## Seasonal genetic variation and genetic structure of Spodoptera exigua in northern China on the basis of 11 years of microsatellite data (#107333)

First submission

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# Seasonal genetic variation and genetic structure of Spodoptera exigua in northern China on the basis of 11 years of microsatellite data

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**Background:** The beet armyworm (BAW), *Spodoptera exigua*, is a destructive migratory pest worldwide that has caused severe economic losses in China's major crop-producing regions. To control this pest effectively, it is crucial to investigate its seasonal genetic variation and population genetic structure in northern China. **Methods:** In this study, we used ten nuclear microsatellite loci to investigate the seasonal genetic diversity and genetic structure of BAW in Shenyang, Liaoning Province, Northeast China, from 2012-2018. **Results:** Microsatellite data revealed moderate levels of genetic variation among 50 seasonal populations of BAW sampled from 2012--2018, along with significant genetic differentiation among these populations. Neighbor-joining dendrograms, STRUCTURE analysis, and principal coordinate analysis (PCoA) revealed two genetically distinct groups: the SY2012-2018 group and the SY2019-2022 group. Our results revealed seasonal variation in the genetic subconstruction at this location, which may be related to the presence of different migratory individuals throughout the year. Accordingly, our unique insights into the population genetics of BAW will contribute to the development of effective management strategies for this migratory pest.

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21	Abstract
22	Background: The beet armyworm (BAW), Spodoptera exigua, is a destructive migratory pest
23	worldwide that has caused severe economic losses in China's major crop-producing regions. To
24	control this pest effectively, it is crucial to investigate its seasonal genetic variation and population
25	genetic structure in northern China.
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28	2012-2018.
29	Results: Microsatellite data revealed moderate levels of genetic variation among 50 seasonal
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32	coordinate analysis (PCoA) revealed two genetically distinct groups: the SY2012-2018 group and
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35	the year. Accordingly, our unique insights into the population genetics of BAW will contribute to
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38	Key words: beet armyworm pnetic differentiation; genetic structure; microsatellites; migration;
39	seasonal genetic variation



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#### Introduction

pest affecting a wide range of crops, including vegetables, maize, cotton, soybeans, and ornamental 42 43 plants (Adamczyk et al., 2009; Guo et al., 2010). In general, BAW larvae feed on the leaves of host plants, resulting in reduced crop yields and potential plant death. Originally from South Asia, 44 this species is now widely distributed across the tropical and temperate regions of Europe, Africa, 45 46 North America, and Asia. (Wei et al., 2010). In China, BAW was first recorded in Beijing in the 47 1890s. It is widely distributed in the primary crop-producing regions of China and has caused significant economic losses in recent years. For example, the beet armyworm has spread to several 48 49 provinces in North China and East China, infesting a total area exceeding 2.7 million hectares (Luo et al., 2000). This pest particularly affects Welsh onions in northern China, which infest over 8,000 50 51 hectares in Tianjin and result in a 30% reduction in annual Welsh onion production (Zheng et al., 2009; Zhu et al., 2010). 52 BAW is a polyphagous insect known for its high fecundity and long-distance flight capabilities 53 54 (Feng et al., 2003; Adamczyk et al., 2009). Typically, the eggs of this species are laid on the 55 undersides of leaves. Newly hatched larvae feed gregariously on the upper surfaces of the leaves. whereas third-instar larvae begin to feed solitarily. By the fourth instar, they start consuming a 56 variety of plant parts, including leaves, petals, and pods. Pupae predominantly overwinter in the 57 58 soil, with no overwintering occurring in South China. BAW can reproduce year-round, and no diapause behavior has been observed (Zheng et al., 2011). Previous studies have indicated that 59 BAW migrates seasonally once a year in eastern China (Si et al., 2012). There are significant 60 interannual and seasonal variations in the capture number of BAW based on 11 years of monitoring 61 62 data in northern China, which will contribute to a deeper understanding of population dynamics in northern China and provide a theoretical basis for regional monitoring, early warning, and the 63 development of effective management strategies for long-range migratory pests (Ma et al., 2024). 64 In addition, chemical pesticides remain the primary method for pest control. However, the 65 prolonged use of these insecticides has led to the rapid development of resistance in BAW, notably 66

The beet armyworm (BAW), Spodoptera exigua (Lepidoptera: Noctuidae), is a major polyphagous



Therefore, effectively controlling this pest is difficult. 68 69 Populations of short-lived organisms can adapt to seasonal changes through various mechanisms, 70 including genetic polymorphisms and phenotypic plasticity. These populations typically harbor 71 significant adaptive genetic variation, enabling them to respond rapidly to environmental shifts (Brennan et al., 2019). The genetic variation and population genetic structure of a species can be 72 73 influenced by numerous factors, including climate change, ecological conditions, natural barriers, 74 migration patterns, and human activities (Fairley et al., 2000; Prugh et al., 2008; Pauls et al., 2013; Nater et al., 2013). Currently, a range of molecular markers are utilized to illuminate the 75 biogeography and evolutionary history of this species (Susanta, 2006). Owing to its moderate 76 77 evolutionary rate and distinct evolutionary pattern, the cytochrome c oxidase subunit I (COI) gene 78 is well suited for reconstructing species phylogenies (Hebert et al., 2003; Wang et al., 2014). 79 Owing to their high codominance and significant polymorphism, microsatellite markers have been widely utilized in population genetics studies (Aggarwa et a 007; Wang et al., 2007; Zhu et al., 80 2020). Previous studies have focused on investigating the genetic variability and structure of BAW 81 82 across different spatial scales (Wang et al., 2014; Wang and Zhou, 2016; Niu et al., 2006; Zhou et al., 2017). For example, previous studies have indicated that both mtDNA and microsatellite data 83 indicate low levels of genetic diversity among all populations. Moderate genetic differentiation 84 85 among some BAW populations and two genetically distinct groups in western China has been 86 detected (Wang et al., 2020). These studies provide valuable information for understanding the dispersal patterns and causes of outbreaks of pest species. However, accurate assessments of the 87 genetic diversity and population genetic structure of this pest across large temporal scales in 88 89 northern China have not been performed. 90 In the present study, we investigated the seasonal genetic variation and structure of BAW in northern China. We utilized microsatellite loci to assess genetic variation, genetic differentiation, 91 and population structure across 50 seasonal populations of BAW collected from October 2012 to 92 2022 in Shenyang, Liaoning Province. Additionally, we discuss potential management strategies 93

to chlorinated hydrocarbons and carbamates (Meinke & Ware, 1978; Chaufaux & Ferron, 1986).



for this species. This research aims to deepen our understanding of the population genetics of this moth and provide a robust theoretical foundation for developing effective pest management strategies.

Sample collection and DNA extraction. A total of 1095 individuals of BAW in the 50 seasonal

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#### **Materials & Methods**

100 populations were collected via three sex pheromone traps (Pherobio Technology Co. Ltd., Beijing, China) in a Welsh onion field (123.57°N, 41.82°E) over a period of 11 years, from June to October 101 2012 to 2022 in Shenyang, Liaoning Province. The number of trapped moths was recorded weekly, 102 and the trap cores were replaced every two weeks. All the BAW samples were preserved in 95% 103 ethanol at -20 °C and stored at the Plant Protection College, Shenyang Agricultural University, 104 105 Shenyang, China. Details regarding the locations of the populations and the number of samples 106 are provided in Table S1. The samples were collected from private land with the permission of the landowners, and none of the field surveys in this study involved endangered or protected species. 107 Genomic DNA was extracted from individual samples via Qiagen's DNeasy Blood & Tissue Kit 108 109 (Oiagen, Valencia, CA) following the manufacturer's protocols. MSI amplification and genotyping. In this study, individuals were genotyped viright loci 110 (Spe06, Spe07, Spe08, Spe09, Spe10, Spe11, Spe12, and Spe15) from a set of eight polymorphic 111 112 microsatellite loci provided by Kim et al. (2012). Each microsatellite locus was assigned a unique 113 fluorophore for fluorescent tagging of the DNA. For these isolated microsatellites, each PCR mixture consisted of 1.0 units of EasyTaq DNA polymerase, 2.5 mM dNTP mixture, 0.5 µL of 114 DNA template, 1× EasyTag® buffer (containing 2 mM MgCl<sub>2</sub>; TransGen Biotech Co., Ltd., 115 116 Beijing, China), and 0.4 μM of each primer, which was labelled with the fluorochromes HEX or FAM (Sangon Biotech, Shanghai, China). The PCR amplification conditions were as follows: 117 initial denaturation at 94 °C for 4 minutes, followed by 30 cycles of denaturation at 94 °C for 30 118 seconds, annealing at 58 °C for 30 seconds, and extension at 72 °C for 30 seconds. A final 119 extension was performed at 72 °C for 5 minutes. After amplification, the products were visualized 120



at Sangon Biotech Co., Ltd. (Shanghai, China) via an ABI 3730XL automated sequencer (Applied 121 Biosystems, Foster City, CA, USA). The microsatellite alleles were analysed with GeneMapper 122 4.0 software (Applied Biosystems). The raw read pe amplified fragment length from 1095 123 124 individuals of BAW were shown in Table S2. Microsatellite data analyses 125 Genetic variation and genetic differentiation. Micro-Checker 2.2.3 was utilized to detect errors 126 127 and null alleles in BAW microsatellite genotypes, excluding individuals with missing data (Van 128 Oosterhout et al., 2004). Genotypic linkage disequilibrium (LD) was assessed for all pairs of loci across populations via GenePop 3.4 with exact probability tests (Raymond & Rousset, 1995). An 129 exact test for Hardy-Weinberg equilibrium (HWE) was conducted for each locus as well as for all 130 131 loci within each population. 132 Genetic diversity indices, including the mean number of alleles (Na), effective number of alleles 133 (Ne), Shannon's information index (I), observed heterozygosity ( $H_0$ ), expected heterozygosity (He), and unbiased expected heterozygosity ( $uH_e$ ), were assessed via GenAlEx 6.5 (Peakall and 134 Smouse, 2006). The allelic richness  $(A_R)$ , fixation index  $(F_{ST})$ , and inbreeding coefficient  $(F_{IS})$ 135 among populations were calculated via FSTAT 2.9.3.2 (Goudet, 2002), whereas the polymorphism 136 information content (PIC) was computed via Cervus 2 (Hearne et al., 1992). To assess the degree 137 of genetic differentiation between pairs of BAW populations, we calculated pairwise  $F_{ST}$  values 138 139 via Arlequin 3.0 (Excoffier et al., 2005) and created associated heatmaps of these values via R 140 statistical software 3.0.2 (Dean and Nielsen, 2007). 141 **Temporal genetic structure.** To investigate the temporal population structure of BAW, we followed a stepwise process. First, we used POPTREE 2 to construct an unrooted tree via the 142 143 neighbor-joining method (Takezaki et al., 2010). Next, we employed a Bayesian clustering model 144 to assess the degree of genetic structure and admixture among populations. The software STRUCTURE 2.3 (Evanno et al., 2005) was utilized to identify clusters of genetically similar 145 populations. We specified an initial range of potential genotype clusters (K) from 1-10 under the 146 admixture model, assuming correlated allele frequencies among populations (Falush et al., 2003). 147



The Markov chain Monte Carlo simulation was run 10 times for each value of K, with a total of 5 148  $\times$  10<sup>5</sup> iterations following a burn-in period of 5  $\times$  10<sup>4</sup>. The most likely number of clusters was 149 determined via the  $\Delta K$  approach (Evanno et al., 2005), as implemented in Structure Harvester 150 0.56.3 (Earl and vonholdt, 2012). The optimal alignment of the ten replicate analyses for the "best" 152 K was achieved via CLUMPP 1.1 (Jakobsson, 2007) and visualized with DISTRUCT 1.1 (Guillot et al., 2012). Third, we conducted principal coordinate analysis (PCoA) on the basis of the 153 154 covariance of the genetic distance matrix via GenAl 6.41 (Piry et al., 1999). Fourth, we assessed 155 the hierarchical partitioning of genetic structure among groups through analysis of molecular variance (AMOVA), which was performed via Arlequin 3.0 (Excoffier et al., 2005). As mentioned 156 previously, populations were grouped into (1) genetic structure, (2) year groups, and (3) month 157 158 groups.

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#### **Results**

#### Seasonal and interannual genetic variation

162 In this study, we used 8 microsatellite loci to genotype 1095 individuals from 50-month 163 populations during eleven years from 2012-2022 in Shenyang, Liaoning Province, Northeast China. Low null allele frequencies per locus were observed, with an average of 0.066 (Table S3). 164 Furthermore, the average  $F_{ST}$  values calculated with and without applying the ENA correction 165 166 were 0.159 and 0.145, respectively, and these values did not differ significantly. Therefore, the presence of null alleles did not affect the  $F_{ST}$  estimations (Table S4). The average number of alleles 167 (Na) varied from 2.680 in spe06 to 10.300 in spe09, with an overall average of 6.270. The 168 maximum polymorphic information content (PIC) was 0.860 for spe09, whereas the minimum was 169 170 0.392 for spe06, resulting in an average PIC of 0.694. Eight microsatellite loci presented a low inbreeding coefficient  $(F_{IS})$  in most BAW populations. The mean observed heterozygosity was 171 0.569, whereas the expected heterozygosity was 0.597 (Table S5). 172 Overall, the eight microsatellite loci selected in this study were modestly polymorphic. Seasonal 173

genetic variation analysis indicated that the average number of alleles (Na) ranged from 2.625 in



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SY1408 to 7.625 in SY2009 and SY2209 (average = 6.270). The effective number of alleles (Ne) 175 ranged from 1.766–4.295 (average = 3.510). The observed heterozygosity (Ho) ranged from 176 177 0.382–0.813, whereas the expected heterozygosity (He) ranged from 0.359–0.673. The unbiased 178 expected heterozygosity (uHe) ranged from 0.367 in SY1408 to 0.750 in both SY2106 and 179 SY2206. Furthermore, the interannual genetic variation analysis revealed that the mean observed heterozygosity (Ho = 0.564) was comparable to the mean expected heterozygosity (He = 0.627) 180 181 across all the BAW populations. The estimates of microsatellite genetic variation varied among 182 the populations. For example, the unbiased expected heterozygosity (uHe) ranged from 0.546 in 2014 to 0.680 in 2020. The average number of effective alleles (Ne) across the different BAW 183 populations was 3.884. Additionally, the average observed number of alleles (Na) across 184 185 microsatellite loci ranged from 6.500 in SY2012 to 10.750 in SY2019, with an overall mean value 186 of 8.864 (Table 1).

**Table 1** 

#### Population genetic differentiation

- On the basis of the microsatellite data, pairwise  $F_{ST}$  values for genetic differentiation ranged from
- 190 0 to 0.473, with 1,038 out of 1,225 comparisons showing significant differences. Overall, a high
- level of genetic differentiation among the population was observed, with an average  $F_{ST}$  of
- 192 0.1452. Only a few pairwise  $F_{ST}$  comparisons did not indicate genetic differentiation, as evidenced
- by the low pairwise  $F_{ST}$  values (Figure 1).

194 **Figure 1** 

#### Temporal genetic structure

#### POPTREE analysis based on microsatellite data

A comparison of samples taken at different times of the year revealed that the BAW genetic signature population in Shenyang significantly increased throughout the year. On the basis of microsatellite data, the unrooted neighbor-joining tree, which included the 50 BAW seasonal populations, revealed two major clades: the SY2012-2018 group and the SY2019-2022 group (Figure 2). One clade corresponded to 28 populations collected from 2012–2018, and the second



clade was composed of the remaining 22 populations collected from 2019–2022. 202 Figure 2 203 204 **Bayesian clustering** 205 Using microsatellite data, we applied a clustering algorithm in STRUCTURE 2.3.3 to analyse the relationships among the 50 BAW populations in Shenyang (Figure 3). The mean LnP (D) values 206 gradually increased from K = 2, suggesting that this is likely the optimal number of primary 207 208 clusters. The highest value of  $\Delta K$  was reached at K=2. This result was consistent with the 209 hypothesis that these populations could be divided into two groups: the SY2012-2018 group and the SY2019-2022 group (Figure 3, Figure 4). One clade corresponded to 28 populations collected 210 from 2012–2018, and the second clade was composed of the remaining 22 populations collected 211 212 from 2019–2022. This finding aligned with the results from the NJ phylogenetic tree analyses. 213 Figure 3 Figure 4 214 Principal coordinate analysis (PCoA) 215 Population-based PCoA was performed on the basis of Nei's genetic distance matrix derived from 216 217 the allele frequencies of the eight microsatellite markers in the 50 BAW populations (Figure 5). The first and second axes explained 37.19% and 62.1% of the total variance, respectively. The 218 219 neighbor-joining tree and Bayesian clustering analyses, which were based on data from 1095 220 individuals, indicated the presence of two distinct groups, corroborating the effectiveness of the 221 PCoA method. Figure 5 222 Analysis of molecular variance (AMOVA) 223 224 The global AMOVA of the microsatellite genotype data from the 50 BAW populations indicated 225 that most genetic variation was partitioned between populations and individuals within those populations. Approximately 82.12% of the total genetic variation was attributed to individuals 226 within populations, whereas 17.88% was attributed to variation among populations. Interannual 227 AMOVA indicated that 15.36% of the total genetic variation could be attributed to differences 228



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between major groups, whereas the majority of the variation (81.16%) was due to variation within populations. When K = 2, AMOVA revealed that 21.74% of the total genetic variation was explained by differences between major groupings, with the most variation (73.13%) occurring within populations. Additionally, the  $F_{CT}$  for K = 2 was 0.217. The variation rates of the two groups were similar and statistically significant. Therefore, the group with K = 2 was considered the optimal grouping for these 50 BAW populations (Table 1). Consistent with the results from the NJ tree, PCoA, and STRUCTURE analyses, the AMOVAs also supported the existence of two distinct genetic groups.

237 **Table 2** 

#### Discussion

A comprehensive understanding of the genetic makeup of migratory pest populations is crucial for developing forecasting tools, biosecurity protocols, and sustainable management practices (Simon and Peccoud, 2018). Despite this importance, there is still a lack of critical insights into long-range dispersal events influenced by allelic drift and migration (Rosetti and Remis, 2012). BAW is a polyphagous species that feeds on more than 300 plant species, indicating its significant adaptive potential. Originally from South Asia, it is now widely distributed across many major cropproducing areas in China. Understanding seasonal genetic variation and genetic structure can offer valuable insights into the evolutionary and ecological processes of this species. In this study, the microsatellite markers used presented a high average number of alleles (Na), ranging from 2.625 to 7.625, indicating that they are potentially informative tools for population genetics analysis of this species (Kalinowski 2004; Rueda et al. 2011). Populations frequently exhibit substantial adaptive genetic variation, enabling rapid responses to environmental changes (Bitter et al., 2019; Brennan et al., 2019). Fluctuating selection can result in stable oscillations in the relative abundance or frequency of various alleles within a population, especially when these alleles correspond to phenotypes adapted to the differing environments encountered throughout the year (such as winter and summer morphs; Bergland et al., 2014). The microsatellite markers used in this study exhibited a high number of alleles, indicating that they are potentially informative tools



for population genetics analysis of this species (Kalinowski 2004; Rueda et al. 2011). The 256 microsatellite data indicated moderate levels of genetic variation among all 50 BAW seasonal 257 258 populations. Seasonal genetic variation analysis indicated that the average number of alleles (Na) 259 ranged from 2.625 in SY1408 to 7.625 in SY2009 and SY2209 (average = 6.270). The effective number of alleles (Ne) ranged from 1.766-4.295 (average = 3.510). This result is consistent with 260 those of previous studies (Niu et al., 2006; Wang et al., 2014; Wang et al., 2016). This high level 261 262 of diversity is likely due to BAW not experiencing significant founder effects, genetic bottlenecks, 263 or strong selection pressures (e.g., from insecticides) in this region. In contrast, several studies have indicated low levels of genetic diversity in other migratory Lepidoptera, such as monarch 264 butterflies (Danaus plexippus) (Lyons et al., 2012). 265 266 In general, species capable of dispersal exhibit minimal genetic differentiation between 267 populations. While BAW is a significant agricultural pest, there is limited information available regarding its dispersal ability (Fu et al., 2017). Migration typically homogenizes genetic 268 differentiation among populations (Wei et al., 2013). Our previous work revealed asymmetric 269 migration between the eastern and western BAW populations in China, with the eastern population 270 271 exhibiting a greater proportion of potential migrants. Additionally, the East Asian monsoon in the eastern range facilitates BAW migration and promotes gene flow (Wang et al., 2023). However, 272 the seasonal population genetic differentiation of this pest has rarely been studied. In this study, a 273 274 high level of genetic differentiation among the seasonal populations was detected (average  $F_{ST}$  = 275 0. 1452) (Figure 1). Similar findings have been reported in the diamondback moth *Plutella* xylostella, where climatic variables contribute to genetic differentiation between temperate and 276 subtropical regions (Wei et al., 2013). Spatiotemporal separation can lead to random genetic drift 277 278 and adaptive mutations, ultimately resulting in reproductive isolation and speciation (Wang et al., 279 2023). Genetic divergence is likely to be particularly pronounced in genes associated with migration and those subjected to strong selection pressures. 280 Understanding population genetic structure provides valuable insights into the evolutionary and 281 ecological processes of species. In this study, neighbor-joining dendrograms, STRUCTURE 282



analysis, and principal coordinate analysis (PCoA) were used to identify two genetically distinct 283 groups: the SY2012-2018 group and the SY2019-2022 group. We also investigated the seasonal 284 285 genetic diversity and genetic structure of BAW in Shenyang, Liaoning Province, Northeast China, 286 via six temporal samples collected over 11 years. POPTREE, PCoA, STRUCTURE, and AMOVAs, which are based on microsatellite data, divided all individuals into two clusters: the 287 SY2012-2018 group and the SY2019-2022 group (Fig 2-Figure 5). One clade corresponded to 28 288 289 populations collected from 2012-2018, and the second clade was composed of the remaining 22 290 populations collected from 2019-2022. This may be attributed to either relatively recent population expansion or the capacity of migratory populations to sustain genetic diversity over time. Such 291 turnover can impact the control of BAW, particularly in terms of pesticide resistance, as it may 292 293 lead to the rapid spread of resistance alleles (Che et al., 2015) 294 Resistance to insecticides in insects exemplifies evolutionary adaptation to environmental changes 295 (Ffrench-Constant et al., 2004). For several decades, cultural and chemical control methods have been employed to prevent the spread and damage caused by BAW. This pest has developed 296 resistance to certain insecticides. For example, it has a long history of exposure to carbamate 297 298 pesticides and has exhibited a high level of resistance to methomyl in California, USA, since 1989 (Meinke and Ware, 1978). Therefore, quantifying the level of resistance to insecticides and 299 investigating the distribution of resistance genes in relation to the genetic structure and gene flow 300 among Chinese BAW populations are essential. Understanding the dispersal ability, genetic 301 302 structure, and population demography of this pest is crucial for both elucidating the theoretical aspects of its evolution and effectively implementing pest forecasting systems. In the future, we 303 will investigate the population genetic differentiation and structure of BAW at the genomic level 304 305 to reveal its evolutionary relationships and reconstruct its population history. This research enhances our understanding of how BAW adapts to climate and ecological factors at the genomic 306 level. Additionally, by elucidating the influence of monsoon patterns on the migration dynamics 307 of BAW, we can improve predictions regarding the magnitude, timing, and geographic distribution 308 of immigrant pest populations in China. Thus, while characterizing the fine-scale seasonal genetic 309



310	structure of BAW in northern China, our work will also clarify its large-scale temporal migration
311	dynamics and provide vital information for refining monitoring, forecasting, and integrated pest
312	management (IPM) strategies.
313	
314	Conclusions
315	This study provides further data on the seasonal genetic variation and genetic structure of BAW in
316	northern China. The results support moderate levels of genetic variation and two genetically
317	distinct groups among 50 BAW seasonal populations from 2012-2018. These unique insights into
318	BAW population genetics will aid in the development of strategies for managing this highly
319	migratory pest.
320	
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339	Competing Interests
340	The authors declare that they have no competing interests.
341	
342	Author Contributions
343	Ming-Li Yu and Xian-Zhi Xiu conceived and designed the experiments, analysed the data, wrote
344	the paper, and contributed reagents/materials/analysis tools.
345	Ming-Li Yu¹ and Jin-Yang Wang performed the experiments, analysed the data, wrote the paper,
346	and prepared the figures and/or tables.
347	Ming-Li Yu, Xian-Zhi Xiu, Xin-Yi Cao and Fa-Liang Qin performed the experiments and analysed
348	the data.
349	Xing-Ya Wang and Li-Hong Zhou contributed reagents/materials/analysis tools, authored or
350	reviewed drafts of the paper, and approved the final draft.
351	
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353	The following information was supplied relating to field study approvals (i.e., approving body and
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355	Agricultural Sciences (project number: 201003025).
356	
357	Data Availability
358	The following information was supplied regarding data availability:
359	The raw measurements are available in Table S2.
360	
361	Supplemental Information
362	The following information was supplied regarding the data availability:
363	The data are available in the Supplemental Information.

364	
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## **PeerJ**

518	Figure legends
519	
520	Figure 1 Heatmap of pairwise $F_{\rm ST}$ values estimated from microsatellite data for 50 Spodoptera
521	exigua populations collected in Shenyang, Liaoning Province in northern China. *indicates
522	significant differences following Bonferroni correction. See Table S1 for population codes.
523	
524	Figure 2 Unrooted neighbor-joining phylogenetic tree based on microsatellite data from 14
525	populations of Spodoptera exigua in northern China. The numbers next to the nodes represent
526	bootstrap values.
527	
528	<b>Figure 3</b> Population structure analysis of 1095 individuals collected from 50 seasonal populations
529	of Spodoptera exigua in northern China on the basis of eight microsatellite loci. The likelihood of
30	the data is plotted against the number of genetic clusters $(K)$ for $(a)$ the mean posterior probability
31	values (mean $lnP(D)$ values) and (b) $\Delta K$ values. (c) Individual Bayesian assignment probabilities
532	for $K = 2$ are shown, with each individual represented by a single vertical line. The sampling
333	location codes can be found in Table S1.
34	
35	Figure 4 Seasonal sampling locations of Spodoptera exigua and distribution of microsatellite
36	lineages in northern China. The red triangle indicates the sampling site in Shenyang, Liaoning
37	Province (a). Lineage 1 is represented in blue, whereas Lineage 2 is shown in red (b). Groups
38	identified by STRUCTURE analysis of the microsatellite data are indicated. The population codes
39	can be found in Table S1. The monitoring and sampling position map was created via ArcGIS Pro
540	(https://www.esri.com/en-us/arcgis/products/arcgis-pro/overview) on the basis of geographic
541	coordinates. The base map utilized in the analysis originates from the World Bank
542	(https://datacatalog.worldbank.org/search/dataset/0038272).
543	

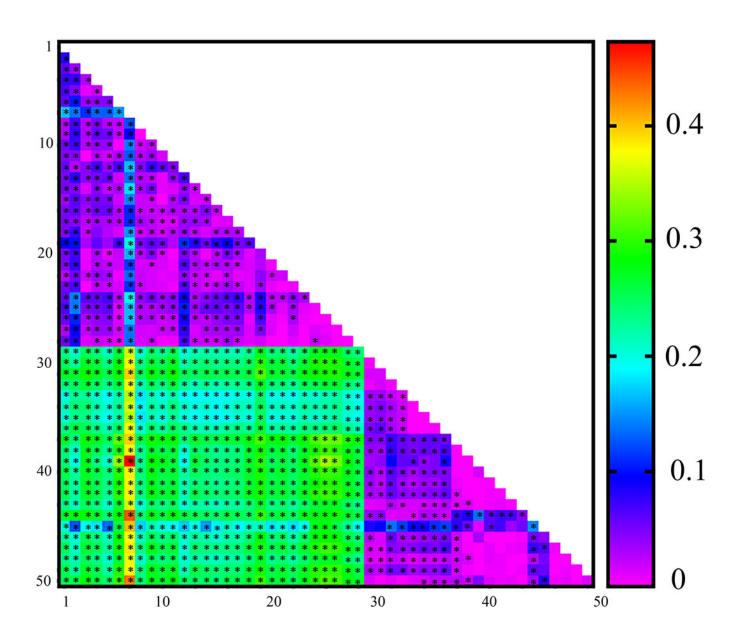




544	Figure 5 Principal coordinate analysis (PCoA) illustrating the relationships among 50 Spodoptera
545	$exigua$ seasonal populations in northern China, based on the genetic distance matrix of $F_{\rm ST}$ values
546	derived from microsatellite data. Population codes are provided in Table S1.
547	
548	Tables
549	
550	Table 1 Seasonal genetic variation in Spodoptera exigua based on eight microsatellite loci in
551	Shenyang, Liaoning Province, Northeast China from 2012-2022
552	
553	Table 2 Results of the molecular variance analysis (AMOVA) for microsatellite markers

Figure 1

Heatmap of pairwise  $F_{ST}$  values estimated from microsatellite data for 50 *Spodoptera exigua* populations collected in Shenyang, Liaoning Province innorthern China. \*indicates significant differences following Bonferroni correction. See Table S1 for population codes.





## Figure 2

Unrooted neighbor–joining phylogenetic tree based on microsatellite data from 14 populations of *Spodoptera exigua* in northern China. The numbers next to the nodes represent bootstrap values.

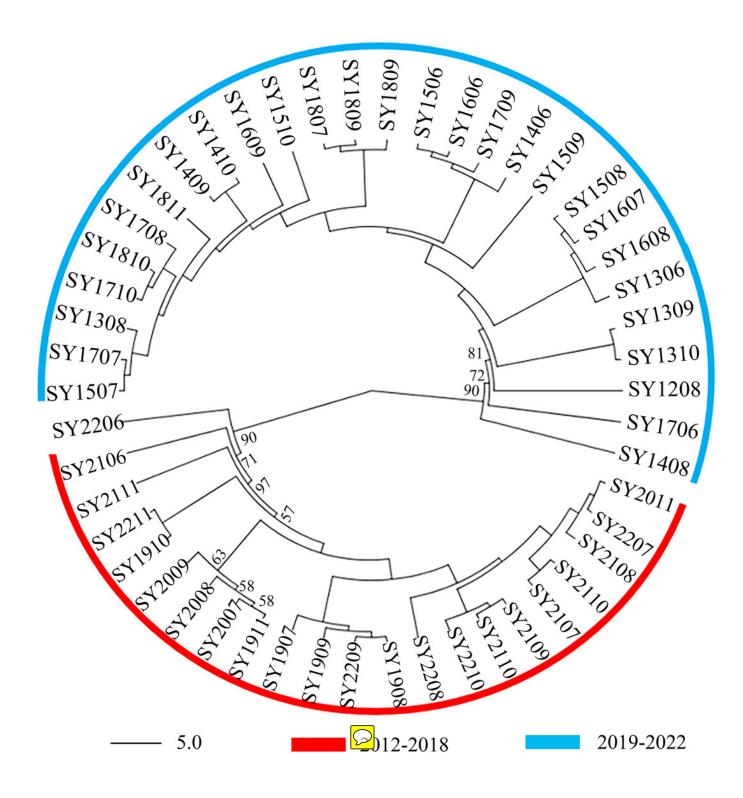
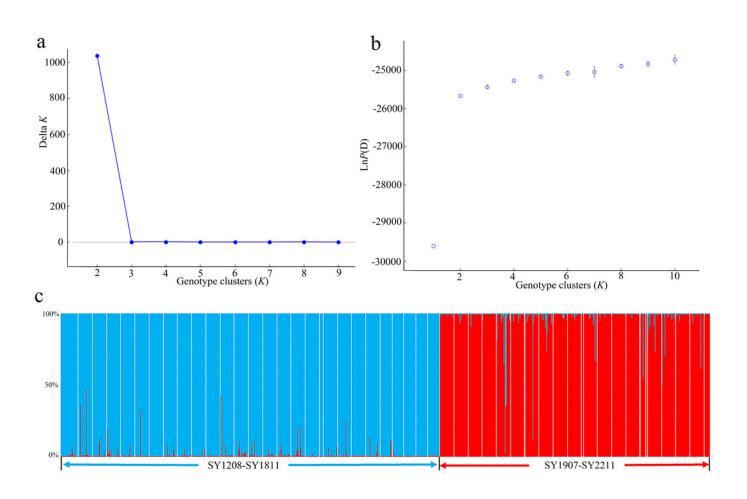




Figure 3

Population structure analysis of 1095 individuals collected from 50 seasonal populations of Spodoptera exigua in northern China on the basis of eight microsatellite loci. The likelihood of the data is plotted against the number of genetic clusters (K) for (a) the mean posterior probability values (mean InP(D) values) and (b)  $\Delta K$  values. (c) Individual Bayesian assignment probabilities for K=2 are shown, with each individual represented by a single vertical line. The sampling location codes can be found in Table S1.



#### Figure 4

**Figure 4** Seasonal sampling locations of *Spodoptera exigua* and distribution of microsatellite lineages in northern China. The red triangle indicates the sampling site in Shenyang, Liaoning Province (a). Lineage 1 is represented in blue, whereas Lineage 2 is shown in red (b). Groups identified by STRUCTURE analysis of the microsatellite data are indicated. The population codes can be found in Table S1. The monitoring and sampling position map was created via ArcGIS Pro ( <a href="https://www.esri.com/en-us/arcgis/products/arcgis-pro/overview">https://www.esri.com/en-us/arcgis/products/arcgis-pro/overview</a>) on the basis of geographic coordinates. The base map utilized in the analysis originates from the World Bank (https://datacatalog.worldbank.org/search/dataset/0038272).

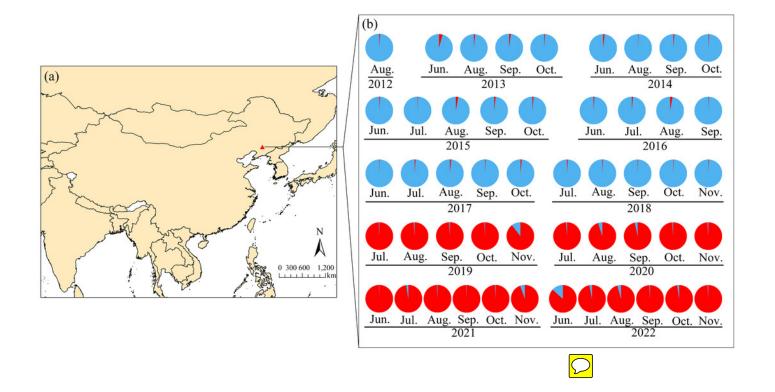
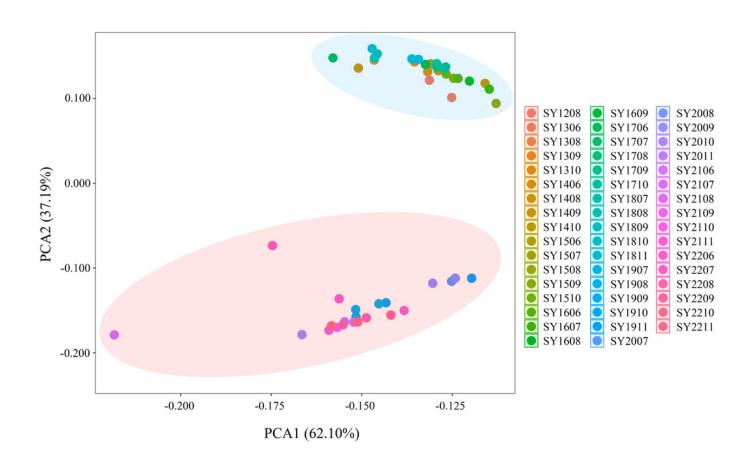




Figure 5

Principal coordinate analysis (PCoA) illustrating the relationships among 50 *Spodoptera* exigua seasonal populations in northern China, based on the genetic distance matrix of  $F_{ST}$  values derived from microsatellite data. Population codes are provided in Table S1.





## **Table 1**(on next page)

Table 1

Seasonal genetic variation in *Spodoptera exigua* based on eight microsatellite loci in Shenyang, Liaoning Province, Northeast China from 2012-2022



1 Table 1 Seasonal genetic variation in Spodoptera exigua based on eight microsatellite loci in

2 Shenyang, Liaoning Province, Northeast China from 2012-2022

Pop	Na	Ne	I	$H_0$	Не	иНе	Ap	Hs
SY2012	6.500	4.295	1.411	0.546	0.658	0.669	0.125	0.6100
SY2013	8.625	3.886	1.425	0.553	0.631	0.634	0.1250	0.6344
SY2014	8.000	2.943	1.212	0.452	0.543	0.546	0.2500	0.5470
SY2015	7.875	3.807	1.362	0.529	0.618	0.621	0.2500	0.6210
SY2016	9.125	3.777	1.403	0.526	0.612	0.615	0.3750	0.6159
SY2017	8.375	3.869	1.359	0.511	0.601	0.604	0.2500	0.6040
SY2018	7.625	3.675	1.279	0.521	0.561	0.564	0.1250	0.5643
SY2019	(10.750)	4.064	(1.517)	0.557	0.656	0.659	0.7500	0.6593
SY2020	(10.625)	4.209	(1.527)	0.688	0.677	0.680	0.6250	0.6801
SY2021	9.375	4.092	1.500	0.647	0.668	0.671	0.1250	0.6714
SY2022	10.625	4.105	1.527	0.670	0.668	0.671	0.2500	0.6710
Mean	6.270	3.510	1.272	0.569	0.597	0.619	0.052	0.6211

<sup>3</sup> Abbreviations:  $N_a$ , observed number of alleles;  $N_e$ , effective number of alleles; I, Shannon's information index;

6

7

<sup>4</sup>  $H_0$ , observed heterozygosity;  $H_0$ , expected heterozygosity;  $H_0$ , unbiased expected heterozygosity;  $H_0$  unbiased expected heterozygosity;

<sup>5</sup> of private alleles;  $H_S$ , gene diversity.







## Table 2(on next page)

Table 2

Results of the molecular variance analysis (AMOVA) for microsatellite markers

Table 2 Results of the molecular variance analysis (AMOVA) for microsatellite markers

Source of variation		Sum of a guayag	Variance	Percentage	F-statistics	
Source of variation	d.f. Sum of squares		components	variation (%)	F-statistics	
Global analysis						
Among populations	49	1327.779	0.561 Va	17.88		
Within populations	2140	5510.564	2.575 Vb	82.12	$F_{\rm ST} = 0.179^{***}$	
Total	2189	6838.343	3.136	100.00		
Hierarchical AMOVA ( $K = 11$ )						
Among groups	10	1042.527	0.487Va	15.36	$F_{\rm ST} = 0.188^{***}$	
Among populations within groups	39	285.207	0.111Vb	3.49	$F_{\rm SC} = 0.041^{***}$	
Within populations	2140	5510.564	2.575Vc	81.16	$F_{\rm CT} = 0.153^{***}$	
Total	2189	6838.343	3.173			
Hierarchical AMOVA ( $K = 2$ )						
Among groups	1	825.349	0.766 Va	21.74	$F_{\rm ST} = 0.269^{***}$	

Among populations within groups	48	502.430	0.181Vb	5.13	$F_{\rm SC} = 0.066^{***}$
Within populations	2140	5510.564	2.575Vc	73.13	$F_{\rm CT} = 0.217^{***}$
Total	2189	6838.343	3.521	100.00	

<sup>2</sup> *d.f.*, degrees of freedom; \*\*\*P < 0.001: significance level.