overall survival was significantly influenced by all variables except Sex. Therefore, in addition to the depth of the tumor and lymph node involvement (as either dummy or continuous variables), the tumor length is also an independent prognostic factor for ESCC. The overall survival rate was higher in a group with smaller tumor length (≤ 3 cm) than those patients with larger tumor length (>3 cm), no matter what the tumor stage was.

Conclusion. The tumor length was found to be an important prognostic factor for ESCC patients without receiving neoadjuvant therapy. The modification of EC staging system may consider tumor length to better predict ESCC survival and identify higher risk patients for postoperative therapy.

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INTRODUCTION

Esophageal cancer is the eighth most common cancer worldwide and the sixth most common cause of death from cancer (*Ferlay et al., 2010*). Esophageal adenocarcinoma (EAC) and squamous cell carcinoma (ESCC) are the predominant histological types. In the western countries, the majority of the esophageal cancers were EAC, while in the eastern countries, ESCC were predominant.

In the current American Joint Committee on Cancer (AJCC) staging system for esophageal tumors, stage depends on the depth of the tumor (T classification), lymph node involvement (N classification), and distant metastasis (M classification), without consideration on tumor length (Edge & Compton, 2010). Studies have shown that pathologic tumor length had prognostic value to EAC (Bolton et al., 2009; Gaur et al., 2011; Yendamuri et al., 2009). In China, a few studies were conducted to show the association between the overall survival (OS) rate and the tumor length, among the survival of ESCC patients (Feng, Huang & Zhao, 2013; Ma et al., 2015; Wang et al., 2011; Wang et al., 2012). However, the optimal cut-off points for tumor length to predict OS has no consensus. In China, ESCC is the predominant histological subtype and EAC remains rare. Esophageal cancer is the fourth most frequently diagnosed cancer and the fourth leading cause of cancer death in China, with an estimated 259,235 new cases and 211,084 deaths in 2008 (Lin et al., 2013). Data from the Chinese National Health and Family Planning Commission databases showed that from 2005 to 2012 the crude incidence and mortality rate for esophageal cancer increased from 19.55 to 38.44/100,000 and 12.62 to 16.77/100,000, respectively (Tang et al., 2015). Survival rate from esophageal cancer is low, despite improvements in care. This leads to great disease burden.

More accurate methods of determining prognosis in patients with esophageal cancer are needed. It will benefit patients with specific and appropriate treatment options. The purpose of this study is to explore the prognostic value of pathological tumor length in the context of other prognostic factors and obtain the optimum cut-off point for tumor length.

MATERIALS & METHODS

Patients

From August 2007 to July 2014, 552 patients were identified to have undergone surgical resection of esophageal tumor with curative intent at Shandong Provincial Hospital affiliated to Shandong University. The patients who had histology other than squamous cell carcinoma were excluded from this study (n = 33). The patients who received neoadjuvant radiotherapy or chemotherapy and noncurative resection (tumor-free margin) were also excluded from this study (n = 21). A total of 498 patients with ESCC were included for analyses. The preoperative evaluation was performed based on a complete medical history, physical examination, endoscopic ultrasound and thoracoabdominal computed

tomography scanning. Either the left or the right thoracotomy was included in this study that mainly depended on the tumor location and possible mediastinal lymph nodes involvement. The surgeons taking into account the preoperative evaluation decided the extent of the esophageal resection and lymph nodes dissection during the operation. Then the specimens were sent for pathology examination after preservation in 10% formalin. The tumor length, differentiation, T classification, number of positive and removed lymph nodes were recorded according to the pathology reports. The TNM staging was performed according to the AJCC 7th edition guidelines. Other collected information were: the types of esophageal resection, ABO blood group, family history and demographic and lifestyle factors. After surgery, all the patients were followed up regularly by telephone interviews. Complete follow-up information for all patients until death or July 2014 was available. All subjects gave written informed consent to the study protocol, which was approved by the Ethical Committees of Shandong Provincial Hospital affiliated to Shandong University (No. 2015-055). The study was carried out in accordance with the approved guidelines.

Statistical analyses

Two methods were conducted to identify the optimum cut-off point for pathologically measured tumor length. First, we used a time-dependent receiver operator characteristic (ROC) curve to evaluate the tumor length in order to predict ESCC survival. Using the ROC curve, we identified the sensitivity and specificity at each tumor length point and the optimal tumor length cut-off point. This was found to be 3.1 cm (Fig. S1A)—based on the highest Youden index (sensitivity + specificity -1) (*Heagerty, Lumley & Pepe, 2000*). Then, we used a regression tree survival analyses for pathologically measured tumor length (Gaur et al., 2011) and found it to be 2.9 cm (Fig. S1B). The mean of these two results -3.0 cm - was used as the threshold value for optimal cut-off. Based on this value, patients were divided into two groups: one with tumor length ≤ 3 cm and the other with tumor length >3 cm. Overall survival rate was estimated by the Kaplan–Meier method using the date of esophageal resection as the starting point and the date of death or last follow-up as the endpoint. The Kaplan-Meier survival curves between two tumor length groups were compared using the log-rank test. The association between risk factors and survival was performed using a univariate Cox proportional hazards (Coxph) regression model. Hazard ratios (HRs) with 95% confidence intervals (CIs) were used to quantify the strength of the association. Risk factors with p-values ≤ 0.1 in univariate analyses were entered into a multivariate Coxph regression model. The results of the multivariate Coxph regression model were visualized in graphs. Tests and graphical diagnostics for proportional hazards were based on the scaled Schoenfeld residuals, which calculated tests of the proportional hazards assumption for each covariate by correlating the corresponding set of scaled Schoenfeld residuals with a suitable transformation of time (*Hess*, 1995). All statistical calculations were conducted using R software (version 3.2.3) (*Team, 2014*). Different packages were used for analysis: survivalROC package in time-dependent ROC curve, rpart.plot and survival packages in regression tree for survival, survival package in overall survival rate estimated by Kaplan-Meier method and the log-rank test, univariate and

multivariate Coxph regression analyses, tests and graphical diagnostics for proportional hazards.

RESULTS

The characteristics of the 498 patients are summarized in Table 1. The median follow-up time for all patients was 47.2 months (IQR: 6.5–64.5 months). The median age was 59 years (range 38-81 years); majority of patients were male (78.9%). Over 50% of the patients had over 10 years of smoking or alcohol intake. Over 80% patients' relatives (first-degree) had history of cancer, which includes oral, esophagus, lung, stomach, nasopharynx, liver, pancreatic, rectal, prostate cancer, etc. Majority patients (63%) had tumor with a length larger than 3 cm.

Figure 1 showed the Kaplan–Meier survival curves. The 1-, 3-, 5-year OS rates were 82.5%, 55.6%, and 35.1%, respectively (Fig. 1A). Figure 1B shows a comparison of survival curves between two groups of patients based on their pathologically measured tumor lengths: larger or less and equal to 3 cm. Patients who had larger tumor length showed a lower survival curve than the patients with smaller tumor length. Specifically, the 3-year and 5-year OS were 44.7% and 26.4% for the group with larger tumor length (>3cm); whereas they were 73.1% and 49.3% for the group with smaller tumor length (\leq 3 cm). This difference is significant (p < 0.001)).

Univariate Coxph regression model showed that the following variables had no statistical association with survival: age, ABO blood group, family history of first degree relatives with cancer, types of esophageal resection, histologic grade, length of smoking time and length of alcohol drinking time (either as dummy variable or continuous variable). However, T status and N status, either as dummy variable or continuous variable, and Sex, tumor length influenced significantly the OS (Table 2).

Based on the results from Univariate Coxph regression, a multivariate Coxph regression model was constructed using variables of Sex, T stauts, N status and tumor length. The restults showed that except Sex, the rest three variables have significant association with survival. Therefore, T status and N status (either as dummy variable or continuous variable), and tumor length remained as independent prognostic factors.

Based on the multivariate Coxph regression model, patients with larger tumor length (>3cm) had an increased risk of death compared with patients with smaller tumor length (\leq 3 cm) after controlling T status and N status. The hazard ratios were 1.52 (95% CI [1.06–2.22]) when treating T and N status as dummy variables and 1.53 (95% CI [1.05–2.21]) when treating T and N status as continuous variables. When T status and N status were treated as continuous variables, the hazard ratios become 1.60 (95% CI [1.21–2.13]) when T status increased by 1 degree after controlling N status and tumor length, and 1.37 (95% CI [1.18–1.60]) when N status increased by 1 degree after controlling T status and tumor length (Table 3).

Figure 2 showed the fitted results of multivariate Coxph regression model. The top four figures showed the variation of T-status (increasing tumor depth) while holding N-status constant at N_0 ; the bottom four figures showed the variation of N-satus (increasing lymph

Table 1 Characteristics of patients.	
Characteristic	No. of patients (%)
Sex	
Male	393(78.9)
Female	105(21.1)
Age, y	
Median	59
Range	38-81
ABO blood group	
А	152(30.5)
В	136(27.3)
0	110(22.1)
AB	64(12.9)
Unkown	36(7.2)
Length of smoking time (year)	
None	157(31.5)
1–10	35(7)
11–20	76(15.3)
21–30	130(26.1)
30-	100(20.1)
Length of alcohol drinking time (year)	
None	211(42.4)
1–10	19(3.8)
11–20	101(20.3)
21–30	100(20.1)
30-	67(13.5)
Family history in first degree relatives with any cancer	
None	423(84.9)
Yes	75(15.1)
Types of esophageal resection	
Ivor-Lewis esophagectomy	143(28.7)
Three-field esophagectomy	47(9.4)
Minimally invasive esophagectomy	17(3.4)
Left thoracic	275(55.2)
Unkown	16(3.2)
Tumor length (cm)	
<u>≤</u> 3	185(37.1)
>3	313(62.9)
Tumor classification	_ /
Tis	7(1.4)
	47(9.4)
T2	105(21.1)
Т3	314(63.1)
T4	21(4.2)

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  Table 1 (continued)
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Characteristic	No. of patients (%)
Unknown	4(0.8)
Lymph node classification	
N0	294(59)
N1	102(20.5)
N2	76(15.3)
N3	22(4.4)
Unknown	4(0.8)
Grade	
Well differentiated	16(3.2)
Moderately differentiated	396(79.5)
Poorly differentiated	66(13.3)
Undifferentiated	1(0.2)
Unknown	19(3.8)



Figure 1 Kaplan-Meier survival curves (S(t)). (A) S(t) vs. time of after surgery with 95% confidence interval for overall ESCC patients. (B) S(t) vs. time of after surgery for ESCC patients stratified by tumor length (3 cm)

node involvement) while keeping T-satus constant at T₃. Each figure showed two curves with the top one (blue) corresponding to smaller tumor length and the low curve (green) to larger tumor length. These choices of N₀ and T₃ are based on the fact that most patients in the study were either in T₃ (63.1%) status or in N₀ (59%) status (Table 1). We used the multivariable Coxph regression model to fit the survival trend of patients with T₃N₀₋₃ status and T₁₋₄N₀ status in the two different tumor length groups (tumor length \leq 3 cm versus >3 cm). The results showed that the OS rate was higher in the smaller tumor length (\leq 3 cm) group independent of the tumor stage. The OS rate declined as either T-status or N-status increases independent of the tumor length.

Variable	Univariate analysis		
	HR (95% CI) p		
Sex			
Male	1.00		
Female	0.64 (0.43-0.94)	0.023	
Age,y	1.00 (0.98–1.02)	0.836	
Blood type			
А	1.00		
В	1.42 (0.83–2.44)	0.198	
0	1.41 (0.81–2.45)	0.225	
AB	1.29 (0.73–2.28)	0.378	
Length of smoking time (year) (As dummy variable)			
None	1.00		
1–10	1.03 (0.57–1.86)	0.910	
11–20	0.95 (0.60–1.50)	0.818	
21–30	0.98 (0.66–1.44)	0.906	
30-	0.94 (0.62–1.41)	0.761	
Length of smoking time (year) (As continuous variable)	1.00 (0.99–1.01)	0.644	
Length of alcohol drinking time (year) (As dummy variable)			
None	1.00		
1–10	1.35 (0.65–2.80)	0.424	
11–20	1.11 (0.75–1.64)	0.593	
21–30	1.04 (0.70–1.55)	0.842	
30-	1.26 (0.81–1.95)	0.303	
Length of alcohol drinking time (year) (As continuous variable)	1.00 (0.99–1.01)	0.596	
Family history in first degree relatives with any cancer			
None	1.00		
Yes	0.94 (0.60–1.49)	0.806	
Types of esophageal resection			
ILE, MIE, 3-field	1.00		
Left thoracic	0.81 (0.60–1.09)	0.162	
Tumor length (cm)			
<u>≤</u> 3	1.00		
>3	2.35 (1.68–3.27)	< 0.001	
Tumor classification (As dummy variable)	2.14 (1.69–2.75)		
Tis/T ₁	1.00		
T ₂	4.94 (1.76–13.82)	< 0.001	
T ₃	8.45 (3.12–22.87)	< 0.001	
T_4	17.93 (5.22–61.54)	< 0.001	
Tumor classification (As continuous variable)			

Table 2 Univariate Cox proportional regression analyses results.

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Table 2 (continued)

Variable	Univariate analysis		
	HR (95% CI)	p	
Lymph node classification (As dummy variable)			
N	1.00		
N ₁	2.17 (1.51–3.11)	< 0.001	
N ₂	3.26 (2.23–4.76)	< 0.001	
N ₃	2.73 (1.45–5.13)	< 0.001	
Lymph node classification (As continuous variable)			
Grade			
Well differentiated	1.00		
Moderately differentiated	1.07 (0.63–1.82)	1.823	
Poorly differentiated	0.33 (0.08–1.46)	1.455	

Fable 3	Multivariate Cox	proportional	regression	analyses resu	lts.
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Variable	Dummy variable result		Continuous variable result	
	HR(95% CI)	<i>p</i> -value	HR(95% CI)	<i>p</i> -value
Sex				
Male	1.00		1.00	
Female	0.74 (0.49–1.10)	0.138	0.72 (0.49–1.08)	0.11
Tumor length (cm)				
<u>≤</u> 3				
>3	1.52 (1.05–2.21)	0.026	1.53 (1.06–2.22)	0.025
Tumor classification			1.60 (1.21–2.13)	< 0.001
Tis/T ₁	1.00			
T ₂	4.02 (1.43–11.33)	0.009		
T ₃	4.82 (1.72–13.50)	0.003		
T_4	7.22 (1.93–27.00)	0.003		
Lymph node classification			1.37 (1.18–1.60)	< 0.001
N ₀	1.00			
N ₁	1.67 (1.16–2.42)	0.006		
N ₂	2.21 (1.48-3.30)	< 0.001		
N ₃	2.14 (1.13-4.07)	0.02		

DISCUSSION

Our study showed that patients with larger tumor length than 3 cm had a higher risk for death than those with smaller tumor length. This indicated that tumor length, in additional to the depth of the tumor and lymph node involvement, is also an independent prognostic factors to ESCC. This finding can be important to the improvement of ESCC prognosis, giving the current poor situation on prognosis of ESCC. In fact, the results in ESCC are inconsistent so far. The worldwide studies have been trying to identify more accurate methods to determine prognosis in patients. This study provides the needed evidence for the prognostic factors to improve the ESCC.





The association between tumor length and OS rate has been controversial. Two studies conducted in Japan and one study in Germany showed that tumor length was associated with survival rate of ESCC patients in univariate analyses, but not in the multivariate analyses (*Bollschweiler et al., 2006; Tachibana et al., 2002; Tachibana et al., 1999*). In these studies, some patients received neoadjuvant therapy, which may have biased the measurement of tumor length because of tumor regression. Another study conducted in Japan also showed that tumor length was not an independent prognostic factor to OS of ESCC patients (*Igaki et al., 2001*). However, the study sample was small—only 96 ESCC cases. In addition, the tumor length was measured with X-ray examination which is known to be a rough measuring method. Further more, the study did not indicate how to figure out the cut-off point. A study conducted in America showed that tumor length had no association with survival of ESCC patients (*Yendamuri et al., 2009*). Here again, the statistical power was small because of small sample size—only 30 cases.

Some studies conducted in China showed that tumor length could be as a prognostic factor in ESCC (*Feng, Huang & Zhao, 2013*; *Ma et al., 2015*; *Wang et al., 2011*). However, the optimum cut-off point of the tumor length was different. One study in Taibei indicated that 3 cm was the cut-off point for prediction (*Wang et al., 2011*); while the other study in Tianjin and the third one in Zhejiang chosed 4 cm as the cut-off point (*Feng, Huang & Zhao, 2013*; *Ma et al., 2015*). A large study conducted in the US showed that tumor length was an independent predictor of mortality and a 3-cm tumor length cut-off was proposed. This agrees with our result using regression tree for survival and time-dependent ROC curve methods, and multivariate Coxph regression.

Our study indicated that T classification and N classification both were independent prognostic factors in EC. This is consistent with the known results that are acknowledged





in the current TNM classification (*Peyre et al., 2008*; *Wijnhoven et al., 2007*). Fig. 3 showed the dependence of Beta for T, N, and tumor length on time. Beta was the scaled Schoenfeld residuals from tests and graphical diagnostics for proportional hazards. If the proportional hazards assumption is true, beta(t) will be a horizontal line centering around 0. Since Beta for each chosen variable was independent of time and basically zero, we concluded that each prognostic factor and the whole model did not violate the proportional hazards assumption (Table S1). Our results showed that the survival rate decreased as either tumor depth or lymph node involvement increased after controlling the other variables. The multivariable Coxph regression model indicated that tumor length was positively associated with survival after adjusting the T and N status. Therefore, the optimal cut-off point of tumor length applied to different T or N status.

Studies have proven that smoking and alcohol intake increased the risk of ESCC (*Chung et al., 2010; Islami et al., 2011*), and positive synergistic effect of alcohol and tobacco use for ESCC existed (*Prabhu, Obi & Rubenstein, 2014*). Our study explored whether smoking and alcohol intake before diagnosis had influence on ESCC survival. However, the study showed that alcohol and tobacco intake had no association with survival. The data on smoking and alcohol were consumption duration not the intensity of tobacco and the amount of alcohol usage, which was an important limitation. One study showed that heavy smoking is strongly associated with poor prognosis (*Shitara et al., 2010*).

In our study, we did not find the association between ESCC and either family history of any cancer in first-degree relatives or family history of EC in first-degree relatives. Very few studies explore the prognostic value of family history of cancer to ESCC. Other studies have shown that family history of any cancer in first-degree relatives was one of the risk factor for ESCC (*Bhat et al.*, 2015; *Gao et al.*, 2009), and *Chen et al.* (2015) proved that family history of EC was the risk factor of ESCC. In addition, Jiang et al. indicated that family history of EC was one of the independent prognostic factors for survival of ESCC (*Yuequan, Shifeng & Bing, 2010*). While there was no obvious evidence to confirm ESCC with positive family history was a hereditary predisposition (*Yuequan, Shifeng & Bing, 2010*). More studies are needed to confirm this association.

In the study, we did not find the association between ABO blood group and survival. This is a controversial area. Several studies explored the ability of ABO blood type to predict prognosis of ESCC, but the results were not consistent. *Wang et al.* (2015) showed that ABO blood group had no association with EC survival. *Qin et al.* (2015) indicated that blood type was not an independent prognostic factor for all ESCC patients, but was an independent prognostic factor for ESCC patients with negative lymph nodes. *Yang, Huang & Feng* (2014) found that ABO blood group was an independent prognostic factors and blood group O had a worse OS than non-O blood groups. The conflicting results could be due to the heterogeneity of study populations or limited study sizes. Although there are several hypotheses which may explain the relationships observed between ABO blood group and cancer, the direct biologic mechanisms underlying the association are inconclusive. Therefore, whether ABO blood group can be used to predict ESCC survival or not still needs to be confirmed.

Our study indicated that tumor length was a prognostic factor for ESCC patients. However, limitations still existed in our study. The study design was a retrospective observational study. As known, generally the evidence level of retrospective study is relatively low compared to that of prospective study. In addition, the sample size of the study was relatively small, which introduced the relatively larger 95% CI of HR (Table 3). In the future, we plan to conduct a multi-center prospective study to verify and quantity the results more precisely.

CONCLUSIONS

In this retrospective study of 498 patients, we have found that tumor length was an important prognostic factor for ESCC patients without receiving neoadjuvant therapy, in addition to the tumor depth and lymph node involvement. As with lung cancer and breast cancer, tumor size may be considered in the modification of EC staging system to better predict ESCC survival and identify high risk patients for postoperative therapy. Further prospective study with larger sample may be needed to confirm our results.

ADDITIONAL INFORMATION AND DECLARATIONS

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Competing Interests

The authors declare there are no competing interests.

Author Contributions

- Xiangwei Zhang conceived and designed the experiments, analyzed the data, wrote the paper, prepared figures and/or tables.
- Yang Wang conceived and designed the experiments, analyzed the data, prepared figures and/or tables.
- Cheng Li, Yuanzhu Jiang, Guoyuan Ma, Guanghui Wang and Wei Dong contributed reagents/materials/analysis tools.
- Jing Helmersson contributed reagents/materials/analysis tools, reviewed drafts of the paper.
- Shaowei Sang and Jiajun Du conceived and designed the experiments, reviewed drafts of the paper.

Human Ethics

The following information was supplied relating to ethical approvals (i.e., approving body and any reference numbers):

Ethical Committees of Shandong Provincial Hospital affiliated to Shandong University (No. 2015-055).

Data Availability

The following information was supplied regarding data availability: The raw data has been supplied as a Supplementary File.

Supplemental Information

Supplemental information for this article can be found online at http://dx.doi.org/10.7717/ peerj.2943#supplemental-information.

REFERENCES

- Bhat GA, Shah IA, Rafiq R, Nabi S, Iqbal B, Lone MM, Islami F, Boffetta P, Dar NA. 2015. Family history of cancer and the risk of squamous cell carcinoma of oesophagus: a case-control study in Kashmir, India. *British Journal of Cancer* 113:524–532 DOI 10.1038/bjc.2015.218.
- Bollschweiler E, Baldus SE, Schroder W, Schneider PM, Holscher AH. 2006. Staging of esophageal carcinoma: length of tumor and number of involved regional lymph nodes. Are these independent prognostic factors? *Journal of Surgical Oncology* 94:355–363 DOI 10.1002/jso.20569.
- Bolton WD, Hofstetter WL, Francis AM, Correa AM, Ajani JA, Bhutani MS, Erasmus J, Komaki R, Maru DM, Mehran RJ, Rice DC, Roth JA, Vaporciyan AA, Walsh GL, Swisher SG. 2009. Impact of tumor length on long-term survival of pT1 esophageal adenocarcinoma. *Journal of Thoracic and Cardiovascular Surgery* 138:831–836 DOI 10.1016/j.jtcvs.2009.02.003.
- Chen T, Cheng H, Chen X, Yuan Z, Yang X, Zhuang M, Lu M, Jin L, Ye W. 2015. Family history of esophageal cancer increases the risk of esophageal squamous cell carcinoma. *Scientific Reports* 5: Article 16038 DOI 10.1038/srep16038.

- Chung CS, Lee YC, Wang CP, Ko JY, Wang WL, Wu MS, Wang HP. 2010. Secondary prevention of esophageal squamous cell carcinoma in areas where smoking, alcohol, and betel quid chewing are prevalent. *Journal of the Formosan Medical Association* 109:408–421 DOI 10.1016/S0929-6646(10)60072-1.
- Edge SB, Compton CC. 2010. The American Joint Committee on Cancer: the 7th edition of the AJCC cancer staging manual and the future of TNM. *Annals of Surgical Oncology* 17:1471–1474 DOI 10.1245/s10434-010-0985-4.
- Feng JF, Huang Y, Zhao Q. 2013. Tumor length in elderly patients with esophageal squamous cell carcinoma: is it a prognostic factor? Upsala Journal of Medical Sciences 118:145–152 DOI 10.3109/03009734.2013.792887.
- Ferlay J, Shin HR, Bray F, Forman D, Mathers C, Parkin DM. 2010. Estimates of worldwide burden of cancer in 2008: GLOBOCAN 2008. *International Journal of Cancer* 127:2893–2917 DOI 10.1002/ijc.25516.
- Gao Y, Hu N, Han X, Giffen C, Ding T, Goldstein A, Taylor P. 2009. Family history of cancer and risk for esophageal and gastric cancer in Shanxi, China. *BMC Cancer* 9:269 DOI 10.1186/1471-2407-9-269.
- Gaur P, Sepesi B, Hofstetter WL, Correa AM, Bhutani MS, Watson TJ, Swisher SG, Group MDAEC, University of Rochester School of M, Dentistry Foregut G.
 2011. Endoscopic esophageal tumor length: a prognostic factor for patients with esophageal cancer. *Cancer* 117:63–69 DOI 10.1002/cncr.25373.
- Heagerty PJ, Lumley T, Pepe MS. 2000. Time-dependent ROC curves for censored survival data and a diagnostic marker. *Biometrics* 56:337–344 DOI 10.1111/j.0006-341X.2000.00337.x.
- Hess KR. 1995. Graphical methods for assessing violations of the proportional hazards assumption in Cox regression. *Statistics in Medicine* 14:1707–1723 DOI 10.1002/sim.4780141510.
- Igaki H, Kato H, Tachimori Y, Sato H, Daiko H, Nakanishi Y. 2001. Prognostic evaluation for squamous cell carcinomas of the lower thoracic esophagus treated with three-field lymph node dissection. *European Journal of Cardio-Thoracic Surgery* 19:887–893 DOI 10.1016/S1010-7940(01)00701-1.
- Islami F, Fedirko V, Tramacere I, Bagnardi V, Jenab M, Scotti L, Rota M, Corrao G, Garavello W, Schuz J, Straif K, Negri E, Boffetta P. 2011. Alcohol drinking and esophageal squamous cell carcinoma with focus on light-drinkers and neversmokers: a systematic review and meta-analysis. *International Journal of Cancer* 129:2473–2484 DOI 10.1002/ijc.25885.
- Lin Y, Totsuka Y, He Y, Kikuchi S, Qiao Y, Ueda J, Wei W, Inoue M, Tanaka H. 2013. Epidemiology of esophageal cancer in Japan and China. *Journal of Epidemiology* 23:233–242 DOI 10.2188/jea.JE20120162.
- Ma MQ, Yu ZT, Tang P, Jiang HJ, Zhao XJ, Zhang JG, Qu DW, Jin QW, Zhang XZ.
 2015. Is tumor length a prognostic indicator for esophageal squamous cell carcinoma? A single larger study among Chinese patients. *International Journal of Clinical and Experimental Pathology* 8:5008–5016.

- Peyre CG, Hagen JA, DeMeester SR, Altorki NK, Ancona E, Griffin SM, Holscher A, Lerut T, Law S, Rice TW, Ruol A, Van Lanschot JJ, Wong J, DeMeester TR. 2008. The number of lymph nodes removed predicts survival in esophageal cancer: an international study on the impact of extent of surgical resection. *Annals of Surgery* 248:549–556 DOI 10.1097/SLA.0b013e318188c474.
- Prabhu A, Obi KO, Rubenstein JH. 2014. The synergistic effects of alcohol and tobacco consumption on the risk of esophageal squamous cell carcinoma: a meta-analysis. *American Journal of Gastroenterology* 109:822–827 DOI 10.1038/ajg.2014.71.
- Qin J, Wu SG, Sun JY, Lin HX, He ZY, Li Q. 2015. Effect of blood type on survival of Chinese patients with esophageal squamous cell carcinoma. *Onco Targets Ther* 8:947–953 DOI 10.2147/OTT.S81936.
- Shitara K, Matsuo K, Hatooka S, Ura T, Takahari D, Yokota T, Abe T, Kawai H, Tajika M, Kodaira T, Shinoda M, Tajima K, Muro K, Tanaka H. 2010. Heavy smoking history interacts with chemoradiotherapy for esophageal cancer prognosis: a retrospective study. *Cancer Science* 101:1001–1006 DOI 10.1111/j.1349-7006.2009.01466.x.
- Tachibana M, Dhar DK, Kinugasa S, Yoshimura H, Fujii T, Shibakita M, Ohno S, Ueda S, Kohno H, Nagasue N. 2002. Esophageal cancer patients surviving 6 years after esophagectomy. *Langenbeck's Archives of Surgery* 387:77–83 DOI 10.1007/s00423-002-0290-1.
- Tachibana M, Kinugasa S, Dhar DK, Kotoh T, Shibakita M, Ohno S, Masunaga R, Kubota H, Kohno H, Nagasue N. 1999. Prognostic factors after extended esophagectomy for squamous cell carcinoma of the thoracic esophagus. *Journal of Surgical Oncology* 72:88–93 DOI 10.1002/(SICI)1096-9098(199910)72:2<88::AID-JSO9>3.0.CO;2-V.
- Tang WR, Chen ZJ, Lin K, Su M, Au WW. 2015. Development of esophageal cancer in Chaoshan region, China: association with environmental, genetic and cultural factors. *International Journal of Hygiene and Environmental Health* 218:12–18 DOI 10.1016/j.ijheh.2014.10.004.
- **Team RC. 2014.** *R: A language and environment for statistical computing.* Vienna: R Foundation for Statistical Computing.
- Wang BY, Goan YG, Hsu PK, Hsu WH, Wu YC. 2011. Tumor length as a prognostic factor in esophageal squamous cell carcinoma. *Annals of Thoracic Surgery* 91:887–893 DOI 10.1016/j.athoracsur.2010.11.011.
- Wang BY, Liu CY, Lin CH, Hsu PK, Hsu WH, Wu YC, Cheng CY. 2012. Endoscopic tumor length is an independent prognostic factor in esophageal squamous cell carcinoma. *Annals of Surgical Oncology* 19:2149–2158 DOI 10.1245/s10434-012-2273-y.
- Wang W, Liu L, Wang Z, Wei M, He Q, Ling T, Cao Z, Zhang Y, Wang Q, Shi M. 2015. Impact of ABO blood group on the prognosis of patients undergoing surgery for esophageal cancer. *BMC Surgery* 15:106 DOI 10.1186/s12893-015-0094-1.
- Wijnhoven BP, Tran KT, Esterman A, Watson DI, Tilanus HW. 2007. An evaluation of prognostic factors and tumor staging of resected carcinoma of the esophagus. *Annals of Surgery* 245:717–725 DOI 10.1097/01.sla.0000251703.35919.02.

- Yang X, Huang Y, Feng JF. 2014. Is there an association between ABO blood group and overall survival in patients with esophageal squamous cell carcinoma? *International Journal of Clinical and Experimental Medicine* 15:2214–2218.
- Yendamuri S, Swisher SG, Correa AM, Hofstetter W, Ajani JA, Francis A, Maru D, Mehran RJ, Rice DC, Roth JA, Walsh GL, Vaporciyan AA. 2009. Esophageal tumor length is independently associated with long-term survival. *Cancer* 115:508–516 DOI 10.1002/cncr.24062.
- Yuequan J, Shifeng C, Bing Z. 2010. Prognostic factors and family history for survival of esophageal squamous cell carcinoma patients after surgery. *Annals of Thoracic Surgery* 90:908–913 DOI 10.1016/j.athoracsur.2010.05.060.