

An extensive comparison of species-abundance distribution models

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A number of different models have been proposed as descriptions of the species-abundance distribution (SAD). Most evaluations of these models use only one or two models, focus only a single ecosystem or taxonomic group, or fail to use appropriate statistical methods. We use likelihood and AIC to compare the fit of four of the most widely used models to data on over 16,000 communities from a diverse array of taxonomic groups and ecosystems. Across all datasets combined the log-series, Poisson lognormal, and negative binomial all yield similar overall fits to the data. Therefore, when correcting for differences in the number of parameters the log-series generally provides the best fit to data. Within individual datasets some other distributions performed nearly as well as the log-series even after correcting for the number of parameters. The Zipf distribution is generally a poor characterization of the SAD.

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10 **Abstract**

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13 ecosystem or taxonomic group, or fail to use appropriate statistical methods. We use likelihood and
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16 series, Poisson lognormal, and negative binomial all yield similar overall fits to the data. Therefore,
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18 best fit to data. Within individual datasets some other distributions performed nearly as well as the
19 log-series even after correcting for the number of parameters. The Zipf distribution is generally a
20 poor characterization of the SAD.

21 **Introduction**

22 The species abundance distribution (SAD) describes the full distribution of commonness and rarity
23 in ecological systems. It is one of the most fundamental and ubiquitous patterns in ecology, and
24 exhibits a consistent general form with many rare species and few abundant species occurring
25 within a community. The SAD is one of the most widely studied patterns in ecology, leading to
26 a proliferation of models that attempt to characterize the shape of the distribution and identify
27 potential mechanisms for the pattern (see McGill et al. 2007 for a recent review of SADs). These
28 models range from arbitrary distributions that are chosen based on providing a good fit to the data
29 (Fisher et al. 1943), to distributions chosen based on the most likely states of generic random
30 systems (Frank 2011, Harte 2011, Locey and White 2013), to models based more directly on
31 ecological processes (Tokeshi 1993, Hubbell 2001, Volkov et al. 2003, Alroy 2015).

32 Which model or models provide the best fit to the data, and the resulting implications for the
33 processes structuring ecological systems, is an active area of research (e.g., McGill 2003, Volkov et
34 al. 2003, Ulrich et al. 2010, White et al. 2012, Connolly et al. 2014). However, most comparisons
35 of the different models: 1) use only a small subset of available models (typically two; e.g., McGill
36 2003, Volkov et al. 2003, White et al. 2012, Connolly et al. 2014); 2) focus on a single ecosystem
37 or taxonomic group (e.g., McGill 2003, Volkov et al. 2003); or 3) fail to use the most appropriate
38 statistical methods (e.g., Ulrich et al. 2010, see Matthews and Whittaker 2014 for discussion of
39 best statistical methods for fitting SADs). This makes it difficult to draw general conclusions about
40 which, if any, models provide the best empirical fit to species abundance distributions.

41 Here, we evaluate the performance of four of the most widely used models for the species abundance
42 distribution using likelihood-based model selection on data from 16,209 communities and nine
43 major taxonomic groups. This includes data from terrestrial, aquatic, and marine ecosystems
44 representing roughly 50 million individual organisms in total.

45 **Methods**

46 **Data**

47 We compiled data from citizen science projects, government surveys, and literature mining to
 48 produce a dataset with 16,209 communities, from nine taxonomic groups, representing nearly
 49 50 million individual terrestrial, aquatic, and marine organisms. Data for trees, birds, butterflies
 50 and mammals was compiled by White et al. (2012) from six data sources: the US Forest Service
 51 Forest Inventory and Analysis (FIA; USDA Forest Service 2010), the North American Butterfly
 52 Association's North American Butterfly Count (NABC; North American Butterfly Assoc. 2009),
 53 the Mammal Community Database (MCDB; Thibault et al. 2011), Alwyn Gentry's Forest Transect
 54 Data Set (Gentry; Phillips and Miller 2002), the Audubon Society Christmas Bird Count (CBC;
 55 National Audubon Society 2002), and the US Geological Survey's North American Breeding Bird
 56 Survey (BBS; Pardieck et al. 2014). The publicly available datasets (FIA, MCDB, Gentry, and
 57 BBS) were acquired using the EcoData Retriever (<http://ecodataretriever.org>; Morris and White
 58 2013). Details of the treatment of these datasets can be found in Appendix A of White et al. (2012).
 59 We modified the data slightly by removing sites 102 and 179 from the Gentry data due to issues
 60 with decimal abundances appearing in raw data due to either data entry or data structure errors. Data
 61 on Actinopterygii, Reptilia, Coleoptera, Arachnida, and Amphibia, were mined from literature by
 62 Baldrige and are publicly available (Baldrige 2013). All abundances in the compiled datasets
 63 were counts of individuals.

64 Table 1: Details of datasets used to evaluate the form of the species abundance distribution. Datasets
 65 marked as Private were obtained through data requests to the providers.

| Dataset | Dataset code | Availability | Total sites | Citation |
|---------------------------|--------------|--------------|-------------|---------------------------------|
| Breeding Bird Survey | BBS | Public | 2769 | Pardieck et al. (2014) |
| Christmas Bird Count | CBC | Private | 1999 | National Audubon Society (2002) |
| Gentry's Forest Transects | Gentry | Public | 220 | Phillips and Miller (2002) |

| Dataset | Dataset code | Availability | Total sites | Citation |
|---------------------------|----------------|--------------|-------------|--|
| Forest Inventory Analysis | FIA | Public | 10355 | USDA Forest Service (2010) |
| Mammal Community DB | MCDB | Public | 103 | Thibault et al. (2011) |
| NA Butterfly Count | NABA | Private | 400 | North American Butterfly Assoc. (2009) |
| Actinopterygii | Actinopterygii | Public | 161 | Baldrige (2013) |
| Reptilia | Reptilia | Public | 129 | Baldrige (2013) |
| Amphibia | Amphibia | Public | 43 | Baldrige (2013) |
| Coleoptera | Coleoptera | Public | 5 | Baldrige (2013) |
| Arachnida | Arachnida | Public | 25 | Baldrige (2013) |

66 **Models**

67 The majority of species abundance distributions (SADs) are constructed using counts of individuals
 68 (for discussion of alternative approaches see McGill et al. 2007 and @morlon2009). As such, the
 69 data are discrete and discrete distributions (i.e., those that only have non-zero probabilities for
 70 positive integer values of abundance) are more appropriate.

71 McGill et al. (2007) classified species abundance distribution models into five different families:
 72 purely statistical, branching process, population dynamics, niche partitioning, and spatial distribution
 73 of individuals. We evaluated models from each of these families, with some models having been
 74 derived from more than one family of processes. We evaluated the log-series, the Poisson lognormal,
 75 the negative binomial, and the Zipf distributions. All distributions were defined to be capable of
 76 having non-zero probability at integer values from 1 to infinity.

77 The log-series is one of the first distributions used to describe the SAD, being derived as a purely
 78 statistical distribution by Fisher (1943). It has since been derived as the result of ecological processes,
 79 the metacommunity SAD for ecological neutral theory (Hubbell 2001, Volkov et al. 2003), and
 80 several different maximum entropy models (Pueyo et al. 2007, Harte et al. 2008).

81 The lognormal is one of the most commonly used distributions for describing the SAD (McGill
82 2003) and has been derived as a null form of the distribution resulting from the central limit theorem
83 (May 1975), population dynamics (Engen and Lande 1996), and niche partitioning (Sugihara 1980).
84 We use the Poisson lognormal because it is a discrete form of the distribution appropriate for fitting
85 discrete abundance data (Bulmer 1974).

86 The negative binomial (which can be derived as a Gamma-distributed mixture of Poisson distri-
87 butions) provides a good characterization of the SAD predictions for several different ecological
88 neutral models for the purposes of model selection (Connolly et al. 2014). We use it to represent
89 neutral models as a class.

90 The Zipf (or power law) distribution was derived based on both branching processes and as the
91 outcome of the McGill and Collin's (2003) spatial model. It was one of the best fitting distributions
92 in a recent meta-analysis of SADs (Ulrich et al. 2010). We use the discrete form of the distribution
93 which is appropriate for fitting discrete abundance data (White et al. 2008).

94 Figure 1 shows three example sites with the empirical distribution and associated models fit to the
95 data Zipf distributions tend to predict the most rare species followed by the log-series, the negative
96 binomial, and Poisson lognormal.

97 **Analysis**

98 Following current best practices for fitting distributions to data and evaluating their fit, we used
99 maximum likelihood estimation to fit models to the data (Clark et al. 1999, Newman 2005, White et
100 al. 2008) and likelihood-based model selection to compare the fits of the different models (Burnham
101 and Anderson 2002, Edwards et al. 2007). These general best practices have recently been affirmed
102 as best practices for species abundance distributions (Connolly et al. 2014, Matthews and Whittaker
103 2014). This requires that likelihoods for the models can be solved for and therefore we excluded
104 models that lack probability mass functions and therefore likelihoods. While methods have been
105 proposed for comparing models without probability mass functions in this context (Alroy 2015),

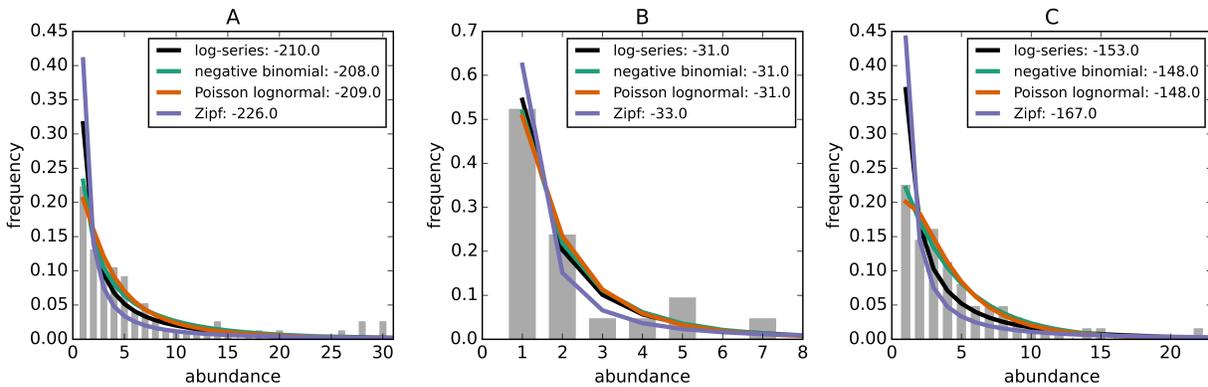


Figure 1: Example species-abundance distributions including the empirical distributions (grey bars) and the best fitting log-series (black line), negative binomial (green line), Poisson lognormal (red line), and Zipf (purple line). Distributions are for (a) Breeding Bird Survey - Route 36 in New York, (b) Forest Inventory and Analysis - Unit 4, County 57, Plot 12 in Alabama, and (c) Gentry - Araracuara High Campina site in Colombia. Log-likelihoods of the models are included in parenthesis in the legend

106 these methods have not been evaluated to determine how well they perform compared to the widely
 107 accepted likelihood-based approaches.

108 For model comparison we used corrected Akaike Information Criterion (AICc) weights to compare
 109 the fits of models while correcting for differences in the number of parameters and appropriately
 110 handling the small sample sizes (i.e., numbers of species) in some communities (Burnham and
 111 Anderson 2002). The Poisson lognormal and the negative binomial each have two fitted parameters,
 112 while the log-series distribution and the Zipf distributions have one fitted parameter each. The
 113 model with the greatest AICc weight in each community was considered to be the best fitting model
 114 for that community. We also assessed the full distribution of AICc weights to evaluate the similarity
 115 of the fits of the different models.

116 In addition to evaluating AICc of each model, we also examined the log-likelihood values of the
 117 models directly. We did this to assess the fit of the model while ignoring corrections for the number
 118 of parameters and the influence of similarities to other models in the set of candidate models. This
 119 also allows us to make more direct comparisons to previous analyses that have not corrected for the
 120 number of parameters (i.e., Ulrich et al. 2010, Alroy 2015)

121 Model fitting, log-likelihood, and AICc calculations were performed using Python (Van Rossum and
122 Drake 2011) and R (R Core Team 2015). Python packages used for analysis include numpy (Oliphant
123 2007, Van Der Walt et al. 2011), matplotlib (Hunter and others 2007), sqlalchemy (Bayer 2014),
124 pandas (McKinney and others 2010), macroecotools <https://github.com/weecology/macroecotools>,
125 retriever (Morris and White 2013), R packages used for analysis include ggplot2 (Wickham 2009),
126 magrittr (Bache and Wickham 2014), tidyr (Wickham 2016), dplyr (Wickham and Francois 2016).
127 All of the code and all of the publicly available data necessary to replicate these analyses is available
128 at <https://github.com/weecology/sad-comparison>. The CBC datasets and NABA datasets are not
129 publicly available and therefore are not included.

130 **Results**

131 Across all data sets, the negative binomial and Poisson lognormal distributions had very similar
132 average log-likelihoods (within 0.01 of one another; Figure 2). The log-likelihoods for each of these
133 distributions averaged 0.8 units higher than for the log-series distribution and 5 units higher than for
134 the Zipf distribution (corresponding to likelihoods that were twice as high and 140 times as high,
135 respectively).

136 Although the negative binomial and Poisson lognormal distributions matched the data most closely,
137 the likelihood provides a biased estimate of these distributions' ability to generalize to unobserved
138 species. AICc approximately removes this bias by penalizing models with more degrees of freedom
139 (e.g. the negative binomial and Poisson lognormal distributions, which have two free parameters
140 instead of one like the log-series and Zipf distributions). After applying this penalty, the log-series
141 distribution would be expected to make the best predictions for 69.2% of the sites. The Poisson
142 lognormal and negative binomial distributions were each preferred in about 12% of the sites, and
143 the Zipf distribution was preferred least often (6.0% of sites; Figure 3).

144 Across all data sets and taxonomic groups, the log-series distribution had the highest AICc weights

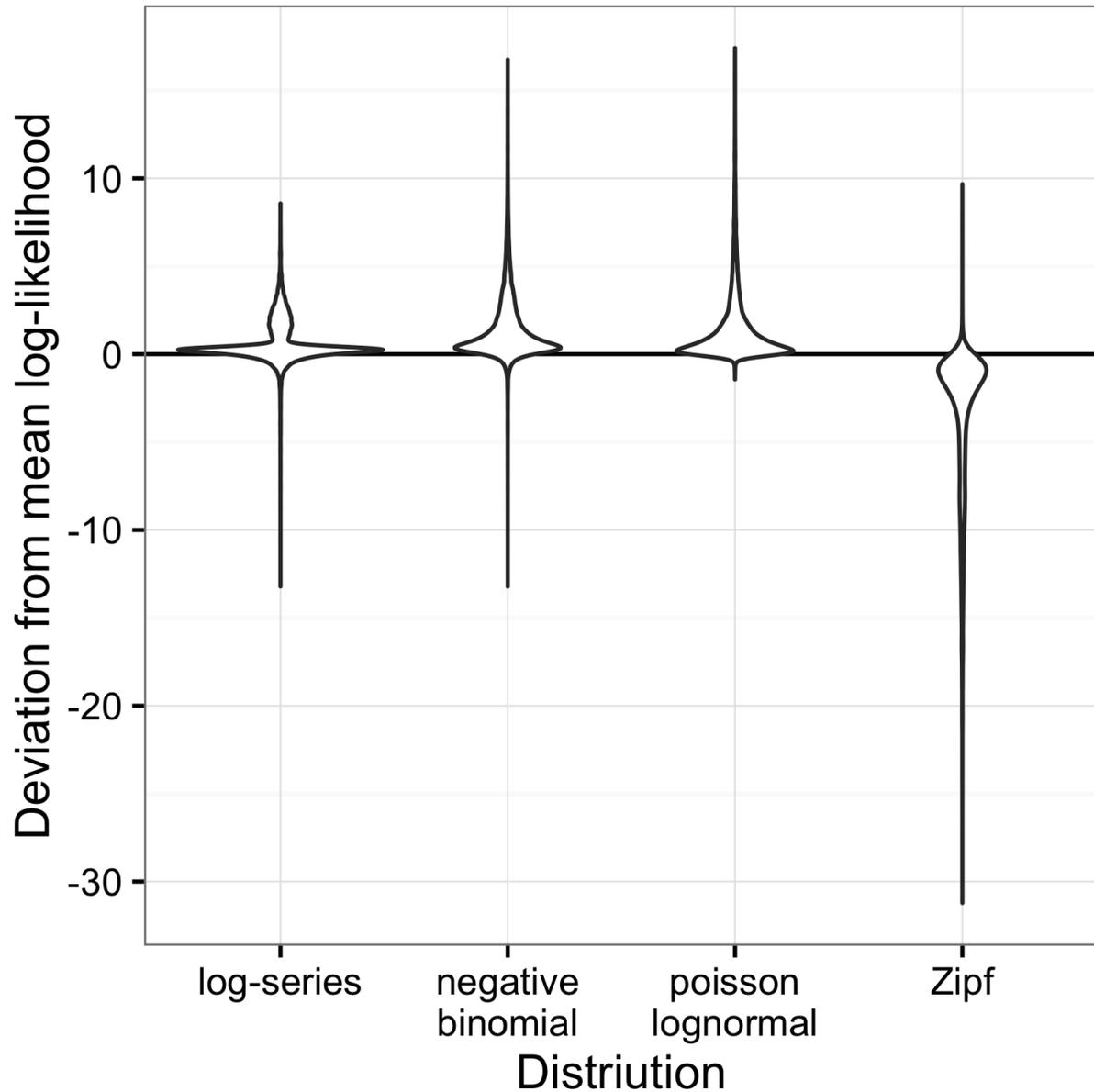


Figure 2: Violin plots of the deviation from the mean log-likelihood for each site for all datasets combined. Positive values indicate that the model fits better than the average fit across the four models.

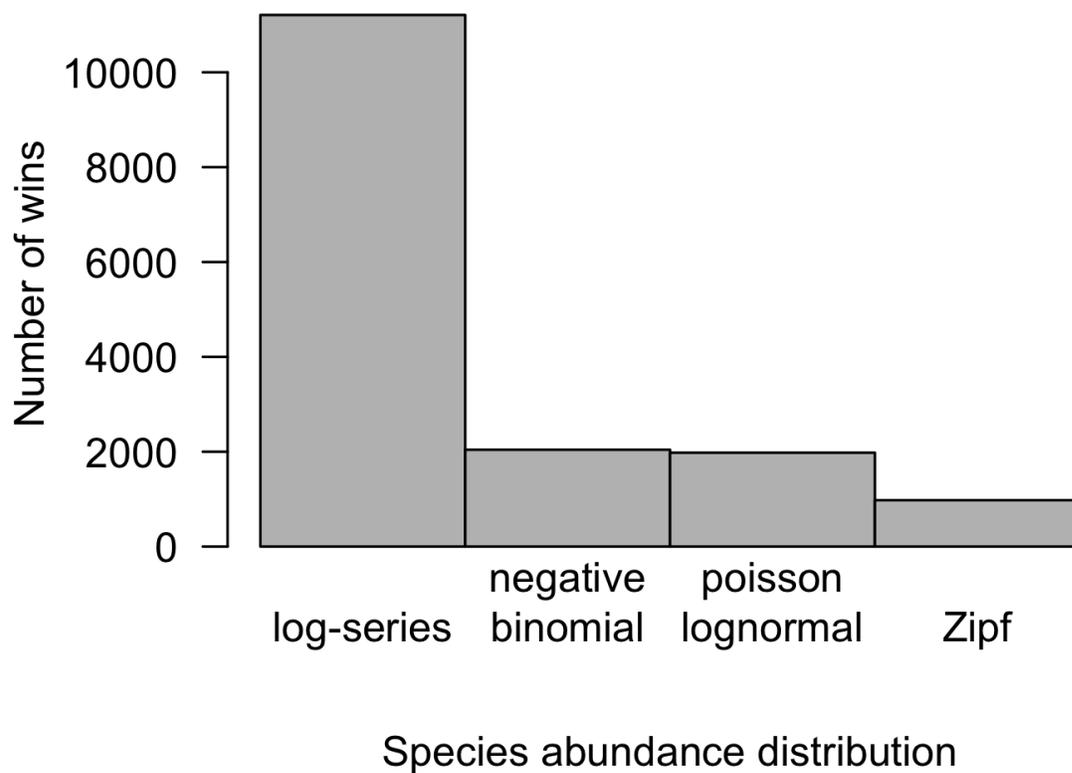


Figure 3: Number of cases where each model provided the best fit to the data based on AICc for all datasets combined.

145 more often than any other model. The negative binomial performed well for BBS, but was almost
 146 never the best fitting model for plants (FIA and Gentry), butterflies (NABA), Acintopterygii, or
 147 Coleoptera. The Poisson lognormal performed well for the bird datasets (BBS and CBC) and the
 148 Gentry tree data, but was almost never best in the FIA and Coleoptera datasets (Figure 4). The Zipf
 149 distribution only performed consistently well for Arachnida.

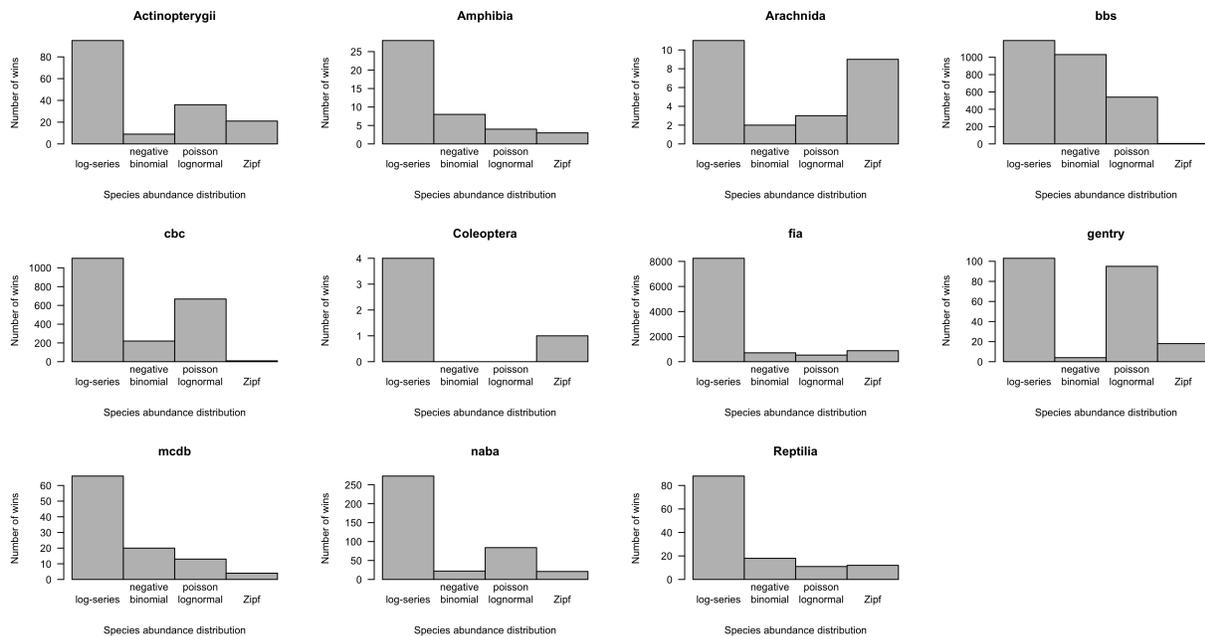


Figure 4: Number of cases where each model provided the best fit to the data based on AICc for each dataset separately.

150 The full distribution of AICc weights shows separation among models (Figure 5). Although the
 151 log-series distribution had the best AICc score much more often than the other models, its lead was
 152 never decisive: across all 16,209 sites, it never had more than about 75% of the AICc weight (Figure
 153 5). Most of the remaining weight was assigned to the negative binomial and Poisson lognormal
 154 distributions (each of which usually had at least 12-15% of the weight but was occasionally favored
 155 very strongly). The Zipf distribution showed a strong mode near zero, and usually had less than 7%
 156 of the weight.

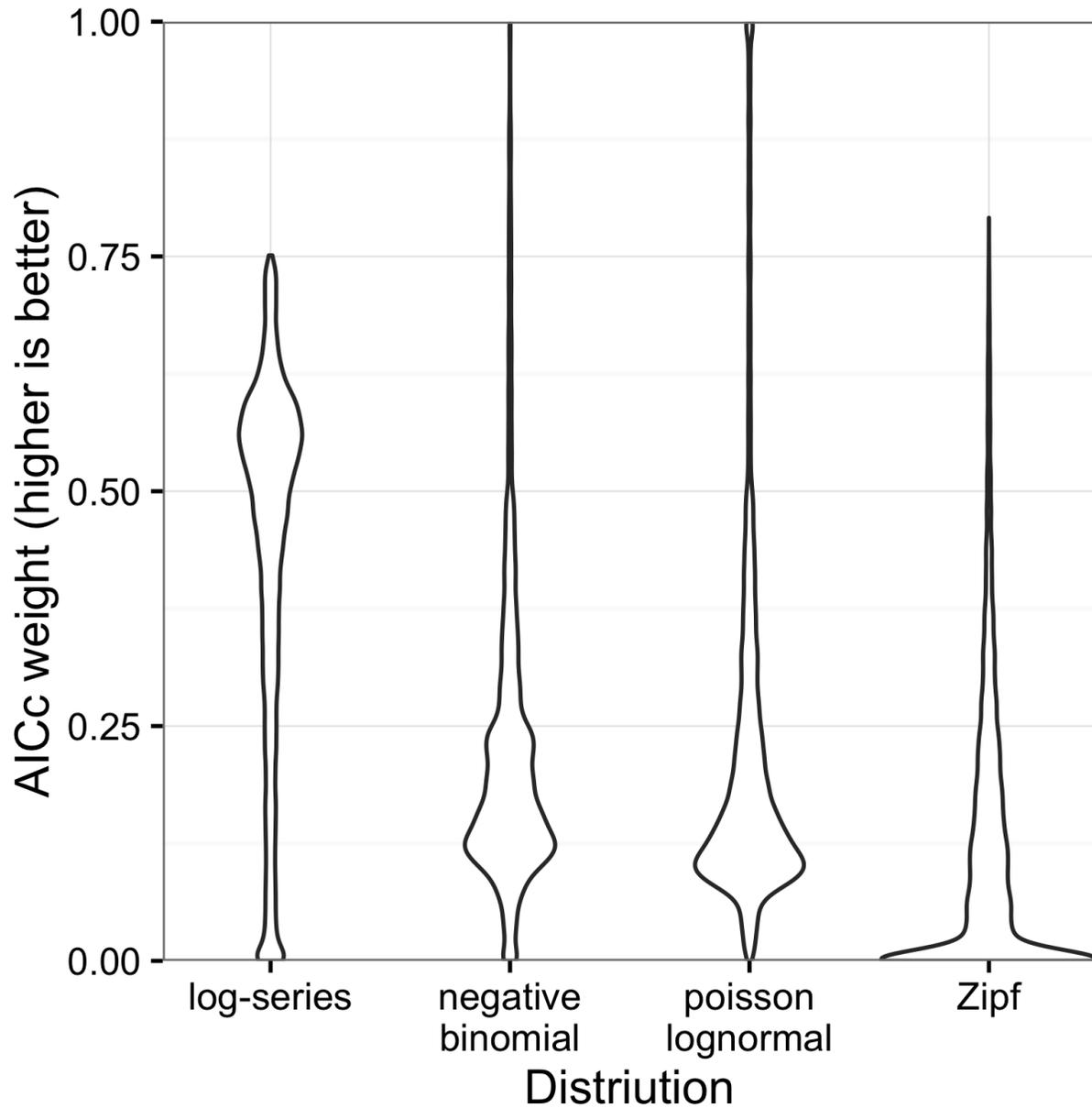


Figure 5: Violin plots of the AICc weights for each model. Weights indicate the probability that the model is the best model for the data

157 Discussion

158 Our extensive comparison of different models for the species abundance distribution (SAD) using
159 rigorous statistical methods demonstrates that several of the most popular existing models provide
160 equivalently good absolute fits to empirical data. Log-series, negative binomial, and Poisson lognor-
161 mal all had model relative likelihoods between 0.25 and 0.5 suggesting that the three distributions
162 provide roughly equivalent fits in most cases, but with the two parameters models doing slightly
163 better on average. Because the log-series has only a single parameter but fits the data almost as well
164 as the two parameter models, the log-series performed better in AICc-based model selection, which
165 penalizes model complexity. These results differ from two other recent analyses of large numbers
166 of species abundance distributions (Ulrich et al. 2010, Connolly et al. 2014) and are generally
167 consistent with a third recent analysis (Alroy 2015).

168 Ulrich et al. (2010) analyzed ~500 SADs and found support for three major forms of the SAD
169 that changed depending on whether the community had been fully censused or not. They found
170 that “fully censused” communities were best fit by the lognormal, and “incompletely sampled”
171 communities, best fit by the Zipf and log-series (Ulrich et al. 2010). In contrast we find effectively
172 no support for the Zipf across ecosystems and taxonomic groups, including a number of datasets
173 that are incompletely sampled. Our AICc value results also do not support the conclusion that the
174 lognormal outperforms the log-series in fully censused communities. The Gentry and FIA forest
175 inventories both involve large stationary organisms and were collected with the goal of including all
176 trees above a certain stem diameter. Therefore, above the minimum stem diameter, they are as close
177 to fully censused communities as is typically possible. In these communities the log-series provides
178 the best fit to the data most frequently. The discrepancy between our results and those found in
179 (Ulrich et al. 2010) may be due to: 1) their use of binning and fitting curves to rank abundance plots,
180 which deviates from the likelihood-based best practices (Matthews and Whittaker 2014) used in this
181 paper; 2) the statistical methods they use to identify communities as “fully censused”, which tend to
182 exclude communities with large numbers of singletons that would be better fit by distributions like

183 the log-series; 3) the use of the continuous lognormal instead of the Poisson lognormal.

184 Connolly et al. (2014) use likelihood-based methods to compare the the negative binomial dis-
185 tribution (which they call the Poisson gamma) to the Poisson lognormal for a large number of
186 marine communities. They found that they Poisson lognormal provides a substantially better fit
187 than the negative binomial to empirical data and that the negativie-binomial provides a better fit
188 to communities simulated using neutral models. They conclude that these analyses of the SAD
189 demonstrates that marine communities are structured by non-neutral processes. Our analysis differs
190 from that in Connolly et al. (2014) in that they aggregate communities at larger spatial scales than
191 those sampled and find the strongest results at large spatial scales. This may explain the difference
192 between the two analysis or there may be differences between the terrestrial systems analyzed here
193 and the marine systems analyzed by Connolly et al. (2014). The explanation for these differences is
194 being explored elsewhere (Connolly et al. unpublished data).

195 Alroy (2015) compared the fits of the lognormal, log-series, Zipf, geometric series, broken stick, and
196 a new model dubbed the “double geometric”, to over 1000 terrestrial community datasets assembled
197 from the literature. To incorporate the geometric series, broken stick, and the double geometric, this
198 research used non-standard methods for evaluating the fits of the models to the data, however the
199 results were generally consistent with those presented here. The central Kullback-Leibler divergence
200 statistics results show that: 1) the Zipf, geometric series, and broken stick all perform consistently
201 worse than the other distributions; 2) the double geometric, log-series, and lognormal all provide
202 the best overall fit for at least one taxonomic group; and 3) the lognormal and double geometric fit
203 the data equivalently well and slightly better than the log-series when not controlling for differences
204 in the number of parameters (Alroy’s tables S1, S2, and S3). Penalizing the two-parameter models
205 (lognormal and double geometric) for their complexity, as we do here with AICc, would likewise
206 improve the relative performance of the log-series distribution.

207 In combination, the results of these three papers suggest that in general the Zipf is a poor characteri-
208 zation of species-abundance distributions and that both the log-series and lognormal distributions

209 provide reasonable fits in many cases. Differences in the performance of the log-series, lognormal,
210 double geometric, and negative binomial, appear to be more minor. How these differences relate to
211 differences in intensity of sampling, spatial scale, taxonomy, and ecosystem type (marine vs. terres-
212 trial) remain open questions. Our analyses suggest that controlling for the number of parameters
213 makes the log-series a slightly better fitting model, at least in the terrestrial systems we studied.
214 Neither of the other papers that include the log-series (Ulrich et al. 2010, Alroy 2015) make this
215 correction and both show that it is still a reasonably competitive model even against those with
216 more parameters.

217 The relatively similar fits of several commonly used distributions emphasizes the challenges of
218 inferring the processes operating in ecological systems from the form of the abundance distribution.
219 It is already well established that models based on different processes can yield equivalent models
220 of the SAD, i.e., they predict distributions of exactly the same form (Cohen 1968, Boswell and
221 Patil 1971, Pielou 1975, McGill et al. 2007). To the extent that SADs are determined by random
222 statistical processes, one might expect the observed distributions to be compatible with a wide
223 variety of different process-based and process-free models (Frank 2009, 2011, Locey and White
224 2013). Regardless of the underlying reason that the models performed similarly, our results indicate
225 that the SAD usually does not contain sufficient information to distinguish among the possible
226 statistical processes—let alone biological processes—with any degree of certainty (Volkov et al.
227 2005), though it is possible that this result differs in marine systems (see Connolly et al. 2014).
228 A more promising way to draw inferences about ecological processes is to evaluate each model's
229 ability to simultaneously explain multiple macroecological patterns, rather than relying on a single
230 pattern like the SAD (McGill 2003, McGill et al. 2006, Newman et al. 2014, Xiao et al. 2015).
231 It has also been suggested that examining second-order effects, such as the scale-dependence of
232 macroecological patterns (Blonder et al. 2014) or how the parameters of the distribution change
233 across gradients (Mac Nally et al. 2014), can provide better inference about process from these
234 kinds of pattern.

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